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A FORECASTING IMPLEMENTATION OF BOX-JENKINS TIME SERIES ANALYSIS

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A FORECASTING IMPLEMENTATION OF BOX-JENKINS TIME SERIES ANALYSIS

ABSTRACT

The ultimate effect of a decision generally depends on the outcome of factors that can not be foreseen at the time decision is made. Among wide variety of forecasting methods, the Box-Jenkins approach is known as the application of the more general and statistical based methods of time series analysis.

This thesis covers the implementation of Box-Jenkins approach by using computer. A computer software is developed for building univariate and bivariate models, and for making the forecasts. The underlying principles of the Box-Jenkins approach are presented, and a methodology of using the approach is suggested.

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BOX-JENKINS ZAMAN SERİSİ ANALİZİNİN BİR TAHMİN UYGULAMASI

ÖZET

Bir kararın nihai etkisi genellikle sonuçları önceden kestirilemeyen bir takım sebeplerin sonuçlarına bağlıdır. Geleceği tahmin etmek amacı ile pek çok metod geliştirilmiştir. Bunlar arasında Box-Jenkins yaklaşımı zaman serileri analizinin genel ve istatistiğe dayalı yöntemlerinin tahmin yapmak için uygulanması olarak bilinir.

Bu çalışma Box-Jenkins yaklaşımının bilgisayar aracılığı ile uygulanmasını kapsamaktadır. Bu amaçla tek zaman serileri ve çift zaman serileri modelleri oluşturmak ve tahmin yapmak için bir bilgisayar yazılımı geliştirilmiştir. Box-Jenkins yaklaşımının kullanılabilmesi için bir metodoloji önerilmiştir.

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AR(p)	Autoregresisve process of order p
ARIMA(p,d,q)	Autoregressive integrated moving average process of order (p,d,q)
ARMA(p,q)	Mixed process of order (p,q)
В	Backward shift operator
Bp	Dead time operator of order b of a transfer function
c _k	Autocovariance coefficient estimate at lag k, k = 0,1,,K
MA(q)	Moving average process of order q
N _t	Noise series for $t = 1, 2,, N$
r _k	Autocorrelation coefficient estimate at lag k, k = 0,1,2,,k
SOE	Sum of errors
SSE	Sum of squares of errors
u _t	White noise process, t = 1,2,,N
v(B)	Transfer function of the filter between an input and an output
v _k	Impulse response weights, k = 0,1,2,,K
X _t	Input series of a system, t = 1,2,,N
Ź _t (1)	Forecast made at origin t for lead time ℓ, ℓ = 1,2,,L
δ(B)	Left hand side linear operator of order r of a transfer function model
θ(B)	Moving average operator of order q
σ²	Variance of white noise series

φ (B)	
Ψ(B)	
ω(B)	

 ${}^{{\scriptscriptstyle \Delta} d}$

Autoregressive operator of order p Linear filter operator Right hand side operator of order s

Differencing operator of order d

of transfer function model

I. INTRODUCTION

1.1 Forecasting and Planning

Forecasting may be defined as the estimation of the future based on the past by using the methodologies which are developed for this purpose. It is used by decision and policy makers as an aid to determine goals and targets, to understand the environment and causal factors that affect it, and to understand the uncertainities in the future and to force decision about the level of risk appropriate.

Planning on the other hand may be defined as attempts at purposeful, future oriented decision making (1). A firm is generally exposed to uncontrollable external events, and controllable internal events. Forecasting applies directly to uncontrollable external events, while decision making applies directly to controllable internal events. Planning is considered the link that integrates them.

1.2. Current Status of Forecasting

1.2.1. Theory and Applications

As a result of increasing uncertainty of the environment, a number of estimation and prediction methods are developed both in theory and practice for organizational forecasting. The current forecasting literature can be classified into two groups with regard to the theory and practice of the forecasting methods.

In the first group, the works on forecasting have generally been written by the specialists who have accomplished the theoretical formulation and verification of specific methods and who are trying to convey the state of the art knowledge to a group of specialists. In this area, the work of Box and Jenkins (2) does an excellent job in developing and providing statistical properties of specific classes of forecasting methods. The second group of the literature is concerned on translating what is theoretically possible and computationally feasible into a form that can be easily understood and applied. In this area, the work of Makridakis et al (3) is a complete study on forecasting where they put a broad range of forecasting methods into a form that can easily be understood and applied by nonspecialists. The work of Nelson (4) is related with the essence of the application of Box-Jenkins models and their forecasting.

There are many research articles on forecasting. Two important applications of Box-Jenkins method are worth to mention. First is the univariate application of the method to actual time series. It includes the comparison of the method with the exponentially smoothed method. It is concluded that two methods perform equally well on the given data (5).

Secondly, in a recent study, the procedural steps in the Box-Jenkins transfer function method are demonstrated in an application to the advertising and sales relationship with particular focus on the advertising lag structure (6).

1.2.2. Forecasting Methods

To deal with the increasing variety and complexity of managerial forecasting problems, a wide variety of forecasting methods are developed that originated from several major fields of study. Although various classification schemes are possible depending on the framework chosen, the generally accepted classification is given here for describing the methods. According to this classification, the existing forecasting methods fall into two major categories: quantitative methods, and qualitative methods.

<u>Quantitative methods</u> can be applied when sufficient information about the past is available in the form of numerical data, and it can be assumed that some aspects of the existing patterns will continue into the future. These methods are further divided into two categories.

(1) <u>Causal methods</u>. The objective of forecasting using causal methods is to determine the cause-effect relationship of the factors to be forecasted with one or more independent variables by assuming that this relationship will hold into the future.

(2) <u>Time series methods</u>. The objective of these forecasting methods is to determine the pattern in the historical data series and extrapolate that pattern into the future. The system is treated as a black box and, as a result, the factors affecting its behavior are not considered.

<u>Qualitative methods</u> require subjective estimation, intuitive thinking, judgement, and accumulated knowledge through the opinions of the experts.

Another useful classification is to divide forecasting approaches into two groups: statistical methods, and filtering methods.

ARIMA schemes which were first introduced by Box and Jenkins are quite sophisticated mathematical models. Their work also includes an extension of ARIMA scheme into multivariate models which is known as transfer function analysis. Regression methods and ARIMA schemes are accepted as statistical approaches to forecasting.

Filtering methods, first introduced by Kalman (7), are engineering approaches. The name of Kalman filter is used synonymously with Bayesian forecasting. In this approach, it is assumed a fixed model with varying parameters and variances.

1.2.3. Selection of the Forecasting Method

The selection of the forecasting method for a given situation is a decision problem. The information supplied from this system is to be used to improve the decision process. There are several useful criteria that are used to select, to compare, and to evaluate the competing methods. The first problem to be solved is to define the <u>variables of the forecast</u> to be analysed and predicted.

An overview of decision criteria will clarify the interrelationships among criteria and need to select a forecasting method that best meets all the requirements of a given situation.

Level of detail is related with the decision on what level the forecast to be made such as regional or market demand, or product or product group demand. An important class of decisions involves the time elements: forecast period, and forecast lead time or forecast horizon. The <u>form</u> of the final forecast is also an important consideration. Various methods provide different outcomes such as mean, an estimate of the standard deviation of forecast error, or a probability interval. <u>Technical sophistication</u> is a determinant on the applicability of a method for a given situation, because the application of a method will be restricted with the capabilities and interests of the people who will make and use forecasts.

Before choosing a method, the extensiveness, accuracy, currency, and representativeness of the <u>available data</u> must be considered, because the ability of many forecasting methods is limited with the amount of available data. The <u>type of data series</u> is another factor on the selection of a forecasting method. The more general classification is macro series and micro series. The <u>pattern of the data</u> must be taken into consideration, because, there are many methods that can only cope with a certain data pattern. Four types of data patterns can generally exist in the data series: horizontal, trend, seasonal, and cyclical. <u>The variability of</u> <u>data series</u> is a result of the process generating the variable under study. A number of simple methods are available for the stationary series, but more sophisticated methods are developed to handle nonstationary series.

The element of <u>cost</u> depends on the developement, installation, maintenance and operation of the method, data requirements, computer requirements, and human sources requirement.

<u>Forecast accuracy</u> have an important impact on the selection of the forecasting method. Altough there are several statistical measures of accuracy, it is not possible to propose a robust measure of a specific method in common use for all situations. The relative importance of a decision criterion depends on the forecasting situation. The criteria mentioned above have interrelationships, therefore the trade offs for a given situation must be correctly established. The costs and benefits of a forecasting application must be considered in the evaluation of alternative forecasting methods.

Chambers et al propose that the stage of the product life cycle for which it is making the forecast is an important consideration. Their approach is based on matching methods with the forecasting need as determined by product life cycle (8).

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In another study, several different criteria for evaluating alternative forecasting methods are described and those criteria are used to match the situation with the most appropriate forecasting method. It is concluded that simple methods can do as well in a wide variety of cases, therefore the mathematically sophisticated methods should not be selected, unless there is a strong evidence that those methods will do better (9).

In a recent study, it is proposed that each method has strengths and weaknesses, every forecasting situation is limited by constraints. An evaluation of twenty common forecasting methods are arrayed against sixteen evaluative dimensions. It is concluded that extrapolations can be improved by combining forecasts, or simulating a range of input assumptions, or selectively applying judgement (10).

The research articles on this subject are numerous, only a few of the studies are included here (11), (12), (13), (14), (15).

1.3. Overview and Framework for Thesis

The need to carry out this study has been arisen from the sales forecasting problem of an existing group of companies. A marketing company, which markets hundreds of glass products, has been chosen within this group in order to define forecasting problem, to establish objectives and determine contents of the study, and to apply resulting model to the chosen sales forecasting situation. Although this study is originated from a specific forecasting situation, the forecasting model developed can be applied to any univariate and bivariate time series forecasting situation.

1.3.1. Definition of the Forecasting Problem

The forecasts are made for product groups which are formed by aggregating the products within the same product class. The unit of measure is determined as units of product sold during a month. Forecasts are made by using only internal data sources.

The data sources are investigated in two dimensions. The first is sales series or as it is frequently called <u>output series</u>. Secondly, the

factors that affect sales, which are called <u>input series</u> or <u>leading</u> <u>indicators</u>, are sought out. These are price, advertisement expenditures, and sales promotion expenditures. It is found out that about fifteen years of monthly data available for sales series, price data are also available for the same length of time, but it is not possible to determine the expenditures of advertisement and promotion on product group basis.

At this point, it is better to distinguish the difference between aggregate and point data. <u>Aggregate data</u> represent the value of a variable accumulated over a period of time, while <u>point data</u> indicate the value of a variable at specific points in time. Of the data mentioned above, price data are point data.

The changes in price series do not exactly match with sales series in this forecasting situation, consequently to establish a proper relationship between price and sales will not be efficient.

Time series data must be collected in equispaced time intervals. In this respect, a month can not be assumed to be a period. A <u>trading</u> days adjustment on data is recommended to increase forecast accuracy (16).

Another fact is that the available data do not properly represent actual demand, because it is collected on delivery basis. Actual demand may be best represented by collecting data on order basis.

The investigation on sales data by using tools, such as graph of data, autocorrelations, differencing, etc., has exhibited horizontal, trend or seasonal patterns. Also, many of the series have shown the indications of nonstationarity.

The problem may be stated as to find and implement a forecasting method that meets all the requirements of the given situation.

1.3.2. Objectives of This Thesis

The main objective of this study is to develop an interactive computer software that is used to build, and to forecast with univariate and bivariate Box-Jenkins models.

The other objective is to explain statistical concepts and underlying Box-Jenkins modelling procedures in a comprehensive way, and to provide guidelines for building Box-Jenkins models properly and quickly.

1.3.3. Evaluation of Box-Jenkins Approach

Box and Jenkins have effectively assembled in a comprehensive way the relevant information required to understand and use time series ARIMA (Autoregressive/Integrated/Moving Average) processes, their names have frequently been used as synonymous with the general ARIMA processes applied to time series analysis, forecasting, and control. Their work includes univariate time series analysis, multivariate time series analysis, and design of discrete control schemes. The theoretical aspects of Box-Jenkins time series are quite sophisticated. The relative development time of the method and to build a forecasting model take time, but the forecasts can be made quickly. Box-Jenkins approach to forecasting is apparently the most accurate with regard to mean square error, and the most developed statistical method presently available.

The time series are fitted with a mathematical model which is optimized on parameters in order to assign smaller errors to history than any other model. Box and Jenkins propose a general class of models for forecasting, their approach is appropriate to handle various data patterns.

Box-Jenkins philosophy of model building for time series includes two principles (17). First is the principle of <u>parsimony</u> that can be described as the smallest number of parameters that should be employed for adequate representation of underlying model of a series.

The second principle is to apply <u>iterative procedure</u> in the selection of a model. There are three stages in their approach to model building.

(1) <u>Identification</u>. The methods are proposed to define models which may be good representation of underlying generating mechanism.

(2) <u>Estimation</u>. The model selected is fitted to data and the parameters are estimated by minimizing sum of squares of errors.

(3) <u>Diagnostic checking</u>. Adequacy of the fitted model is tested by using available statistical measures; and then, causes of lack of fit, if it exists are diagnosed.

These three stages of the approach are iterated until an appropriate representation is found. The selected model is then applied to make the forecasts.

In this study, the basic theory, modelling procedures, and relevant algorithms are largely drawn from the work of Box and Jenkins. The basic notation and terminology are also adopted from their work.

II. BOX-JENKINS MODEL BUILDING AND FORECASTING

In this chapter, first the underlying theory of Box-Jenkins models is presented. Then, the statistical tools used in time series analysis, and methods for building, identifying, fitting and checking models for time series are illustrated in accordance with the three stage Box-Jenkins approach to model building: identification, estimation, and diagnostic checking. Finally, the forecasting version of the developed model is briefly explained.

A phenomenon that evolves through time according to probabilistic laws is called <u>stochastic process</u>, it is simply referred as <u>process</u> in this study. The time series to be analysed may then be considered as one particular realization of a variable, from an infinite population of such realizations of that variable, produced by the underlying probability mechanism which is generally called <u>generating mechanism</u> of the process. The three stage Box-Jenkins procedure is designed to find a model that is a good representation of the unknown underlying process.

2.1. Univariate Box-Jenkins Models

ARIMA is an acronym for Autoregressive/Integrated/Moving Average. An ARIMA process refers to the particular generating mechanism which describes the evolution of observations through time, and the derivation of the conditional distribution of future realizations. The general nonseasonal model is represented by ARIMA (p,d,q), where p denotes order of the autoregressive process, d denotes degree of differencing to achieve a stationary mean, and q denotes order of the moving average process. There is no limit to the variety of ARIMA models; p,d, and q are nonnegative integers.

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2.1.1. Fundamentals of ARIMA Processes

The underlying logic in Box-Jenkins approach is to exploit the dependency relationships in successive observations of the time series. Thus, a time series of that type may be considered as generated from a time series of random "shocks" u_t . These shocks are random drawings from a fixed distribution, and they are usually assumed to be normally, independently, identically distributed having a mean zero and a constant variance σ_u^2 . The sequence of random variables u_t , u_{t-1} , u_{t-2} ,... is often called white noise process.

The white noise process u_t is supposed to be transformed to the process Z_t , which represents time series observations, by a <u>linear filter</u> so that

$$Z_{t} = \mu + u_{t} + \Psi_{1} u_{t-1} + \Psi_{2} u_{t-2} +$$
(1)

$$= \mu + \Psi(B) u_{+}$$

where

B is backward shift operator such that $B^{m}Z_{t} = Z_{t-m}$, m is nonnegative integer, $\Psi(B) = 1 + \Psi_{1}B + \Psi_{2}B^{2} + ...$ is the linear operator, μ is a parameter that represents the level of the process.

If this sequence of Ψ_1 , Ψ_2 ,... weights are finite, or infinite and convergent, the filter is said to be <u>stable</u> and the process Z_t to be stationary. The parameter μ is then the mean about which the process varies. Otherwise, Z_t is nonstationary and μ can be regarded as a reference point for the level of the process. Box-Jenkins models are based on this idea of linear filtering.

An ARIMA(p,0,0) or AR(p) process is called autoregressive process of order p so that the current value of the deviations from mean of the process is represented by a linear combination of previous values of the process and current shock u_t . Let \tilde{z}_t be deviations from μ , $\tilde{z}_t = z_t - \mu$. Then AR(p) process can be written as

$$\tilde{z}_{t} = \varphi_1 \tilde{z}_{t-1} + \varphi_2 \tilde{z}_{t-2} + \dots + \varphi_p \tilde{z}_{t-p} + u_t$$
 (2)

or

$$\varphi(B) \tilde{z}_t = u_t$$

where

 $\varphi(B) = 1 - \varphi_1 B - \varphi_2 B^2 - \dots - \varphi_D B^D$ is the

autoregressive operator of order p.

It is required that $\varphi(B)$ must be stationary.

Another useful representation is moving average process of order q, ARIMA(0,0,q) or MA(q). The current value of \tilde{z}_t is expressed as a finite number of previous shocks plus current shock u_t . That is

$$\tilde{z}_{t} = u_{t} - \theta_{1}u_{t-1} - \theta_{2}u_{t-2} - \dots - \theta_{a}u_{t-a}$$
 (3)

or

$$\tilde{z}_t = \theta(B) u_t$$

where $\theta(B) = 1 - \theta_1 B - \theta_2 B^2 - \ldots - \theta_q B^q$ is the moving average operator of order q. It is required that $\theta(B)$ must be invertible.

It is interesting to note that an AR process of infinite order can be represented by an MA process of finite order. Also, an MA process of infinite order can be represented by an AR process of finite order. This characteristic of ARIMA process suffices to explain how to build parsimonous models.

The general ARIMA process of order (p,d,q) is defined by

 $\varphi(B) \nabla^{d} z_{+} = \Theta_{0} + \Theta(B) u_{+}$

where

 ∇^{d} = (1-B)^d is the differencing operator,

 $\boldsymbol{\theta}_{0}$ is a constant which denotes the deterministic shift in the process.

Letting $\omega_t = \nabla^d z_t$ then Equation (4) can be written as

$$\varphi(B) \omega_{+} = \Theta_{0} + \Theta(B) u_{+}$$

or

 $\varphi(B) \widetilde{\omega}_{t} = \Theta(B) u_{t}$

(5)

(4)

(7)

where ...

$$\tilde{\omega}_t = \omega_t - \mu_\omega$$
.

Another useful representation is provided by letting $\zeta(B) = \varphi(B) \nabla^d$. Then the process becomes

$$\zeta(B)$$
 $\tilde{z}_{t} = \Theta(B) u_{t}$ (6)

The process defined by Equation (6) can be represented as a linear filter of infinite order such that

$$\tilde{z}_t = \frac{\Theta(B)}{\zeta(B)} u_t$$

or

$$\tilde{z}_t = \zeta^{-1}$$
 (B) $\Theta(B)$ u

2.1.2. Identification of ARIMA Models

The statistical tools for the analysis of time series are proposed as <u>autocorrelation function</u> (acf), <u>partial autocorrelation function</u> (pacf), and <u>differencing</u>. Box and Jenkins also propose that spectral analysis is a useful device of analysing time series (18).

Estimates of autocorrelations of any time series are computed from the sample data. Let ${\bf c}_k$ be the autocovariance coefficient at lag k, it can be written as

$$c_{k} = \frac{1}{N} \sum_{t=1}^{N-k} [(z_{t} - \bar{z})(z_{t+k} - \bar{z})]$$
(8)

for k = 0, 1, 2, ..., K

where \bar{z} denotes the sample mean

$$\bar{z} = \frac{1}{N} \begin{array}{c} N \\ \Sigma \end{array} \begin{array}{c} z \\ t \\ t \\ t \end{array}$$

N is the number of observations in the series. Let r_k be the autocorrelation coefficient at lag k, it is computed as follows:

(9)

$$r_{k} = \frac{c_{k}}{c_{0}}$$
, $k = 0, 1, 2, ..., K.$ (10)

A graph of autocorrelation coefficients is called autocorrelation function (acf).

Partial autocorrelations are used to measure the degree of association between z_t and z_{t+k} , but the effects of intervening z's are somehow partialled out. A plot of partial autocorrelation coefficients is called partial autocorrelation function (pacf).

ARIMA processes provide a general class of models. The selection of the most appropriate model for the given time series requires experience. In general, there are several useful guidelines for stationary series.

(1) A pure stationary AR process has a theoretical acf that drops off to zero, but it has a pacf that cuts off to zero after lag p, the order of AR process. The standard error of the partial autocorrelation coefficients after lag p is approximated as

$$se_{r_k} = 1/\sqrt{n}$$

- (2) A pure MA process has a theoretical acf that cuts off to zero after lag q, the order of MA process, but it has a theoretical pacf that drops off to zero. The estimated autocorrelations of order q+1, and higher, are approximately, independently distributed with variance 1/n.
- (3) A stationary mixed ARMA process has
 - a) a theoretical acf that tails off toward zero after the first q-p lags;
 - b) a theoretical pacf that tails off toward zero after the first q-p lags.

2.1.3. Estimation of ARIMA Models

At the estimation stage, the values of the parameters, of the model selected at the identification stage that seems promising to provide parsimonous and statistically adequate representation of the given time series, are computed by minimizing conditional sum of square of errors so that

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(11)

$$SSE(\underline{\varphi},\underline{\theta}) = \sum_{t=1}^{n} [u_t \mid \underline{\varphi},\underline{\theta},\underline{\omega}]^2$$
(12)

where

$$u_{t} = \omega_{t} - \sum_{i=1}^{p} \varphi_{i} \omega_{t-i} + \sum_{j=1}^{q} \theta_{j} u_{t-j}$$
(13)

The minimization of sum of squares function is accomplished on a computer by using a nonlinear least squares estimation method developed by Marquardt (19). A back forecasting procedure is proposed to approximate unconditional sum of squares of erros (20).

2.1.4. Diagnostic Checking of ARIMA Models

After having the estimated parameters of the specified ARIMA model, diagnostic checks are applied to verify the model adequacy. There are two general factors of interest for diagnostic checking.

(1) The residuals or errors left over after fitting an ARIMA model are expected to be a white noise process. Any recognizable pattern in the estimated autocorrelation function of errors could point out the model inadequacy. A lack of fit test is proposed to test whether the autocorrelation estimates $r_k(u)$ for the residuals are significantly different from zero. The statistic for this purpose is called Q statistic and it is computed as follows:

$$Q_{u} = n \frac{K}{k=1} [r_{k}(u)]^{2}$$
 (14)

where K is maximum time lag considered.

 Q_u statistic is approximately distributed as x^2 with K-p-q degrees of freedom (21), (22).

- (2) The sampling statistic of the current optimum solution are studied to see if the model is overfitted. There are two summary statistics provided :
 - (a) standard errors of the estimated parameters;

(b) correlation matrix of the estimated parameters.

High standard errors of the estimated parameters, and high correlation between parameters may be the signs of overfitting.

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2.1.5. Seasonal Models

The general seasonal ARIMA model is represented as $ARIMA(p,d,q)(P,D,Q)^S$ where

P denotes order of seasonal AR process,

D denotes degree of seasonal differencing,

Q denotes order of seasonal MA process, and

S is period of seasonality.

It can be written as

$$\varphi(B) \Phi(B) \nabla^{d} \nabla^{D} z_{t} = \Theta(B) \Theta(B) u_{t}$$

where $\varphi(B)$ and $\theta(B)$ are nonseasonal operators as previously identified, $\Phi(B)$ is seasonal AR operator of order P,

- $\Theta(B)$ is seasonal MA operator of order Q, and
- ∇^{D} is the differencing operator so that $\nabla^{D} = (1-B^{S})^{D}$.

The general seasonal ARIMA model is in multiplicative form. Thus, it can be represented as a two stage filtering operation on the process. First is the nonseasonal filtering such that

$$a_{+} = \Theta^{-1} (B) \varphi(B) \omega_{+}$$
(16)

where

 $\omega_{+} = \nabla^{d} \nabla^{D} z_{+}$

Then, applying seasonal filtering on a₊,

 $\Phi(B) a_{+} = \Theta(B) u_{+}$ (17)

which has the same meaning with Equation (15).

Although the mathematics of seasonal ARIMA processes seems sophisticated, it should be kept in mind the the fundamentals of seasonal model building are similar to nonseasonal model building, considering the fact that it takes place at the second stage of the filtering procedure. At the model identification stage, the autocorrelations and partial autocorrelations should be axamined for a seasonal pattern at spikes S lags apart.

(15)

2.2. Transfer Function Models

A transfer function model describes the dynamic response of input variables or leading indicators of a system on the output of that system. In this study, only the bivariate case is studied, namely there is one input series. After a brief explanation of transfer function models, the model building is described in terms of three main stages: identification, estimation, and diagnostic checking.

2.2.1. Fundamentals of Transfer Function Models

 $Z_{t} = V_{0} X_{t} + V_{1} X_{t-1} + \dots + N_{t}$

A very general form of the transfer function can be written as

or

$$Z_t = v(B) X_t + N_t$$

where

- z_+ denotes output series for t = 1, 2, ..., N,
- X_{+} denotes input series for t = 1, 2, ..., N,
- N_t denotes the sum of all effects of all variables
 - other than X_t, usually called <u>noise</u>,
- v(B) is called the transfer function of the system, the weights v₀, v₁,... are called impulse response function of the system.

The infinite series $v_0 + v_1 B + v_2 B^2 + \ldots$ must be convergent for the system to be stable. The stability condition implies that a finite incremental change in the input results in a finite incremental change in the output. Sometimes, there is a delay of the effect of input to output. Impulse response weights for the periods that input effect lags output are theoretically zero.

It is not practical to represent the system with a high order of transfer function. The parsimonous form of the model is represented by the ratio of two polynomials, such that

 $v(B) = \frac{\omega(B)}{\delta(B)} B^{b}$

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(18)

where

Thus, the model becomes

$$\delta(B) \ z_{t} = \omega(B) \ x_{t-b} + n_{t-b}$$

where

$$\delta(B) = 1 - \delta_1 B - \delta_2 B^2 - \dots - \delta_r B^r$$

 $\omega(B) = \omega_0 - \omega_1 B - \omega_2 B^2 - \ldots - \omega_S B^S$

 $z_t = \nabla^d Z_t$, the differenced output series,

 \boldsymbol{x}_t = $\boldsymbol{\nabla}^d$ \boldsymbol{X}_t , the differenced input series,

 n_t is the noise of the model, $n_t = \nabla^d N_t$, r,s,b are nonnegative integer constants.

An ARIMA model is applied to the noise series so that

$$\varphi(B) n_t = \Theta(B) a_t$$
 (21)

Then, the output z_t can be written in the following form

$$z_{t} = \frac{\omega(B)}{\delta(B)} x_{t-b} + \frac{\theta(B)}{\varphi(B)} u(t)$$
 (22)

which is the general form of bivariate transfer function for the transformed series.

<u>17</u>

(20)

2.2.2. Identification of Transfer Function Models

If the estimated autocorrelation and cross correlation functions of x_t and z_t series fail to damp out quickly, then a degree of differencing d is necessary to induce stationary. The stationarity assumption implies that the constituent processes x_t and z_t have constant means and variances. The estimates of cross covariance coefficients $c_{xy}(k)$ are computed by the following formula.

$$c_{xy}(k) = \frac{1}{N} \sum_{t=1}^{N-k} (x_t - \bar{x})(z_{t+k} - \bar{z}) , k=0,1,2,...,K$$
 (23)

Let $c_{xx}(k)$ and $c_{zz}(k)$ be the autocorrelations of input and output series respectively, and S_x and S_z be the estimates of σ_x and σ_z respectively. The cross correlation coefficients $r_{xy}(k)$ are estimated as follows:

$$r_{xy}(k) = \frac{c_{xy}(k)}{S_x S_y}$$
, k=0,1,2,...,K (24)

where

$$S_{x} = \sqrt{C_{xx}(0)}$$
$$S_{z} = \sqrt{C_{zz}(0)}.$$

A plot of $r_{xy}(k)$, k=0,1,...,K is called cross correlation function.

The first step at the identification stage is the differencing operation to achieve stationarity in the series.

Secondly, input series x_t and output series y_t are prewhitened. Prewhitening of the series is accomplished by fitting an ARIMA model to input series x_t , such that

 $\theta_{x}^{-1}(B) \Phi_{x}(B) x_{t} = \alpha_{t}$ (25)

which transforms the correlated input series x_t into the uncorrelated white noise series α_t . The same prewhitening operation is applied to output series z_t as follows:

$${}^{\beta}t = {}^{\theta}x^{(B)} {}^{\Phi}x^{(B)} {}^{z}t$$
(26)

Then, autocorrelations $r_{\alpha\alpha}(k)$ and $r_{\beta\beta}(k)$, and crosscorrelations $r_{\alpha\beta}(k)$ are computed so that

$$r_{\alpha\beta}(k) = \frac{c_{\alpha\beta}(k)}{S_{\alpha}S_{\beta}}$$
, $k = 0, 1, 2, ..., K$ (27)

where

 $c_{\alpha\beta}(k)$ is cross covariance coefficient at lag k,

$$S_{\alpha} = \sqrt{c_{\alpha\alpha}(0)}$$
, and
 $S_{\beta} = \sqrt{c_{\beta\beta}(0)}$.

Finally, impulse response weights estimates v_k are computed as

$$v_{k} = \frac{S_{\beta}}{S_{\alpha}} r_{\alpha\beta}(k) , \quad k = 0, 1, 2, ..., K$$
 (28)

The preliminary estimates v_k are statistically inefficient, but it can provide a rough basis for selecting appropriate operators $\delta(B)$ and $\omega(B)$ of the transfer function model.

Knowing the $\,v_{\rm k}^{}\,$ values, r,s, and b may be guessed by employing the following guidelines:

- (1) First b values of impulse response weights, vo, V1,..., vb-1, will not be significantly different from zero.
- (2) If $r \leq s$ then a further s-r+1 values, v_b , v_{b+1} ,..., v_{b+z-r} , will not show any clear pattern.
- (3) Value v_j with $j \ge b+s-r+1$ will follow a fixed pattern.

It is possible to find a rough estimate of number of impulse response weights which are significantly different from zero except first b-1 values. Let h be this estimate, then the noise series n_t can be estimated as

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$$n_t = z_t - \sum_{i=0}^{h} v_i x_{t-i}$$

or

$$n_t = z_t - v(B) x_t$$

An ARIMA model for the noise series is then specified so that

$$\varphi (B) n_{t} = \Theta(B) u_{t}$$
(30)

2.2.3. Estimation of Transfer Function Models

Having specified transfer function model for the given series, the estimates of parameters can be obtained by minimizing the conditional sum of squares of errors function:

$$SSE(b, \underline{\delta}, \underline{\phi}, \underline{\theta}) = \sum_{\substack{i=a+p+1}}^{n} [u_{t} \mid \underline{x}, \underline{z}, \underline{u}]^{2}$$
(31)

where a is the larger of r and s+b. Marquardt's nonlinear least squares estimation method is employed to solve iteratively for the best values of the parameters.

The calculation of u's is accomplished in the following way: First, the output y_t from transfer function is computed as

$$y_{t} = \delta^{-1}(B) \omega(B) x_{t-b}$$
 (32)

Secondly, having calculated y_t series, the noise series n_t can be obtained from

$$n_t = z_t - y_t \tag{33}$$

Finally, u's can be obtained from

$$u_{t} = \theta^{-1}(B) \varphi(B) n_{t}$$
(34)

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(29)

2.2.4. Diagnostic Checking of Transfer Function Model

The residuals, u_t , are assumed to be normally, independently, identically distributed having mean zero and variance σ_u^2 . If the autocorrelation function $r_{uu}(k)$ shows a clear pattern, this indicates model inadequacy. Q_{uu} statistic to test residual autocorrelations is computed as

$$Q_{uu} = (n - a - p) \sum_{k=1}^{K} r_{uu}^{2}(k)$$
 (35)

 Q_{uu} is approximately distributed as x ² with K-p-q degrees of freedom.

If the cross correlation function $r_{\alpha \mu}(k)$ shows any significant spikes, then transfer function model is suggested to be inadequate. The statistic to test cross correlations $r_{\alpha \mu}(k)$ is computed as

$$Q_{\alpha U} = (n - a - p) \sum_{k=0}^{K} r_{\alpha U}^{2} (k)$$
 (36)

 $Q_{\alpha u}$ is approximately distributed as X ² with K-r-s degrees of freedom (23).

The same summary statistics given for the univariate model checking are also provided to test overfitting.

2.3. Forecasting

The model fitted to the time series may not be the forecast function. The minimum mean square error forecasts are obtained from the difference equation form of the model. Also, probability limits of the forecasts are provided.

2.3.1. Forecasting with ARIMA Models

The exact forecast function is computed from the fitted function in Equation (5). Assuming a deterministic shift θ_0 exists in the process, it is written as

$$\Theta_0 = (1 - \varphi_1 - \varphi_2 - \dots - \varphi_n) \overline{\omega}$$
(37)

The model fitted to the series can be written as

$$(1 - \varphi_1 B - \varphi_2 B^2 - \dots - \varphi_p B^p) (1 - B)^d z_t = \Theta_0 + \Theta(B) u_t$$
 (38)

which is the final form.

By doing necessary multiplications in Equation (38), the model can be written as

$$(1 - \zeta_1 B - \zeta_2 B^2 - \dots - \zeta_{p+d} B^{p+d}) z_t = \Theta_0 + \Theta(B) u_t$$
 (39)

Finally, the forecast function can be written as

$$z_{t} = \sum_{i=1}^{p+d} \zeta_{i} z_{t-i} + \Theta_{0} + u_{t} - \sum_{i=1}^{q} \Theta_{i} u_{t-i}$$
(40)

On the other hand, the forecast function may also be represented as a linear combination of current and previous shocks u_t , u_{t-1} ,... That is

$$z_{t+\ell} = (u_{t+\ell} + \Psi_1 \ u_{t+\ell} + \dots + \Psi_{\ell-1} u_{t-1}) + (\Psi_{\ell} u_t + \Psi_{\ell+1} \ u_{t-1} + \dots) (41)$$
$$= e_t(\ell) + \hat{z}_t(\ell)$$

where

 $e_t(\ell)$ denotes the forecast error for lead time ℓ , $\hat{z}_t(\ell)$ denotes forecast made at origin t for lead time ℓ .

From Equation (41) $e_t(\ell)$ can be written as

$$e_t = u_{t+\ell} + \Psi_1 u_{t+\ell-1} + \cdots + \Psi_{\ell-1} u_{t+1}$$

(42)

Since the expected value of $e_t(\ell)$, $E[e_t(\ell)]$, is zero, forecasts are unbiased. The variance of the forecast error is then obtained as

$$v(\ell) = var[e_{t}(\ell)] = (1 + \Psi_{1}^{2} + \Psi_{2}^{2} + \dots + \Psi_{\ell-1}^{2}) \sigma_{u}^{2}$$
(43)

The upper and lower probability limits are computed as

$$z_{t+\ell}(\ell) = \hat{z}_t(\ell) \pm c \sqrt{v(\ell)}$$
(44)

where c is the value of standard normal distribution depending on the probibility that a future value lies in the interval.

2.3.2. Forecasting with Bivariate Transfer Function Models

The model fitted to the series in Equation (22) may be reorganized to obtain the following form of the model:

$$\varphi(B) \ \delta(B) \ \nabla^{d} z_{t} = \varphi(B) \ \omega(B) \ \nabla^{d} x_{t-b} + \Theta(B) \ \delta(B) \ u_{t}$$
(45)

By doing necessary operations in Equation (45), the final forecast function becomes

$$z_{t} = \sum_{i=1}^{p+r+d} \Delta_{i} z_{t-i} + \Omega_{0} x_{t-b} - \sum_{j=1}^{p+s+d} \Omega_{i} x_{t-b-i} + \sum_{k=1}^{q+r} \Theta_{i} u_{t-i}$$
(46)

The variance of the forecast error for lead time ℓ is given by

$$v(\ell) = \sigma_{\alpha}^{2} \sum_{j=b}^{\ell-1} v_{j}^{2} + \sigma_{u}^{2} \sum_{j=0}^{\ell-1} \Psi_{j}^{2}$$
(47)

The probability limits are then computed in the same way as in Equation (44).

III. METHODOLOGY FOR THE IMPLEMENTATION

The three basic analytical stages of Box-Jenkins approach comprise several procedural steps in the computer implementation of the method. In this chapter, the analysis is presented in terms of procedural steps. Illustrative outlines are given in order to demonstrate how to operate the computer model during the application of the method.

The computer model does not provide automatic model selection, therefore external intervention is necessary at decision points during the execution of the program. The user should strictly follow the procedural steps, which are given in the following sections, during model building and forecasting.

3.1. Methodology for ARIMA Models

The methodology of using ARIMA processes includes six procedural steps. The complete analysis is outlined in Figure 3.1.

<u>Step 1</u>. The user has facilities such as to enter new data series, to access existing data series, or to update existing data series. The data series are saved for later use, the maintenance of data files is the responsibility of the user.

<u>Step 2</u>. The analysis of data is performed in this step. The computer software will provide autocorrelation function and partial autocorrelation function. It is recommended to apply an iterative process to achieve a stationary mean in the series by providing values of d, S, and D. A subclass of models, which, are candidates for a good representation of the process under study, is determined in this step.

<u>Step 3</u>. The user will specify p, q and P, Q parameters of ARIMA(p,d,q) $(P,D,Q)^{S}$ model. If a deterministic shift is recognized in the series, the control parameter M is set to one. The computer program will provide starting values of the parameter $\underline{\phi}$, $\underline{\theta}$, $\underline{\theta}$, $\underline{\theta}$, θ_{o} for the optimization.

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<u>Step 4</u>. Having starting points of the parameters, the computer will carry out the minimization procedure to fit the chosen model to the given data series. The stationarity of AR process, and invertibility of MA process should be checked on the resulting values of the parameters.

An ARMA(0,q) process is always stationary. For an ARMA(1,q) process, the absolute value of φ_1 must be less than one. The stationarity requirement for an ARMA(2,q) process consists of three conditions:

$$|\varphi_{2}| < 1$$

 $\varphi_{2} - \varphi_{1} < 1$ (48)
 $\varphi_{2} + \varphi_{1} < 1$

The stationary conditions are mathematically complicated for p > 2, at least, a rough check may be made

$$\varphi_1 + \varphi_2 + \dots + \varphi_p < 1 , p > 2$$
 (49)

The invertibility conditions of MA processes are similar to the stationarity requirements of AR processes. An ARMA(p,0) process is always invertible. An ARMA(p,1) process is invertible if $|\theta_1| < 1$. The invertibility requirements of ARMA(p,2) are a set of three conditions:

$$|\theta_2| < 1$$

$$\theta_2 - \theta_1 < 1$$
(50)

$$\theta_2 + \theta_1 < 1$$

If q > 2, then a rough check might be

 $\theta_1 + \theta_2 + \ldots + \theta_q < 1 , q 2$ (51)

The relevant statistics for diagnostic checking are provided in this step.

<u>Step 5</u>. If the current fitted model is decided to forecast future values of the series, the exact forecast function is computed by unscrambling the operators which are employed during model building stages. The model is saved for later use.

<u>Step 6</u>. The forecasts of future values are provided with the upper and lower probability limits. Sum of errors, SOE, is computed for <u>tracking</u> signal test (24). It is a measure of accuracy of the forecasting model.

Let current time be N, the forecast error at time t be $u_t = z_t - \hat{z}_t$, thus

$$SOE = \sum_{t=1}^{N} u_{t}$$
(52)

If forecast is unbiased, $E[u_t]$ will be zero, that is to say forecast error u_t is considered a random variable having mean zero. A significant departure of SOE from zero may indicate the inadequacy of present model.

The printouts are also designed for each step except step one. They, include detailed information about the process.

3.2. Methodology for Transfer Function Models

The analysis of transfer function method for bivariate time series resembles to that of ARIMA method in many respects. It includes eight procedural steps. The complete analysis is outlined in Figure 3.2.

<u>Step 1</u>. The contents of this step is the same with the univariate case, except there are two data series.

<u>Step 2</u>. Prewhitening of input series is accomplished by employing steps two through five of ARIMA case on input series. It aims to remove systematic variation in the series.

<u>Step 3</u>. Prewhitening of output series is accomplished by using the model built in the last step, step 2. The cross correlation function and direct estimate of impulse response function are provided. Initial estimate of noise series n_t are computed depending on direct estimates of impulse response weights. The user should specify candidate models for the representation of the given data series.

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Figure 3.2 Outline of Box-Jenkins Transfer Function Analysis Applied to Forecasting

<u>Step 4</u>. Starting values of the specified transfer function model parameters $\underline{\delta}$ and $\underline{\omega}$ are computed in this step.

<u>Step 5</u>. This step basically includes steps two and three of ARIMA case which are applied to noise series n_+ .

<u>Step 6</u>. The minimization procedure is executed to estimate values of parameters $\underline{\delta}, \underline{\omega}, \underline{\varphi}$, and $\underline{\theta}$ of the specified transfer function model. The summary statistic is provided for diagnostic checking. The stability of the transfer function model should be checked (25). The stability of the first order model requires absolute value of δ_1 must be less than one, and for the second order model, the parameters δ_1, δ_2 satisfy

$$\delta_2 + \delta_1 < 1$$

$$\delta_2 - \delta_1 < 1$$

$$-1 < \delta_2 < 1$$

<u>Step 7</u>. The forecast function is computed by unscrambling the operators. The parameter values of the forecast function are saved for making the forecasts later.

<u>Step 8</u>. This step is mainly the same with the Step 6 of ARIMA methodology.

The printouts, which include detailed information about the process, are provided in steps two through eight.

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(53)

IV. CONCLUSION

The Box-Jenkins time series analysis is theoreticaly and statistically very powerful approach for making the forecast of the future. The forecasting application of the approach provides minimum mean square error forecast that there is no other forecast that conditioned only the same history of the series which can produce sum of squares of errors will on the average be smaller. Contrary to this fact there are a few difficulties hindering widespread adoption of the approach as a basis for forecasting.

The Box-Jenkins methods are not suited to the handling of large number of series if the forecasts are needed quickly and cheaply. The methods perform well on situations where good data exist over a reasonable period of time, and where high accuracy of forecasts are necessary for an important planning issue. In certain cases, the method will be either uneconomic to perform, or inferior to some other time series methods, because it is not suitable for short series, and for longer forecast lead times.

Application of Box-Jenkins time series analysis to operational forecasting requires a substantial amount of computing at each stage of model building and forecasting procedures. A computer program is designed with the objective of offering maximum flexibility and ease of use to nonprogramming users. The program can be operated in a fairly mechanical manner, but time series model building considerably requires significant human judgement. The user of the computer program should understand both the methodology suggested and each aspect of the computer output. The computer model for Box-Jenkins approach can not be applied meaningfully unless the methodology and the underlying principles are well understood. The level of understanding of the various steps in the procedures of applying Box-Jenkins method will be a major determinant in its use. In addition, there is a subtantial need for experience and some trial and error in successfully applying the method. In this study, Marquardt's nonlinear least squares estimation algorithm is applied for iteratively minimizing sum of squares of residuals. A study of further development of the computer model could be designing an efficient nonlinear optimization method to maximize likelihood function instead of using Marquardt's algorithm. It is generally expected that maximum likelihood estimates of the parameters on the average will be more efficient than least squares estimates of the parameters.

This study may also be further extended in two dimensions:

- (1) In this study, it has been dealt only with the bivariate case of transfer function analysis. The forecasting version of multivariate transfer function analysis may be applied for making the use of full capabilities of transfer function method.
- (2) Box-Jenkins time series analysis can be applied for process control. The objective to design control schemes is to minimize overall measure of error at the targeted output (26). This model may be developed to design discrete control schemes.

The application of Box-Jenkins time series analysis to forecasting has considerably lagged behind its theoretical formulation and verification. This study should be granted successful if it provides any impetus on the application of Box-Jenkins time series analysis to forecasting, and if it encourages further study on the subject.

APPENDICES

APPENDICES

Appendix	A	COLL	ECTI	ON	0F	TIM	E SER	IES	USED	FOR	APPL	ICAT	IONS	
A.1.	S	eries	1:	Мс	ontl	hly	Sales	of	Glass	Prod	uct	A ·		
A.2.	S	eries	2 :	Мс	ntl	nly	Sales	of	Glass	s Pro	duct	В	•	
A.3.	S	eries	3 :	Мс	ntl	hly	Sales	of	Glas	s Pro	duct	С		
A.4.	S	eries	4:	Мс	ntl	nly	Sales	of	Glass	s Pro	duct	D		
Appendix	В	APPL	ICAT	101	IS	·.								
B.1.	A	pplica	tion	1	•	Sto for	chast Seri	ic M es f	Model 1	Buil	ding	and	Foreca	sting
B.2.	А	pplica	tion	2	:	Sto	chast	ic M	Model	Buil	ding	and	Foreca	sting
		• •				for	Seri	es a	2		•			
B.3.	A	pplica	tion	3	:	Sto	chast	ic M	Model	Buil	ding	for	Series	3
B.4.	A	pplica	tion	4	:	Sto	chast	ic M	Model	Buil	ding	for	Series	4

APPENDIX A

COLLECTION OF TIME SERIES USED FOR APPLICATIONS

A.1 Series 1 : Monthly Sales of Glass Product A

Period	Observation	Period	Observation	Period	Observation
1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 9 20 21 22 23 24 25 26 27 28	3829 2592 3498 3028 3414 3104 4627 2698 2422 4196 4627 4762 3295 2521 2988 2878 2878 2846 4809 3128 2495 2906 3453 4905 5302 3384 3014 2350 2295 2655	47 48 49 50 51 52 53 54 55 56 57 58 59 60 61 62 63 64 65 66 67 68 69 70 71 72 73 74 75	5998 4860 3953 4860 3366 2424 3678 2729 3800 3677 3922 3846 3073 3810 3890 3013 2406 2978 3083 4165 4722 4065 2704 2392 4039 5487 5010 4022 6133	93 94 95 96 97 98 99 100 101 102 103 104 105 106 107 108 109 110 111 112 113 114 115 116 117 118 119 120	5743 5729 9142 9440 1175 2142 3437 6757 9124 4811 2719 6662 3504 6318 5327 3092 2086 4266 3262 5155 4796 3680 7197 7398 14374 8265 6181 4939
29 30 31 32 33 34 35 36 37 38 39 40 41 42 43 44 45 46	2613 3169 3845 3983 2906 2471 2630 2665 2459 3876 3623 3559 4143 3317 3586 2976 4552	76 77 78 79 80 81 82 83 84 85 86 87 88 89 90 91 92	4818 3589 3648 3417 4247 4639 6091 6005 4816 2371 3302 4477 4598 8998 1534 6186 5987		

A.2 Series 2 : Monthly Sales of Glass Product B

Period	Observation	Period	Observation	Period	Observation
$ \begin{array}{c} 1\\2\\3\\4\\5\\6\\7\\8\\9\\10\\11\\12\\13\\14\\15\\16\\17\\18\\19\\20\\21\\22\\23\\24\\25\\26\\27\\28\\29\\30\\31\\32\\33\\34\\35\\36\\37\\38\\39\\40\end{array} $	1728 1306 1528 1491 1489 1600 1291 1164 1147 1150 1884 1627 2163 1485 1732 1109 1362 1839 1395 877 1035 1204 1383 1639 1793 1844 1330 1696 1788 1871 1980 1215 729 1240 1572 2254 2235 2031 1592 2103	41 42 43 44 45 46 47 48 95 51 52 53 55 55 57 58 960 61 62 63 64 566 67 89 70 71 72 73 74 75 77 78 980	$\begin{array}{c} 1830\\ 1662\\ 2344\\ 1502\\ 1400\\ 1825\\ 1702\\ 1931\\ 1915\\ 2453\\ 1646\\ 1217\\ 1029\\ 1700\\ 1400\\ 1300\\ 1400\\ 1300\\ 1400\\ 1301\\ 1470\\ 2434\\ 1451\\ 1642\\ 1404\\ 1769\\ 1484\\ 1683\\ 1243\\ 2519\\ 1906\\ 1418\\ 1570\\ 1437\\ 1611\\ 922\\ 2405\\ 2410\\ 2113\\ 1816\\ 1571\\ 1402 \end{array}$	81 82 83 84 85 86 87 88 89 90 91 92 93 94 95 96 97 98 99 100 101 102 103 104 105 106 107 108 109 110 111 112 113 114 115 116 117 118 119 120	$\begin{array}{c} 1641\\ 1716\\ 2214\\ 1386\\ 1066\\ 1247\\ 1513\\ 1783\\ 3638\\ 738\\ 818\\ 1647\\ 1250\\ 1704\\ 3355\\ 2335\\ 540\\ 928\\ 1519\\ 2262\\ 2786\\ 1303\\ 1213\\ 2447\\ 1442\\ 1676\\ 1718\\ 863\\ 1064\\ 1282\\ 1511\\ 2037\\ 1428\\ 1153\\ 1670\\ 1539\\ 2935\\ 1974\\ 2241\\ 1469\\ \end{array}$

A.3 Series 3 : Monthly Sales of Glass Product C

. .

Period	Observation	Period	Observation	Period	Observation
1 2 3 4 5 6 7 8 9 10 11 23 4 5 6 7 8 9 10 11 23 24 25 6 27 28 9 30 31 23 34 35 36	2790 2143 2809 2513 2910 3962 2100 1960 2031 3305 3358 3365 3060 4102 3492 3072 3198 3509 2763 2489 2606 3242 2429 2880 2732 2973 2453 2354 2966 3188 2928 2490 2589 2827 2811 3111	37 38 39 40 41 42 43 44 45 46 47 48 49 50 51 53 54 55 56 57 58 9 60 61 62 63 64 56 67 68 970 71 72	3395 4071 2693 2869 3022 5868 4925 4762 3385 3199 3222 2254 2813 3828 2739 2855 3193 4947 4050 3358 3749 3308 3576 5400 5015 2764 2326 2700 2397 3366 2683 2911 4547 3139 3164 2956	$\begin{array}{c} 73\\74\\75\\76\\77\\78\\79\\80\\81\\82\\83\\84\\85\\86\\87\\88\\89\\90\\91\\92\\93\\94\\95\\96\\97\\98\\99\\100\\101\\102\\103\\104\\105\\106\\107\\108\end{array}$	3027 2900 3737 3476 3385 3345 2949 3352 3939 4251 5716 3784 2234 3207 4928 5276 8890 4849 3788 9384 6441 4169 7602 6420 1788 2686 4193 9114 8386 4404 3925 7212 4167 5577 5853 4015

A.4	Se
A.4	- 26

eries 4 : Monthly Sales of Glass Product D

Period	Observation	Period	Observation
1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 9 20 21 22 23 24 25 26 27 28 9 30	$\begin{array}{c} 329\\ 308\\ 262\\ 242\\ 551\\ 200\\ 284\\ 35\\ 140\\ 584\\ 936\\ 594\\ 314\\ 293\\ 258\\ 381\\ 406\\ 618\\ 391\\ 619\\ 709\\ 187\\ 127\\ 52\\ 65\\ 10\\ 29\\ 30\\ 22\\ 31\\ \end{array}$	31 32 33 34 35 36 37 38 39 40 41 42 43 44 45 46 47 48 49 50 51 52 53 54 55 56 57 58 59 60	25 33 28 28 33 29 225 252 301 512 957 218 400 625 442 473 550 405 155 218 223 343 381 278 649 586 778 437 313 715

APPENDIX B APPLICATIONS

In this appendix four univariate applications of the Box-Jenkins approach are presented for the time series given in Appendix A.

B.1. Application 1 : Stochastic Model Building and Forecasting for Series 1

B.1.1. Model Identification

Number of observations, N : 108 Mean of the observations, $\bar{z} = 3975$ Degree of nonseasonal differencing, d = 1 Period of seasonality, S = 6 Degree of seasonal differencing, D = 1 Number of observations in the differenced series, n = 101 Mean of the differenced series, $\bar{\omega} = -24.9$ Variance of the differenced series, $\bar{\sigma}_{\omega}^2 = 6.078 \times 10^6$

****** Autocorrelations of the Series ****** Approximate standard error : 9.950E-02

Lag			Autocor.
0		! ************************************	1.00E+00
1	******		-3.9E-01
2	**		-9.0E-02
3	**		-7.9E-02
4		***	9.44E-02
5		*****	2.15E-01
6	*****		-5.4E-01
7		*****	2.36E-01
8		*	6.09E-02
9		**	7.48E-02
10	*		-5.6E-02
11	**		-7.3E-02
12	e a construction de la construction de la construction de la construction de la construction de la construction	***	1.05E-01
13			-2.9E-02
14			-1.8E-02
15	**		-6.8E-02
16		*	5.84E-02
17	· · · · · · · · · · · · · · · · · · ·	*	5.16E-02
18	*		-3.1E-02

k A	* Partial Autocorrelations of the Se pproximate standard error : 9 950E-0	ries ** 2	
Lag			Partials
1	******************	1	-3.9E-01
2	***********		-2.9E-01
3	*************		-3.1E-01
4	**********		-1.7E-01
5		<u> </u> **********	1.95E-01
6	**************************************	1	-4.8E-01
7	**********		-2.4E-01
8	*********	1 1	-1.5E-01
9	********		-1.5E-01
10	****		-5.8E-02
11	•	****	1.001-01
12	********	1	-2.2E-01
13	****		-6.1E-02
14		1	4.35E-03
15	*****		-1.0E-01
16			1.15E-02
17		******	1.40E-01
18	****	1 · · · · · · · · · · · · · · · · · · ·	-5.8E-02

 $ARIMA(1,1,1)(0,1,1)^6$ model is specified for Series 1. It can be written as follows:

$$(1 - \varphi_1 B) \omega_t = \Theta_0 + (1 - \Theta_1 B^6)(1 - \Theta_1 B) u_t$$

or

 $(1 - \varphi_1 B)(1 - B)(1 - B^6) Z_t = \Theta_0 + (1 - \Theta_1 B^6)(1 - \Theta_1 B) u_t$

B.1.2. Model Estimation and Diagnostic Checking Results

(a) Preliminary estimates:

$$\Theta_{0} = -19.24$$
 $\varphi_{1} = 0.2280$
 $\Theta_{1} = 0.5913$
 $\Theta_{1} = 0.2398$

SSE = 3.743 x 10⁸

 $\sigma_{11}^{2} = 4.377 \times 10^{6}$

(b) Final estimates:

Parameter	Value	Standard Error
θ	-5, 1219	3,8380
φ1	0.1751	0.1180
θ	0.9060	0.0534
θ1	0.8039	0.0796

SSE = 2.335×10^8 $\hat{\sigma}_u^2$ = 2.407 x 10⁶

(c) Residuals:

** Residual autocorrelations ** Aproximate standard error : 9.950E-02

Lag			-Autocor.
0	, i	*****	1.00E+00
1			1,38E-02
2	**	1 Construction of the second secon	-6.1E-02
3	*****	1 A second se	-1.3E-01
4	- *	I show the second se	-3.3E-02
5	•	 ***	8.26E-02
6			-4.4E-03
7		****	1.05E-01
8		 ***	9.47E-02
9		 **	5.66E-02
10	*		-3.1E-02
11			6.72E-03
12		 *	2.59E-02
13] *	3.51E-02
14	*		-3.2E-02
15		la de la constante de la constante de la constante de la constante de la constante de la constante de la const La constante de la constante de la constante de la constante de la constante de la constante de la constante de	-6.9E-04
16	**		-5.8E-02
17		*	4.02E-02
18		¦ *	3.05E-02

Chi-square statistic for residual autocorrelations, Q = 6.311, Degrees of freedom = 14

(d) Correlation matrix of the estimated parameters:

	θο	φ1	θ	θ	
θο	1.000	0.011	0.057	-0.088	
φı	0.011	1.000	0.520	-0.170	
θι	0.057	0.520	1.000	-0.360	
θ	-0.088	-0.170	-0.360	1.000	

(e) Forecast function:

 $z_{N+\ell} = -5.122 + 1.175 z_{N+\ell-1} - 0.175 z_{N+\ell-2} + z_{N+\ell-6} - 1.180 z_{N+\ell-7} + 0.175 z_{N+\ell-8} - 0.906 u_{N+\ell-1} - 0.804 u_{N+\ell-6} + 0.728 u_{N+\ell-7}$

for l = 1, 2, ...,

B.1.3. Forecasting

Forecast base time : 108 Forecast lead time : 12

			90% Probability Limits	
Time	Actual	Forecast	Lower limit	Upper limit
109	2.086000E+03	3.282533E+03	7.224050E+02	5.842660E+03
110	4.266000E+03	4.432797E+03	1.781555E+03	7.084040E+03
111	3.262000E+03	4.200969E+03	1.525212E+03	6.876727E+03
112	5.155000E+03	5.243843E+03	2.550875E+03	7.936811E+03
113	4.796000E+03	6.388686E+03	3.679726E+03	9.097646E+03
114	3.680000E+03	4.543673E+03	1.819004E+03	7.268342E+03
115	7.197000E+03	3.431161E+03	5.931711E+02	6.269151E+03
116	7.398000E+03	4.348077E+03	1.478152E+03	7.218002E+03
117	1.437400E+04	4.070256E+03	1.177509E+03	6.963004E+03
118	8.265000E+03	5.099952E+03	2.185939E+03	8.013966E+03
119	6.181000E+03	6.237366E+03	3.302477E+03	9.172255E+03
120	4.939000E+03	4.385930E+03	1.430353E+03	7.341506E+03

B.2. Application 2 : Stochastic Model Building and Forecasting for Series 2

B.2.1. Model Identification

Number of observations, N : 108 Mean of the observations, \bar{z} = 1629

Degree of nonseasonal differencing, d = 1 Period of seasonability, S = 6 Degree of seasonal differencing, D = 1 Number of observations in the differenced series, n = 101 Mean of the differenced series, $\bar{\omega} = -0.03$ Variance of the differenced series, $\hat{\sigma}_{\omega}^2 = 5.523 \times 10^5$

****** Autocorrelations of the Series ****** Approximate standard error : 9.950E-02

Lag			Autocor.
0		*****	1.00E+00
1	*******		-3.3E-01
2	*****		-2.0E-01
3	and the second second second second second second second second second second second second second second second		1.69E-02
4	•	****	1.42E-01
5		 *	5.72E-02
6	*****		-4.5E-01
7		*****	2.36E-01
8		*	5.12E-02
9	*		-4.2E-02
10	*		-3.0E-02
11		*	5.41E-02
12		*	5.63E-02
13	*		-4.7E-02
14		*	5.56E-02
15	· ***		-1.1E-01
16		***	1.06E-01
17	*		-5.6E-02
18			-2:7E-03

** Ar	<pre>k Partial Autocorrelations of the Series pproximate standard error : 9.950E-02</pre>	**	
Lag			Partials
. 1	*****		-3.3E-01
2	*****	• • •	-3.5E-01
3	****		-2.4E-01
4	**		-3.1E-02
5		*******	1.07E-01
6	*****		-4.3E-01
7	*********	•	-1.4E-01
8	*********	а. А	-2.0E-01
g	***********	1.	-1.8E-01
10	******		-1.0E-01
11	*		-1.7E-02
12	*****		-2.1E-01
19	****		-4.7E-02
11		**	2.42E-02
15	*****	1	-1 7E -01
10	<u>ጥ ጥ ጥ ጥ ጥ ጥ ጥ ጥ ጥ ጥ ጥ ጥ ጥ ጥ ጥ ጥ ጥ ጥ ጥ </u>	 *****	6.97E-02
10		[\ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \	2 41F - 02
17	Ψ.Ψ.	ነ ^{ጥጥ} 	_2 GF_02
18	**	i	2.05-02

 $ARIMA(1,1,1)(0,1,1)^6$ model is specified for Series 2. It can be written as follows:

$$(1 - \varphi_1 B)\omega_+ = (1 - \Theta_1 B^6)(1 - \Theta_1 B) u_+$$

or

 $(1 - \varphi_1 B)(1 - B^6)(1 - B) Z_t = (1 - \Theta_1 B^6)(1 - \Theta_1 B) u_t$

B.2.2. Model Estimation and Diagnostic Checking Results

(a) Preliminary estimates:

$$\varphi_1 = 0.6116$$

 $\theta_1 = 0.3907$
 $\theta_1 = 0.4909$
SSE = 5.022 x 10⁷
 $\hat{\sigma}_{11}^2 = 5.415 \times 10^5$

(b) Final estimates:

Parameter	Value	Standard Error
φ1	0.1964	0.1283
θι	0.8632	0.0702
θ	0.6167	0.088

SSE = 2.748×10^7 $\hat{\sigma}_{u}^2$ = 2.804×10^5

(c) Residuals:

** Residual autocorrelations **
Aproximate standard error : 9.950E-02

Lag			-Autocor.
0	•	***************************************	1.00E+00
1			7.11E-03
2 ·	****		-1.1E-01
3	*		-2.7E-02
4			5.06E-03
5	***		-8.7E-02
6	*		-3.6E-02
7		 ***	6.97E-02
8		l* Constant of the second second second second second second second second second second second second second s	3.05E-02
- 9			-1.1E-02
10	· · · · ·		1.55E-02
11		¦**	5.12E-02
12		***	7.42E-02
13		 **	5.84E-02
14	. 1	 **	5.35E-02
15	*		-4.0E-02
16	*		-2.5E-02
17	· · ·	Figure 1. A second s Second s Second se	7.36E-04
18	*		-3.7E-02

Chi-square statistic for residual autocorrelations, Q = 4.612, Degrees of freedom = 15

(d) Correlation matrix of the estimated parameters:

	φ1	θι	θι	
 Ψ1	1.000	0.630	-0.230	
θι	0.630	1.000	-0.350	
θ1	-0.230	-0.350	1.000	

(e) Forecast function:

$$z_{N+} = 1.196 z_{N+\ell-1} - 0.196 z_{N+\ell-2} + z_{N+\ell-6} - 1.200 z_{N+\ell-7} + 0.196 z_{N+\ell-8} - 0.863 u_{N+\ell-1} - 0.617 u_{N+\ell-6} + 0.532 u_{N+\ell-7}$$

B.1.3. Forecasting

Forecast	base	time	:	108
Forecast	lead	time	:	12

Timo	A atual	Trans as at	90% Probabi	lity Limits
1106	ACCUAL	forecast	Lower limit	Upper limit
109	1.064000E+03	7.529725E+02	-1.20757E+02	1.626702E+03
110	1.282000E+03	1.526588E+03	6.056317E+02	2.447545E+03
111	1.511000E+03	1.305219E+03	3.674625E+02	2.242975E+03
112	2.037000E+03	1.663040E+03	7.126843E+02	2.613395E+03
113	1.428000E+03	2.228076E+03	1.265984E+03	3.190169E+03
114	1.153000E+03	1.096733E+03	1.231777E+02	2.070288E+03
115	1.670000E+03	7.425469E+02	-3.44556E+02	1.829650E+03
116	1.539000E+03	1.468209E+03	3.503645E+02	2.586054E+03
117	2.935000E+03	1.237422E+03	9.880754E+01	2.376036E+03
118	1.974000E+03	1.593393E+03	4.359598E+02	2.750826E+03
119	2.241000E+03	2.158067E+03	9.824136E+02	3.333719E+03
120	1.469000E+03	1.026651E+03	-1.66886E+02	2.220189E+03

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B.3 Application 3 : Stochastic Model Building for Series 3

B.3.1. Model Identification

```
Number of observations, N : 108
Mean of the observations, \bar{z} = 3716
```

Degree of nonseasonal differencing, D: 1 Period of seasonality, S = 6 Degree of seasonal differencing, D = 1 Number of observations in the differenced series, n = 101 Mean of the differenced series, $\bar{\omega} = 2.980$ Variance of the differenced series, $\hat{\sigma}_{\omega}^2 = 3.569 \times 10^6$

** Autocorrelations of the Series ** Approximate standard error : 9.950E-02

Jog			Autocor.
n Dag		****	1.00E+00
1	**		-7.3E-02
$\overline{2}$	*****		-6.3E-01
3		*	4.53E-02
4		*****	4.97E-01
5	**		-7.0E-02
6	*****		-5.7E-01
7		*	5.87E-02
8		*****	3.66E-01
ğ			-1.9E-02
10	*****		-2.5E-01
11	5.	*	6.38E-02
12		****	1.89E-01
13	*		-3.9E-02
14	**		-7.1E-02
15	*		-4.0E-02
16		*	6.19E-02
17			3.89E-03
18	*		-3.3E-02

×× Aj	* Partial Autocorrelations of the Series oproximate standard error : 9.950E-02	ጙ ጞ	•
Lag			Partials
1	****		-7.3E-02
2	******		-6.4E-01
3	******		-1.3E-01
4		*******	1.52E-01
5			-8.7E-03
6	**********		-3.8E-01
7	*******		-1.7E-01
8	*********	1	-3.0E-01
9 -	*****		-1.0E-01
10		1	-1.2E-03
11		*	2.39E-02
12	*********	1 1	-1.9E-01
13	*******		-1.5E-01
14	**	1	-3.2E-02
15	******	t	-1.3E-01
16		¦ *	2.02E-02
17	**	1	-4.7E-02
18	****	t 1	-9.0E-02

ARIMA(2,1,0)(1,1,0)⁶ model is specified for Series 3. It can be written as follows:

 $(1 - \phi_1 B^6)(1 - \phi_1 B - \phi_2 B^2) \omega_t = u_t$

or

 $(1 - \Phi_1 B^6)(1 - \varphi_1 B - \varphi_1 B^2)(1 - B^6)(1 - B) z_t = u_t$

B.3.2. Model Estimation and Diagnostic Checking Results

(a) Preliminary estimates:

 $\varphi_{1} = -0.1190$ $\varphi_{2} = -0.6350$ $\Phi_{1} = -0.0885$ SSE = 1.950 x 10⁸ $\hat{\sigma}_{u}^{2} = 2.074 \times 10^{6}$

(b) Final estimates:

Parameter	Value	Standard Error
φ1	-0.322	0.0881
φ2	-0.533	0.0869
Φ_1	-0.633	0.0910

SSE = 1.497×10^8 $\hat{\sigma}_u^2 = 1.528 \times 10^6$

** Residual autocorrelations **

Aproximate standard error : 9.950E-02

Lag			-Autocor.
0		*****	1.00E+00
1	**		-5.5E-02
2	*		-3.0E-02
3	****		-1.2E-01
4			-1.5E-02
5	*****		-1.4E-01
6	**		-4.9E-02
7	****		-1 [,] .2E-01
8		*****	1.59E-01
9		**	5.30E-02
10	****		-9.5E-02
11		\ *	3.57E-02
12	**		-4.6E-02
13		¦ *	2.64E-02
14		**	4.84E-02
15	***	•	-9.1E-02
16			-8.3E-03
17		****	9.74E-02
18	•	*	2.73E-02

Chi-square statistic for residual autocorrelations, Q = 11.879, Degrees of freedom = 15

(d) Correlation matrix of the estimated parameters:

	φ1	φ ₂	Φ,	
φ1	1.000	0.170	0.230	
Ψ₂	0.170	1.00	-0.120	
Φ_1	0.230	-0.120	1.000	

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(e) Forecast function:

 $z_{N+\ell} = 0.678 \ z_{N+\ell-1} - 0.211 \ z_{N+\ell-2} + 0.533 \ z_{N+\ell-3} + 0.367 \ z_{N+\ell-6}$ - 0.249 $z_{N+\ell-7}$ + 0.078 $z_{N+\ell-8}$ - 0.195 $z_{N+\ell-9}$ + 0.633 $z_{N+\ell-12}$ - 0.403 $z_{N+\ell-13}$ + 0.134 $z_{N+\ell-14}$ - 0.338 $z_{N+\ell-15}$

l = 1,2,... for

B.4. Application 4 : Stochastic Model Building for Series 4

B.4.1. Model Identification

Number of observations, N : 60 Mean of the observations, \overline{z} = 654

Degree of nonseasonal differencing, d = 1 Period of seasonality, S = 1 Degree of seasonal differencing, D = 0 Number of observations in the differenced series, n = 59 Mean of the differenced series, $\bar{\omega} = 6.54$ Variance of the differenced series, $\hat{\sigma}_{\omega}^2 = 4.77 \times 10^4$

** Autocorrelations of the Series ** Approximate standard error : 1.302E-01

1.20			Autocor.
0		*****	1.00E+00
ĩ	******		-2.1E-01
5	******		-2.0E-01
2			7.88E-03
3			-6.2E-02
4	ሳጥ የ ህ		-2.6E-02
. ວ	т	i የ	4.94E-02
Б.		ј ጥጥ 1 ለታቀቀ	8.57E-02
1			-1.5E-01
8	*****		8 568-02
· 9		* * *	6 19F-02
10		; **	5.08F-04
11		l de la seguidad de la constante de la constante de la seguidad de la seguidad de la seguidad de la seguidad de I	-1 2E-01
12	*****		-1.3E-01
13	**		-5.3E-02
14		*	4.19E-02
15	****		-1.4E-01
16		*****	2.34E-01
17	**		-5.4E-02
18	*****		-1.6E-0 <u>1</u>

*	* Partial Autocorrelations of the Serie	es **	
Aj Lag	pproximate standard error : 1.302E-01		Partials
1	*****		-2.1E-01
2	******		-2.5E-01
З	********		-1.1E-01
4	**********		-1.6E-01
5	*********		-1.2E-01
6	*****		-5.0E-02
7		*****	5.17E-02
8	***********		-1.4E-01
9		****	4.25E-02
10		*****	4.82E-02
11		****	8.04E-02
12	*********	t i se se se se se se se se se se se se se	-9.9E-02
13	*********		-1.1E-01
14	*****	1	-4.6E-02
15	****		-2.2E-01
16		 	8.74E-02
17	********		-8.8E-02
18	***********	1 1	-1.6E-01

ARIMA(1,1,1) model is specified for Series 4. It can be written as follows:

 $(1 - \varphi_1 B) \omega_t = (1 - \Theta_1 B) u_t$

or

 $(1 - \varphi_1 B)(1 - B) z_t = (1 - \Theta_1 B) u_t$

B.4.2. Model Estimation and Diagnostic Checking Results

(a) Preliminary estimates:

 $\varphi_{1} = 0.9544$ $\varphi_{2} = -0.7560$ SSE = 2.168 x 10⁷ $\hat{\sigma}_{U}^{2} = 8.430 x 10^{5}$

(b) Final estimates:

Parameter	Value	Standard Error
φ1	0.9629	0.0221
θι	0.8032	0.0883

SSE = 3.916×10^{6}

 $\hat{\sigma}_{\rm u}^2 = 6.870 \times 10^4$

(c) Residuals:

** Residual autocorrelations **
Aproximate standard error : 1.302E-01

Lag	5		Autocor.
0		<i>************************************</i>	* 1.00E+00
1	*****		-1.4E-01
2	*****		-1.3E-01
3			1.73E-02
4	•		-1.9E-02
5			7.78E-04
6		****	1.01E-01
7		****	1.24E-01
8	*****		-1.5E-01
9		*	3.84E-02
10		 *	4.40E-02
11			1.91E-02
12.	****		-1.0E-01
13	*		-3.3E-02
14		 **	5.00E-02
15	*****		-1.4E-01
16		*****	1.89E-01
17	**		-6.1E-02
18	******		-1.8E-01
	Ch	i-square statistic for residual autocorrelations,	

Q = 11.452, Degrees of freedom = 16

(d) Correlation matrix of the estimated parameters:

	φι	θι	
φ1	1.000	0.440	
θι	0.440	1.000	

(e) Forecast function:

 $z_{N+\ell} = 1.963 \ z_{N+\ell-1} - 0.963 \ z_{N+\ell-2} - 0.803 \ u_{N+\ell-1}$

for *l* = 1,2,...

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