

**DOGUS UNIVERSITY  
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**TEST OF THE CAPITAL ASSET PRICING MODEL  
IN TURKEY**

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**Master of Science in Financial Economics Thesis**

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## **PREFACE**

Nowadays it is not easy for individual investors and portfolio managers to achieve a best decision in investments and portfolio analysis. Therefore, the role of the Capital Asset Pricing Model in investment decision-making is indeed very important. This study does not only makes a contribution to the investment management in Turkish capital market but also has widened my personal knowledge horizon with regard to the understanding of market behaviour.

I would like to express my appreciation to the staff of the Institute of Social Sciences of Dogus University for giving me the background for writing this thesis. Particularly, I would like to thank my thesis advisor Prof. Dr. Cudi Tuncer Gürsoy for his great help, Assoc.Prof. Dr. Alövsat Müslümov and Dr. Fuat Beyazıt for their valuable advices, the institute secretary Kesper Diler for her help, my husband, my family and my friends for their continued support.

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## ÖZET

Sermaye piyasasında denge koşulları açıklamaya çalışan Finansal Varlık Fiyatlandırma Modeli, sermaye piyasalarındaki risk-getiri ilişkilerini tanımlamak amacıyla araştırmacılar tarafından daha önce de test edilmiştir. Deneysel test sonuçları, modeli kuvvetle destekleyen yönde olduğu kadar aynı ölçüde yadsıyan nitelikte olmuştur. Finansal Varlık Fiyatlandırma Modeli, uygulanabilir portföy yönetiminde geniş kullanım alanı bulunduğu için, akademik araştırmaların temel tartışma konularından biri haline gelmiştir.

Bu çalışmanın amacı, Finansal Varlık Fiyatlandırma Modeli'ni Türk sermaye piyasasındaki beta-getiri ilişkisini tanımlamak, betanın menkul değerlerin riskini ölçmede ne orada başarılı olduğunu göstermektir. Finansal Varlık Fiyatlandırma Modeli testi, İstanbul Menkul Kıymetler Borsasına kayıtlı menkul kıymetlerin getirileri kullanılarak, geleneksel yöntem ve koşullu analiz yöntemi olmak üzere iki şekilde yapılmıştır. Fama ve MacBeth (1973)'in üç basamaklı geleneksel yaklaşımı, beta-getiri ilişkilerini ve risk priminin pozitif olduğunun testini amaçlamış, ancak elde edilen sonuçlar istatistiksel olarak anlamlı olmamıştır. Geleneksel test yönteminin bu sonucu vermesinin temel nedeni pazar risk priminin pozitif ve negatif olduğu dönemleri birbirinden ayırmamasıdır. Bundan dolayı, Pettengill vd'in (1995) koşullu analiz tekniği kullanılarak analiz dönemi pazar risk priminin pozitif ve negatif olduğu haftalardan oluşan iki gruba ayrılmıştır. Böylece istatistiksel olarak, beta ve getiri arasında önemli ve sağlam sonuçlara ulaşılmıştır. Ayrıca, sonuçlar, pazar risk priminin negatif olduğu dönemlerde, risk ile getiri arasında negatif ilişki olduğunu da göstermiştir. Yüksek betalı portföylerin, pazar primi pozitif olduğu dönemlerde yüksek getiri, pazar primi düşük olduğu dönemlerde düşük getiri olduğunu bulunmuştur. Türk sermaye piyasasında, beta ve getiri arasında sistematik koşula bağlı bir ilişki olduğu sonucuna varılmıştır.

Bu çalışmanın sonunda, elde edilen tüm bulgular, betanın, yatırımcılar ve portföy yöneticileri için halen kullanışlı bir risk ölçüm aracı olarak kullanılabileceğini ancak bu yapılırken pazar risk priminin pozitif ya da negatif olmasına dikkat edilmesi gerektiğini göstermektedir.

### **Anahtar Kelimeler:**

Finansal varlık fiyatlandırma modeli; Pazar risk primi; Beta ve getiri

## SUMMARY

The Capital Asset Pricing Model, as the dominating capital market equilibrium model, has been previously tested by various researches in order to define the relationship between risk and return in different capital markets. The empirical test results show the strong support as well as the evidences against the CAPM. The CAPM is the actual issue of many debates in academic research since it continues to be widely used in practical portfolio management.

This study has its aim to test the CAPM in order to determine the relationship between beta and returns in the Turkish capital market and the usefulness of beta as a single measure of a security risk. The test of the CAPM in this thesis has been conducted in both traditional and conditional way using the returns on securities listed on Istanbul Stock Exchange. Using the Fama and MacBeth (1973) three-step traditional approach which aims at testing the relationship between beta and return and whether the market risk premium is positive, has given us the result which shows the insignificant relationship between risk (measured by beta) and return. We believe that the traditional test results have been biased due to the fact that the traditional test does not take into account the condition of positive and negative market excess returns. Therefore, we have splitted the up-market weeks and down-market weeks to conduct the conditional test according to the Pettengill et al. (1995) method. The statistically significant and consistent relationships between beta and returns have been found. The results also indicate that there is an inverse relationship between risk and returns in the periods when the market excess returns are negative. It has been found that the portfolios with higher betas have higher returns when the market risk premium is positive and lower returns when the market risk premium is negative.

We have concluded that the systematic conditional relationship exists between beta and returns in the Turkish capital market. The overall evidence in this study indicates that beta is still a useful measure for risk for investors and portfolio managers to make investment decisions but the analysis should be made by means of the conditional model.

### **Keywords:**

Capital asset pricing model; Market risk premium; Beta and return



## CONTENTS

	Page No
PREFACE	i
ÖZET	ii
SUMMARY	iii
LIST OF FIGURES	v
LIST OF TABLES	vi
1. INTRODUCTION	
1.1 Purpose of the Thesis	1
1.2 Scope of the Thesis	1
1.3 Methodology of the Thesis	1
1.4 Limitations	2
2. THEORY OF THE CAPITAL ASSET PRICING MODEL	3
2.1 Main Implications and Assumptions of the CAPM	3
2.2 The Capital Market Line (CML)	6
2.3 The Characteristic Line	8
2.4 The Secutiry Market Line (SML)	11
2.5 The Relationship between the CML and the SML	13
2.6 Testing the CAPM	14
3. EARLIER EMPIRICAL STUDIES IN TESTING THE CAPM	18
4. TEST OF THE CAPM IN TURKISH CAPITAL MARKET	38
4.1 Data and Methodology	38
4.1.1 Traditional Approach	39
4.1.2 Conditional Approach	41
5. EMPIRICAL RESULTS	44
5.1 The results of the traditional (unconditional) test	44
5.2 The results of the conditional test	50
6.CONCLUSION	58
REFERENCES	60
APPENDIXES	63
Appendix I: Portfolio Formation and Estimation for Test Period 1 (1999-2000)	63
Appendix II: Portfolio Formation and Estimation for Test Period 2 (2000-2001)	64
Appendix III: Portfolio Formation and Estimation for Test Period 3 (2001-2002)	65
Appendix IV: Portfolio Formation and Estimation for Test Period 4 (2002-2003)	66
Appendix V: Portfolio Formation and Estimation for Test Period 5 (2003-2004)	67
Appendix VI: Weekly Portfolio Returns and Market Risk Premium (Test Period 1999-2000)	68
Appendix VII: Weekly Portfolio Returns and Market Risk Premium (Test Period 2000-2001)	71
Appendix VIII: Weekly Portfolio Returns and Market Risk Premium	74

(Test Period 2001-2002)	
Appendix IX: Weekly Portfolio Returns and Market Risk Premium	77
(Test Period 2002-2003)	
Appendix X: Weekly Portfolio Returns and Market Risk Premium	80
(Test Period 2003-2004)	
Appendix XI: Portfolio Returns and Betas	83
Appendix XII: Example of Betas Estimation	85
Appendix XIII: Example of T-bill Rate Adjustment	86
CURRICILIM VITAE	87

## LIST OF FIGURES

	Page No.
Figure 2.1 The Capital Market Line	7
Figure 2.2 The Characteristic Line	9
Figure 2.3 The Security Market Line	11
Figure 5.1 Traditional test results for the period of 1999-2000	47
Figure 5.2 Traditional test results for the period of 2000-2001	47
Figure 5.3 Traditional test results for the period of 2001-2002	48
Figure 5.4 Traditional test results for the period of 2002-2003	48
Figure 5.5 Traditional test results for the period of 2003-2004	49
Figure 5.6 Conditional test results for the up-market period of 1999-2000	53
Figure 5.7 Conditional test results for the down-market period of 1999-2000	53
Figure 5.8 Conditional test results for the up-market period of 2000-2001	54
Figure 5.9 Conditional test results for the down-market period of 2000-2001	54
Figure 5.10 Conditional test results for the up-market period of 2001-2002	55
Figure 5.11 Conditional test results for the down-market period of 2001-2002	55
Figure 5.12 Conditional test results for the up-market period of 2002-2003	56
Figure 5.13 Conditional test results for the down-market period of 2002-2003	56
Figure 5.14 Conditional test results for the up-market period of 2003-2004	57
Figure 5.15 Conditional test results for the down-market period of 2003-2004	57

## LIST OF TABLES

	Page No
Table 4.1 Sample periods and number of stocks	40
Table 5.1 Portfolio betas	44
Table 5.2 Summary results for the traditional test	46
Table 5.3 Summary results for the conditional test for the up-market	52
Table 5.4 Summary results for the conditional test for the down-market	52

## **1. INTRODUCTION**

### **1.1 PURPOSE OF THE THESIS**

The purpose of this study is to test the validity of the CAPM, which is widely used in portfolio management to predict the securities returns according to the relationship between risk and return, in Turkish capital market. The study is also aimed to determine the type of relationship between risk and return with regard to the market conditions.

### **1.2 SCOPE OF THE THESIS**

The scope of this thesis covers the testing of the CAPM using data from the Istanbul Stock Exchange (ISE), particularly all stock returns available on ISE meeting the data requirements of the test, for the period of 1995-2005.

### **1.3 METHODOLOGY OF THE THESIS**

There are various methodologies of testing the CAPM. However, most of them use main principles of the three-step approach of Fama and MacBeth (1973), which is one of the earliest and basic CAPM test methodologies. The basis of this method is the portfolio formation procedure according to the ranked beta estimates. Betas are estimated by running time-series regressions. The final step includes the cross-sectional regression of estimated betas and returns.

One of the recent test approaches is the Pettengill et al. (1995) methodology. This approach also based on the implications of the Fama and MacBeth method but the final step is adjusted by adding into the regression of dummy variable which allows to separate the periods of negative and positive market excess returns. This conditional approach, generally, brings to the different test results.

We have used both above-mentioned methods to test the CAPM in Turkish capital market adjusting some data according to the Turkish market conditions.

#### **1.4 LIMITATIONS**

The scope of this thesis is limited to the main assumptions of the CAPM. Our test has been conducted without relaxing of any model assumptions. The study is focused on the beta as the measure of risk and the single main variable affecting the security returns. However, the model might be broaden in the future to test the other market variables that can influence the security returns in the Turkish capital market.

## **2. THEORY OF THE CAPITAL ASSET PRICING MODEL**

### **2.1 Main Implications and Assumptions of the CAPM**

The Capital Asset Pricing Model (CAPM) was developed from the modern portfolio theory created by Markowitz in 1952. He proposed a portfolio selection technique that maximizes expected utility to a combination of portfolio return and risk. More precisely, the CAPM comes out of two developments: Markowitz portfolio theory, which is showing how to create an efficient frontier, and the theory of James Tobin, who in a 1958 paper suggested that if investors hold risky securities and are able to borrow - buying stocks on margin, or lend - buying risk-free assets, and investors do so at the same rate, then the efficient frontier is a single portfolio of risky securities plus borrowing and lending, and that dominates any other combination. Tobin's Separation Theorem suggests that investors can separate the problem into first finding that optimal combination of risky securities and then deciding whether to lend or borrow, depending on their attitude toward risk. He then showed that if there is only one portfolio plus borrowing and lending options at the same risk-free rate, efficient frontier converts to a linear relationship between risk and return called the capital market line.

These theories were later expanded independently by Treynor (1961), Sharpe (1964), Lintner (1965) and Mossin (1966) who introduce a single index model known as the CAPM. Since then, the CAPM has been the dominating capital market equilibrium model. It continues to be the most widely used model in practical portfolio management and in academic research. Its central implication is that the contribution of an asset to the variance of the market portfolio – the asset's systematic risk, or beta risk – is the correct measure of the asset's risk and the only systematic determinant of the asset's return. The model examines the type of risk-return relationship, the determination of the expected rates of return, as well as the security prices behavior. Being the most basic asset-pricing theory, CAPM tries to explain the difference between expected returns of different capital assets in term of their market risk.

As every theoretical model the CAPM has its assumptions. Being based on the Markowitz portfolio theory, the CAPM includes all its assumptions. The CAPM's basic assumptions are as follows:

- 1) *All investors are risk-averse and rational expected-utility maximizers.* It means that all investors use the Markowitz portfolio selection model, i.e. for holding a higher risk investors want to be compensated with higher return. Hence, for instance, between two assets with equal rate of return an investor will select the asset with lower risk. If it was not true and investors were not risk-averse then logically all investors would hold the single security with the highest return in order to maximise their utility preferring more return to less return. However, investors are typically risk-averse and hold diversified portfolios investing in many types of assets with different levels of risk. Every investor will choose the portfolio that best fits or maximises his or her expected utility, meaning that an investor will have the highest possible indifference curve.
- 2) *All investors select from alternative portfolios on the basis of expected return and risk.* Rational investors prefer to hold portfolios that are placed on the efficient frontier. Investor's portfolio choice depends on the investor's level of risk-aversion that is characterized by the shape of investor's risk-return utility function. A highly risk-averse investor will choose a portfolio somewhere from the lower left part of the efficient frontier with the lower risk and therefore with lower return, whereas an investor with low degree of risk-aversion will hold a portfolio from the right upper part of efficient frontier with higher return and higher risk. So, every investor will choose his or her optimal portfolio from the efficient set of portfolios.
- 3) *All investors operate on an identical single-period time horizon.* It means that this behavior includes only one investment horizon (which is the same for all investors) ignoring everything that might happen beyond it, since the shape of the return distribution generally differs for different investment horizons.
- 4) *All investors have the opportunity of unlimited borrowing and lending at a common risk-free rate.* The variance and covariance of the risk-free asset with other assets equals zero. Generally, treasury bills are taken as a proxy for the risk-free asset. However this assumption ignores the fact that zero variance does not mean that the



risk-free asset is free of all risk such as inflation. It accepts the nominal value of treasury bills.

- 5) *All investors have homogeneous expectations with regard to means, variances and covariances of asset returns.* This assumption suggests that all investors have the same economic view of the world, particularly the prospects of each asset. All investors use the same expected returns and covariances of asset returns to create the efficient frontier and the optimal portfolio. Therefore, every investor faces the same mean-variance efficient set. This assumption also requires that all investors make their decision according to the same information.
- 6) *There are no transaction costs on trade in securities and no taxes (or capital gains and dividends are taxed at the same rates).* It is assumed that there are no market imperfections such as transaction costs. This assumption makes it possible to arbitrage assets, which are mispriced by transaction costs in reality, and therefore leading prices to equilibrium. The assumption with regard to taxes is made due to the fact that investors have different tax brackets, which making them to choose different assets to invest in. Tax rates differ according to the type of income such as dividends, capital gains or interest.
- 7) *All investors are price-takers* meaning that the portfolio choices and transaction scales of individual investors do not noticeably affect the asset prices. So, it is assumed that the market is perfectly competitive as there is a large number of investors, each with wealth that is small compared to the total market value of capital assets.

Certainly, these assumptions ignore many market imperfections and world complexities. There are many empirical tests of the CAPM made by researchers by relaxing one or more of these assumptions, which have become a subject of considerable theoretical debates.

The implications of these assumptions provide the basis of the CAPM theory. The most important CAPM implications are as follows:

- In equilibrium, all investors will choose to hold the market portfolio of risky assets irrespective of their risk preferences. Notice, that the *market portfolio* includes all risky assets traded. Obviously, different investors hold different portfolios,

combinations of which in total form the market portfolio plus the risk-free assets. This issue is called as the *separation principle*.

- The risk of every individual security is defined by its covariance with market portfolio, as all investors hold the market portfolio. So, the remaining non-systematic risk is diversified away. The only systematic or market risk is taken into account. A standardised measure of covariance with the market portfolio is known as *beta coefficient* of the asset. Beta measures the extent to which asset return and market return move together. Beta is the proper measure of risk for individual assets and portfolios regardless of whether these portfolio are efficient or not. Yet, the risk of efficient portfolios can be measured by beta or variance (or standard deviation). The inefficient portfolios risk can be measured only by beta, which is defined as:

$$\beta_i = \frac{\text{Cov}(R_i, R_m)}{\text{Var}(R_m)} \quad (2.1)$$

- Since non-systematic risk is diversified away, the return on investor's portfolios is affected only by systematic risk (as measured by beta), which must be compensated to investors. The security market line (SML) explains this issue by showing the relationship between expected return and beta of the asset or portfolio.

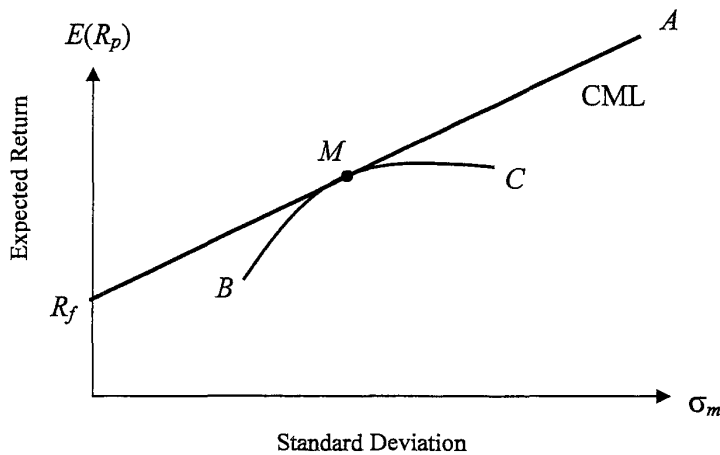
## 2.2 The Capital Market Line (CML)

One of the properties of CML is the availability of risk-free asset and assumption that the investors can borrow and lend unlimited amount at a risk free rate  $R_f$ . Thus, investors have the opportunity to construct their portfolios using risky assets and risk-free assets. The weights of risky and risk-free assets in the portfolio depend on the investor's risk-aversion level. The expected return on a portfolio  $R_p$  thus:

$$E(R_p) = WR_f + (1-W) E(R_m) \quad (2.2)$$

where  $R_m$  is the market return (return on the market portfolio of risky assets) and  $W$  is the weight of risk-free assets in the portfolio, therefore  $(1-W)$  is the proportion of risky assets in the portfolio. Using lending and borrowing at the risk-free rate, investors face the low return when lending, the high return when borrowing, and the middle return when investing only in risky assets (Sharpe, Alexander and Bailey, 1999). In other words,

lending decreases expected return on a portfolio while borrowing increases expected return, as it is clearly seen on the Figure 1 representing the CML.



**Figure 2.1 The Capital Market Line**  
 Source: Farrell, J., (1997), *Portfolio Management*

The  $R_f M$  part of the CML represents all portfolios consisting from market portfolio  $M$  and risk-free lending. The  $MA$  part of the CML shows combinations of risky assets and borrowed funds at risk-free rate. Obviously, portfolios located on the  $MA$  line, earn a higher return than those portfolios located on the  $R_f M$  fragment. So, the CML represents all possible alternative portfolios constructed from risky assets of portfolio  $M$  and risk-free lending and borrowing.

The  $BMC$  curve represents the frontier for inefficient portfolios except portfolio  $M$ . Portfolios allocated on the  $BMC$  curve become inefficient due to the fact that portfolios plotted on the CML offer higher return than those on the  $BMC$  line at the same level of risk. Any combination along the CML (except portfolio  $M$ ) is superior to any combination of risky assets alone. The investors will, therefore, prefer to invest in portfolios allocated on the CML constructing them from risk-free assets or loans and the portfolio  $M$ .

The portfolio containing all risky assets represented by the market portfolio  $M$ , which is the tangency point of efficient frontier to the CML on the figure, consists of all investors' portfolios, i.e. all investors share the market portfolio. It is a result of the CAPM

assumption of homogeneous expectations of investors. Since, all investors attempt to optimize their personal portfolios, each investor will choose the same market portfolio, which has the optimal weights of all risky assets (Ross, Westerfield and Jaffe, 2001). If one security was not included in the market portfolio, it would mean that there is no demand on that security, thus, its price would be zero or the price would be declining until it reaches the price which will be attractive for investors to hold that security in their portfolios. So, this price adjustment leads to the fact that all securities are included in the market portfolio.

The market portfolio is the only portfolio on the CML consisting from risky assets. All other portfolios on the CML include risk-free assets or risk-free loans. Therefore, the CML equation is presented by the risk-free rate and the return on market portfolio:

$$E(R_p) - R_f = \frac{E(R_m) - R_f}{\sigma_m} \sigma_p \quad (2.3)$$

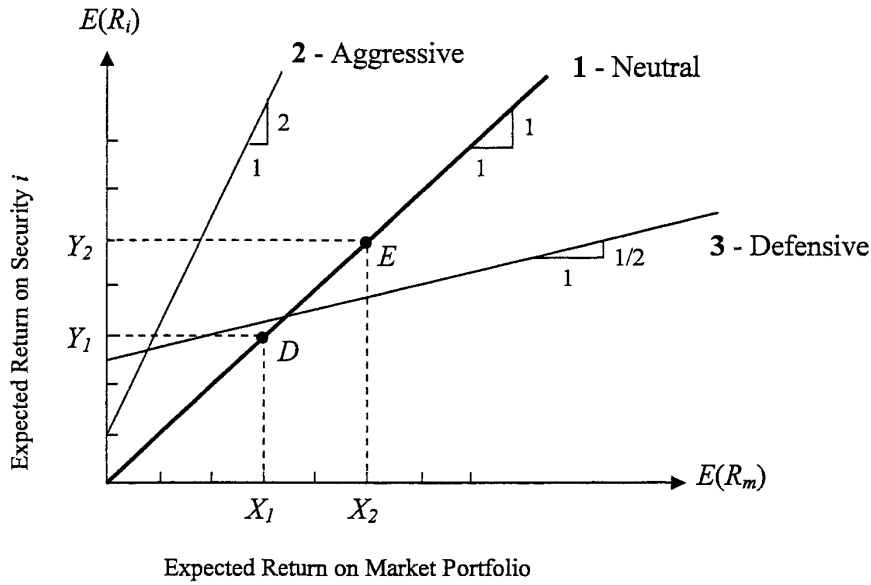
The CML formula shows that for a portfolio to be located on the CML, the expected return in excess of the risk-free rate must be proportional to the standard deviation of that portfolio,  $\sigma_p$ . The slope of the CAPM is the *market price of risk* and signifies the rate, at which investors will be compensated for

To summarize, the CML provides the required risk premium for any portfolio plotted on the CML comprising the risk-free asset and the market portfolio. In other words, the CML shows the risk-return relationship and the appropriate measure of risk for an efficient portfolio, which is the standard deviation of portfolio's return (Farrell, 1997). The CML also indicates that the risk-return relationship is linear. So, the CML gives the explanation of risk premium of the portfolios. However, there exists another issue. It is the relationship between the expected return of a particular security and the expected return on the portfolio. This explanation is provided by the characteristic line.

### 2.3 The Characteristic Line

Since the specific (non-systematic) risk can be diversified away by constructing of a portfolio, investors require to be rewarded only for bearing a systematic (market) risk. The systematic risk provides the extent to which the expected return on individual security

varies with expected return on the market. The extent of this co-movement is provided by the characteristic line.



**Figure 2.2 The Characteristic Line**

Source: Levy and Post, (2005), *Investments*

As it is shown on the Figure 2 the characteristic line's slope relates the expected return on the security  $E(R_i)$  to the expected return on the market portfolio  $E(R_m)$ . The line 1 shows that for every percentage point increase in the expected return on the market portfolio, for example from  $X_1$  to  $X_2$ , the expected return on the security  $i$  rises by the same percentage form  $Y_1$  to  $Y_2$ , meaning that the slope of the line equals 1, which represents the *beta coefficient*  $\beta_i$ . Beta indicates how the return on the security  $i$  responds to a given variation in the market return. It measures the systematic risk of an asset with regard to the market portfolio. Beta coefficient is defined as the ratio of the security's covariance of return with the market  $Cov(R_i R_m)$  to the variance of the market  $Var(R_m)$  and is calculated as follows:

$$\beta_i = \frac{Cov(R_i R_m)}{Var(R_m)} = \frac{\rho_{im} \sigma_i \sigma_m}{\sigma_m \sigma_m} = \frac{\rho_{im} \sigma_i}{\sigma_m} \quad (2.4)$$

Beta coefficient depends on the correlation coefficient between security  $i$  and market, and the standard deviations of the security and market. So, if beta equals 1 the characteristic line will be neutral, as it is seen form the Figure 2. The line 2 represents an aggressive

security since its beta equals 2, meaning that securities with a beta greater than 1 have more risk than the market. The line 3 has a beta lower than 1, thus it has less risk than market and referred to as defensive securities (Levy and Post, 2005).

However, in reality, it is not possible to know or calculate correct betas and expected returns, since nobody knows what will be the return on assets. Thus, beta measurement cannot be done ex-ante. Therefore, beta is generally calculated using historical returns on securities, i.e. ex-post observations. The ex-post returns on assets considered to be the indicator of the future asset returns. Using the past data of asset returns for the long period it is possible to calculate the average return on security, market return and beta, which will be the approximate measures in defining future outcomes. A regression line (characteristic line) represents a set of past relationships. It is defined as follows:

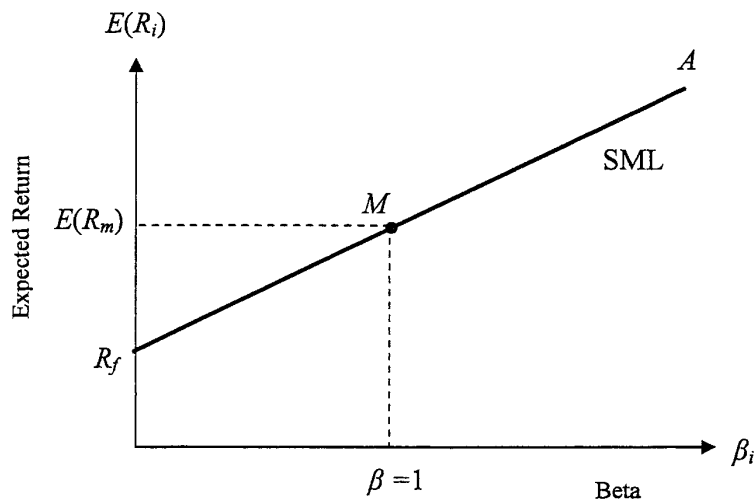
$$R_i = \alpha_i + \beta_i R_m + e_i \quad (2.4)$$

where the *alpha parameter*  $\alpha_i$  is the intercept of the regression line, which indicates the return on the security  $i$  in the case of market return equals zero. It shows the return on the security with regard to the market or systematic risk. Consequently, it indicates what return the security offers for unsystematic risk. The positive alpha represents the excess return on the security and the negative alpha – the negative return. Sometimes alpha is called the abnormal return because it has an economic meaning as the measure of return in excess of the risk-adjusted return. The  $e_i$  parameter is the deviation from the line, called *random or residual return*. It represents the unexpected return arising from influences not determined by the model. According to the type and scale of influence the residual return may take any value. It depends on the number of observations. The bigger the number of observations, the smaller the residual return value. Thus this random return has the mean  $E(e_i) = 0$  and variance  $\sigma_{e_i}^2$ . Variance  $\sigma_{e_i}^2$  represents the unsystematic risk or non-market risk of the security. Since the CAPM considers the 'general market effect' as the 'systematic effect' only, which influences return on securities, then if market effect will be removed from the security, there should be no significant correlation between securities (Farrel, 1997). Therefore, the residual returns are uncorrelated between securities.

The additional point in the explanation of risk premium is how the securities in portfolios are valued. It is given by the Security Market Line.

## 2.4 The Security Market Line (SML)

The CAPM suggests that only market risk is relevant in assessing the required risk premiums for individual securities and portfolios, which is measured by beta coefficient reflecting the sensitivity of the return on securities to movements in the market return (Pike and Neale, 1999). On the other hand, the volume of the risk premium on individual securities or portfolios depends on the correlation of security or portfolio with the market. If a security is perfectly correlated with the market, the risk premium for the market would be appropriate for that security, otherwise the required return depends on beta, which is explained by the SML. The CAPM supposes that when the efficient capital market is in equilibrium, i.e. all securities are accurately priced, the relationship between risk and return is given by the SML. The SML is commonly used as a means of testing the CAPM. It is the direct analog of the single-index model, which is the statistical model for representing a return-generating process.



**Figure 2.3 The Security Market Line**

*Source: Farrell, J., (1997), Portfolio Management*

The SML equation shows that the required return on securities consists of the return on a risk-free asset and a premium for risk related to the market's own risk premium, which varies with regard to the beta of the security:

$$E(R_i) = R_f + \beta_i [E(R_m) - R_f] \quad (2.5)$$

The risk premium  $[E(R_i) - R_f]$  is equal to the market risk premium  $[E(R_m) - R_f]$  multiplied by the beta of security. Since  $[E(R_m) - R_f]$  is greater than zero because investors require a positive risk premium, the greater the beta the higher the risk and thus the required risk premium. For example, if beta equals 1, the required return will then equal the average return for all securities (the return on the market portfolio), otherwise, the higher the beta, the higher the risk premium and the total required return. However, a relatively high beta does not guarantee a relatively high return. The actual return depends partly on the behaviour of the market, which operates as a proxy for general economic factors, the same way as the expected returns for the individual security depends on the expected return for the market (Myers, 2003). In rising market it is worth to hold a high beta securities or aggressive securities, which are allocated on the *MA* segment of the SML of the Figure 3, while defensive securities, which are allocated on the *RM* segment of the SML, offer some protection against falling market. Nevertheless, holding a single high beta security is risky, even on a rising market. Undiversified portfolios of investments, whatever their beta values, are subject of specific risk factors. Therefore, portfolio construction is very important to eliminate the unique risks of individual securities.

The SML follows from the fact that the market portfolio is allocated on the efficient frontier. In order for the market portfolio to be mean-variance efficient, the securities should be positioned on the SML. If at least one asset lies above or below the SML, then the market considered to be inefficient. In other words, the SML provides a necessary and sufficient condition for mean-variance efficiency of the market portfolio (Pike and Neale, 1999). It shows whether the security is overpriced or underpriced according to its beta. For example, if security has beta, which equals zero, then  $E(R_i) = R_f$ , which means that the return of the security with no systematic risk equals the risk-free rate. If beta equals 1, then  $E(R_i) = E(R_m)$  meaning that the price of security varies in tandem with the market and thus the security has the same market risk as the market yielding the same rate of return as the market portfolio. Securities with beta higher than 1 (aggressive securities) have more systematic risk than the market portfolio. It means that, in equilibrium, the aggressive securities will have a higher expected return than the market. The defensive securities with beta lower than 1, have smaller expected return than the market. There are also assets with



negative beta. Since the most assets have a positive betas, the assets with negative betas are very uncommon. However, the negative beta can be obtained by short-selling and using derivative securities. According to the SML, the assets with negative beta have the expected return lower than riskless asset  $R_f$ . These securities can only be used as the part of portfolio to reduce the risk of portfolio and stabilise its return. So, the SML can also be used for assets with negative betas.

## 2.5 The Relationship between the CML and the SML

The SML represents the relationship between expected return and beta and the CML in turn provides the relationship between expected return and standard deviation. Both, beta and standard deviation are measures of risk. However, the SML differs from CML, as the SML applies to all individual assets or portfolios, regardless of whether they are efficient. On the contrary, the CML applies only for efficient portfolios. Since investors do not require a compensation for non-systematic risk, the CML does not apply to individual assets or portfolios which are inefficient (Levy and Post, 2005). The variance of the asset, which consists of the market risk and non-market risk, can be explained as:

$$\sigma_i^2 = \beta_i^2 \sigma_m^2 + \sigma_{e_i}^2 \quad (2.6)$$

where the first component  $\beta_i^2 \sigma_m^2$  is the company's systematic risk (market risk), which provides the asset's variance that is relevant to overall market volatility, and the second component  $\sigma_{e_i}^2$  is the company's non-market risk, which provides the asset's variance that is not relevant to the market. According to this equation beta will be equal:

$$\beta_i = \sqrt{(\sigma_i^2 - \sigma_{e_i}^2) / \sigma_m^2} \quad (2.7)$$

Replacing beta in the SML equation gives the following:

$$E(R_i) = R_f + [E(R_m) - R_f] \times \sqrt{(\sigma_i^2 - \sigma_{e_i}^2) / \sigma_m^2}$$

Since efficient portfolios do not have non-market risk then  $\sigma_{e_i}^2 = 0$ , and therefore the above equation reduces to the CML:

$$E(R_i) = R_f + \frac{E(R_m) - R_f}{\sigma_m} \sigma_i \quad (2.8)$$

So, the efficient portfolio includes only systematic risk  $\sigma_p^2 = \beta_p^2 \sigma_m^2$ , as  $\sigma_{e_i}^2 = 0$ . Thus the risk can be measured by beta (SML) or by variance or standard deviation (CML). However, it is not applicable for inefficient portfolios since  $\sigma_{e_i}^2 > 0$ . In this case only the SML applies.

Inefficient portfolios, which lie below the CML, have non-systematic risk that is measured by the distance between those inefficient portfolios and CML. Certainly, such deviations from the CML do not imply that the market is in disequilibrium. It only means that the assets are inefficient and the investor should not hold them in isolation. The equilibrium status of the market can be determined by checking whether all assets lie on the SML rather than the CML. In equilibrium all assets lie on the SML and below CML.

## 2.6 Testing the CAPM

To summarize all above-mentioned theoretical issues, the CAPM can be defined as the financial theory that describes the relationship between an asset's market risk and its expected return. To test this definition many empirical tests of the CAPM have been conducted by researchers. Majority of those empirical tests on the CAPM follow the methodology pioneered by Black, Jensen and Scholes (1972) and Fama and MacBeth (1973), referred to as the "traditional approach". Under this method, the data set is divided into two: the estimation and the testing periods. In the estimation period, the beta is estimated by running a regression of realized returns of an asset against market returns, which is also called as time series test or the first-pass regression. Then the resulting beta of the first regression is used to proxy for the true beta of the asset and is regressed against the excess return of the asset in the testing period. This second stage is also referred to as cross-sectional test or second-pass regression.

Generally, this regression takes the following form:

$$\hat{R}_{it} - \hat{R}_{ft} = \hat{\gamma}_{0t} + \hat{\gamma}_{1t} \hat{\beta}_i + \hat{\mu}_{it} \quad (2.9)$$

where the left-hand side is the return of the asset in excess of the risk-free rate and  $\hat{\beta}_i$  is the estimated beta from the first regression. The most crucial feature of this method is its test

of hypothesis whereby it specifies that if  $\gamma_1$  is insignificantly different from zero, then the CAPM is taken to be rejected. This simplicity explains its popularity among researchers.

Although substantial criticism has already been raised in the early years of the CAPM and the Arbitrage Pricing Theory has been developed as an alternative equilibrium model (Roll, 1977), the CAPM remains popular. This may at least partially be due to the fact that early empirical tests (Black et.al., 1972, Fama and McBeth 1973) found support for the model in its original form or in the zero beta version of the model developed by Black and Fischer (1972). Roll's critique, however, indicates the significant objectional points with regard to the CAPM. Firstly, he argues that the only test of the CAPM is whether the market portfolio is mean-variance efficient. He says that the CAPM will always hold if the market proxy that is used is mean-variance efficient. If the proxy is not mean-variance efficient, the relationship between expected return and risk will not hold. Secondly, Roll points out that, since the market portfolio is not identifiable, the CAPM cannot be really tested as the market proxies that are used do not include all assets such as real estate, human capital e.t.c, which are not possible to measure. In addition, it was indicated that if asset's performance is measured relative to an index that is ex-post efficient, then from the mathematics of efficient sets, no security will have abnormal performance when measured as a departure from the SML. Conversely, if performance is measured relative to an ex-post inefficient index, then any ranking of portfolio performance is possible, depending on which inefficient index has been chosen. So, Roll concludes that the CAPM is useless because it is not feasible. Roll instead advocates the use of the Arbitrage Pricing Theory. However, perhaps the usefulness of the Roll's critique is reminding us that if we find that the CAPM tests do not hold, then the so-called "market" (or really market proxy) is not mean-variance efficient.

Regardless Roll's critique many empirical tests have been done since then, resulting in rejection of the CAPM as well as supporting it. Among studies rejecting the model, Fama and French (1992) gained the most prominence since it was able to cast the strongest doubt on the model's ability to predict the relationship between beta risk and return. Rather than attributing asset returns to beta as specified in the CAPM, their study found that firm size and the book-to-market equity ratio are far superior in explaining asset returns. This result

is against the CAPM hypothesis. This study also produces a controversial finding on the validity of CAPM: first, as the main model in investigating the relationship beta-return; and second, the beta role in explaining financial asset return. This was widely reported in the financial press as the death of beta. (The detailed overview of this paper is presented below). Other studies related to static CAPM are reported in Banz (1981), Reinganum (1981), Gibbons (1982), Basu (1983), Chan, Chen and Hsieh (1985), Shaken (1985), Bhandari (1988) and Jagannathan and Wang ZhengYu (1996). They have found that the static CAPM is unable to explain the cross-sectional variation of average returns. Tinic and West (1984) found that January has a larger risk premium than the other months and further that the significant risk-return relation only exists in January. When data for the January months are excluded, the estimates of risk premiums are not significantly different from zero, casting serious doubt on the validity of CAPM. Lakonishok and Shapiro (1984) examined the monthly returns of all stocks traded on the New York Stock Exchange for at least four years between 1962 and 1980. They have found that return on individual security is not specifically related to its beta, but is significantly related to the market capitalization values. Lakonishok and Shapiro (1986) updated their analysis to 1981 and included firm size to help explain average portfolio returns. They have concluded that the traditional (beta) as well as the alternative (residual Standard deviation) risk measure still cannot explain the cross-sectional variation in returns, only size can significantly explain it.

Nevertheless, Pettengill, Sundaram and Mathur (1995) make a successful attempt to explain the weaknesses of flat relationship between beta and return. They employ the conditional CAPM between beta and return in up-market and down market. The result finds that the positive relationship between beta and return during up-market and negative relationship during down market. This result supports the continued importance of beta as a measure of market risk. It is followed by Jagannathan and Wang (1996) and Fletcher (2000) who examine the relationship between beta and returns in international stock returns using the model of Pettengill et al (1995). The aim of these studies is to determine whether beta has a role to play in explaining cross-sectional differences in country index returns. These studies support the model proposed by Pettengill et al (1995). In up market months, there is a significant positive relationship between beta and return. In down market months, there is a significant negative relationship between beta and return. In addition, the

relationship is symmetric between up market and down market months. A related study on the conditional CAPM has also been done by Hodoshima, Gomez and Kunimura (2000) that investigate the relationship between beta and return in Nikkei stock market. The result shows the significant conditional relationship between beta and return. (The detailed overview of these papers are presented below).

### 3. EARLIER EMPIRICAL STUDIES IN TESTING THE CAPM

One of the earliest empirical studies of the CAPM is that of Black, Jensen, and Scholes (1972). They have found that the data are consistent with the predictions of the CAPM, given the fact that the CAPM is an approximation to reality just like any other model. Black, Jensen, and Scholes (1972) have used all NYSE stocks during 1931–65 to form 10 portfolios with different historical beta estimates. They have used the time-series method with the realized returns to test the following equation  $\tilde{R}_{jt} = \alpha_j + \beta_j \tilde{R}_{Mt} + \tilde{\epsilon}_{jt}$ . If the CAPM is valid then the intercept  $\alpha_j$  should equal zero. They argue that cross-sectional test can be misleading because of the process structure which appears to be generating the data. They believe that time series test is more powerful since it is free of the difficulties associated with cross-sectional methods. However, they have conducted the cross-sectional test too. So, the time series test starts with regression of average monthly excess returns on beta. The 30-day T-bill rate has been used as the risk-free asset. To avoid an aggregation problem the tests have been run on grouped data. That is, they have formed portfolios of securities and estimated the average return and average beta for these portfolios. In order to avoid the selection bias problem in the ranking procedure, the data from subsequent time period have been used to estimate the portfolio risk. This ranking process is independent of the measurement errors in beta of individual securities. The coefficient beta has been estimated for the five-year periods and then grouped into 10 portfolios in accordance with the beta ranking from high to low (first portfolio contains high betas, last portfolio – low betas). The average monthly excess return on the market proxy used in the study is 1.42 percent. The estimated slope for the resulting regression line is 1.08 percent instead of 1.42 percent as predicted by the CAPM. The estimated intercept is 0.519 percent instead of zero as predicted by the CAPM. The t-statistics that Black, Jensen, and Scholes report indicate that the slope and the intercept of their regression line are significantly different from their theoretical values. Test results show that high-beta securities had negative intercepts and low-beta securities had positive intercepts. Thus the high-risk securities earned less and low-risk securities earned more on-average over 35-year period than the amount predicted by CAPM. The second cross-sectional test indicates that the relation between mean returns and beta of portfolios is linear. However, the intercept and slope of the cross-sectional relation varied over subperiods. This does not necessarily mean that the data do not

support the CAPM. As Black (1972, 1993) has pointed out, these results can be explained in two plausible ways. One is measurement and model specification error that arises due to the use of a proxy instead of the actual market portfolio. This error biases the regression line's estimated slope toward zero and its estimated intercept away from zero. The other plausible explanation is simpler: if no risk-free asset exists, then the CAPM does not predict an intercept of zero. In fact, Black, Jensen, and Scholes have concluded that the data are consistent with Black's (1972) version of the model  $ER_i = ER_z + (ER_m - ER_z) \beta_i$ .

Another classic empirical study of the CAPM is by Fama and MacBeth (1973). They examine whether there is a positive linear relation between the average return and beta and whether the squared value of beta and the volatility of the return on an asset can explain the residual variation in average returns across assets that is not explained by beta alone. The test has been conducted using monthly return data (including dividends and capital gains) for all common stocks traded on the New York Stock Exchange for the period from 1926 to 1968. As a proxy for the market index, "Fisher's Arithmetic Index", i.e. an equally weighted average of the returns on all stocks listed on the NYSE has been used. Since using portfolios rather than individual securities in the CAPM tests causes the loss of information, in order to avoid it a wide range of values of estimated portfolio betas have been obtained by forming portfolios on the basis of ranked beta values for individual securities. Then to minimize the regression phenomenon, when in a cross-section of betas, high observed estimated betas are above the true betas and the low observed estimated betas are below the true ones, the betas for formed portfolios have been estimated using the subsequent period data. So, using the first 4 years (1926-29) of monthly returns 20 portfolios have been formed. The following 5 year-period data (1930-34) has been used to recalculate stock's betas and portfolio betas to start the test. Portfolio betas has been calculated as a simple average of individual securities included in the portfolio for the following 4 years (1935-38). This process has been repeated monthly to adjust portfolio betas to allow for delisting of securities. So, using the returns on the 20 portfolios for that 4-year subperiod (1935-38) the following cross-sectional regression has been run for each month of the subperiod:  $R_{pt} = \hat{y}_{0t} + \hat{y}_{1t} \beta_{p,t-1} + \hat{y}_{2t} \beta_{p,t-1}^2 + \hat{y}_{3t} S_{p,t-1}(\varepsilon_i) + \eta_{p,t}$ , where  $\hat{y}_{0t}, \hat{y}_{1t}, \hat{y}_{2t}, \hat{y}_{3t}$  are the stochastic coefficients,  $S(\varepsilon_i)$  is the standard deviation of the least-squares residuals  $\varepsilon_i$ . The results from this regression, i.e. the time series month by

month values of regression coefficients are the inputs for the 4-year subperiod (1935-38) test. The regressions have been run for every subperiod. More precisely, there are nine overlapping portfolio formation periods. After every portfolio formation period there is an initial estimation period and then testing period. The t-statistics for testing the hypothesis that  $\hat{y}_it = 0$  has also been done. The test results show that the values support the conclusion that on average there is a positive relationship between risk and return.

The work of Pettengill, Sundaram and Mathur (1995) tests the conditional relationship between beta and returns. They argue that the validity of Sharpe-Lintner-Black model has not been directly tested since the previous tests made by other researchers use realized returns instead of expected returns, and the model does not show a direct relationship between portfolio betas and portfolio returns when the realized market return is less than the risk-free return reasoning in the fact that returns for high beta portfolios are less than return for low beta portfolios. Pettengill, Sundaram and Mathur consider these as the main reason of finding negative results or weak and inconsistent relationship between beta and returns in the previous tests. They say that testing results of Fama and MacBeth are positive because of relationship between beta and returns of tested period are positive on average. Pettengill, Sundaram and Mathur have analyzed months with positive and negative market risk premiums separately. A monthly comparison of the CRSP index (as the market proxy) and the 90-day T-bill rates (as a measure of the risk-free rate) over the period 1936 through 1990 has been made by Pettengill et al.. The results show that the T-bill rate exceeds the market return in 280 out of 660 total observations. It has been concluded that the existence of a large number of negative market return periods show that previous tests for unconditional positive relationship between beta and returns are biased against finding a systematic relationship. So, they have conducted two tests: 1) test of a systematic, conditional relationship between betas and realized returns; 2) test of a positive long-run tradeoff relationship between beta and return. The period from 1926 to 1990 has been taken as the sample period of the test. The CRSP equally-weighted index has been taken as a proxy for the market index. Monthly returns for securities and CRSP index returns data have been collected from the CRSP monthly databases. The three-month Treasury bill rates have been taken as the measure of the risk-free rate. Firstly, to test a systematic, conditional relationship between betas and returns Pettengill, Sundaram and



Mathur have used a modified version of the three-step portfolio method of Fama and Macbeth (1973). The sample period has been divided into 15-year subperiods, which then has been divided into a portfolio formation period, a portfolio beta estimation period, and a test period of a five years each. Betas have been estimated for each security by regressing the security's return against the market return. Securities have been equally divided into 20 portfolios according to the estimated beta rankings (securities with lowest betas have been placed in the first portfolio, securities with the highest betas – in the last portfolio). Portfolio betas have been estimated by regressing portfolio returns against the market returns. The third step has been modified taking into consideration the conditional nature of the relationship between beta and returns. The regression coefficients from the following equation have been examined:  $R_{it} = \hat{y}_{0t} + \hat{y}_{1t} * \delta * \beta_i + \hat{y}_{2t} * (1 - \delta) * \beta_i + \varepsilon_t$ . where  $\delta = 1$ , if  $(R_{mt} - R_{ft}) > 0$  (when market excess returns are positive) and  $\delta = 0$ , if  $(R_{mt} - R_{ft}) < 0$  (when market excess returns are negative). This equation has been tested for each month of the test period by calculating  $y_1$  or  $y_2$  depending on the sign of market returns. Secondly, to test a positive long-run tradeoff relationship between beta and return a standard t-test has been used in order to determine if the market returns are, on average, positive. For symmetry test the risk premiums  $\hat{y}_1$  (for up market – market with positive return) and  $\hat{y}_2$  (for down market – market with negative return) have been compared with some sign and mean value adjustments. The empirical results of the above-mentioned tests provide strong support for a systematic but conditional relationship between beta and realized returns. The results of traditional (unconditional test) shows the significant relationship between beta and returns for the total sample period 1936-1990, but not for the subperiods 1951-1970 and 1971-1990. The results of conditional test shows that there is a positive relationship between beta and returns during periods with a positive market returns and an inverse (negative) relationship during periods with a negative market returns. That is for periods with negative market return high beta portfolios have lower returns than low beta portfolios. These findings are hold for each subperiod and across all months in a year. The results of positive long-run tradeoff test indicate a significant positive reward for holding market risk during the overall sample period, which has been found having on average a positive excess return. Also the strong consistency in the relation between beta and return and symmetry in the risk premium in up and down markets have been found during the observations.

Jagannathan and Wang (1996) tested the static CAPM assuming that betas do not remain constant over time and including the human capital when measuring the return on aggregate wealth. They argued that two assumptions of the CAPM that betas of assets remain constant over time and that the return on stocks measures the return on the aggregate wealth portfolio, are unreasonable. Jagannathan and Wang have conducted the conditional test of the CAPM relaxing the assumption that betas remain constant over time allowing betas vary over time. In order to use the better proxy for return on market portfolio they have constructed the model with three betas. They have used the returns on US stocks listed on the NYSE and AMEX for the period of 1962-1990. Using Fama and French (1992) approach 100 portfolios have been created. For every calendar year, starting from 1963, all firms have been sorted into size deciles according to their market values. Then for each size decile betas for each firm have been estimated using past return data and CRSP value-weighted index as the market proxy. The next step was sorting firms within each size decile into beta deciles according to their beta estimates. Then the return on each of 100 portfolios has been calculated for the next 12-month period. This procedure has been repeated for the each calendar year to obtain a time-series of monthly returns for 1963-1990 years for 100 portfolios. Using this return data, the traditional CAPM specification has been examined, resulting in the fact, that the strong size effect suggests that the conventional specification of the CAPM is inconsistent with the data. With regard to the human capital, the authors have argued that there is an important difference between human capital and other physical assets owned by firms, paying attention that the entire cash flow coming from the use of the physical assets used by firms is promised away by issuing financial securities. However, for the human capital only a portion of the labor income is secured by issuing mortgages. Thus, it has been concluded that the factors affecting the return on human capital cannot be identified by examining returns on mortgages and other financial assets. Therefore, the return on human capital has been assumed to be an exact linear function of the growth rate in per capita labor income. So, the main model has been developed including the return on the market portfolio and the growth rate per capita labor income. In order to visually compare the performance of the different specifications, the fitted expected return, which has been computed by using the estimated parameter values in a model specification, has been plotted against the realized

average return. Then it has been shown that when the CAPM holds in a conditional sense (i.e. expected returns and betas vary over time in systematic manner), unconditional expected returns on assets are linear in the average beta and a measure of beta instability over time. It is also has been noticed that when betas remain constant over time, the model collapses to the familiar static CAPM. Jagannathan and Wang have demonstrated that the empirical support for their conditional CAPM specification is strong, and when betas and expected returns are allowed to vary over time by assuming that the CAPM holds period by period. The size effect and the statistical rejections of the model specifications become much weaker. Although the conditional model performed better than the static model, the authors advocated caution in interpreting these results as strong support for the conditional CAPM, explaining it by the inherent word dinamism, which might need another missing explanatory variables in the static model. So, due to the fact that the model might be inaccurate it has been concluded that the CAPM, like any other model, is only an approximation of reality.

The paper of Fletcher (1997) tests the conditional cross-sectional relationship between beta and return using Pettengill et al. (1995) approach and the role of size in UK stock market. The test covers the period of 1975-1994. The 30-day UK Treasury Bill monthly returns have been used as a risk-free rate. The return on the FTA Index (Financial Times All Share Index) and EWI (equally-weighted index) have been used as the market proxy. First, the relationship between portfolio beta and size and the monthly returns of the 100 size-beta portfolios have been tested without separationong up markets and down markets. The betas of the portfolios have been estimated for the full period of 1975-1994 with respect to the FTA and EWI proxies. Betas have been estimated from the regression model of Fama and MacBeth (1973). Then cross-sectional regression, estimated by Ordinary Least Squares has been run each month. That regression equation has also been used to test the size effect by including the additional size variable. For the conditonal CAPM test the sample period has been divided into up market months and down market months. The sample period has been divided into two ten-year sub-periods. All securities have been ranked on the basis of market value and grouped into 10 portfolios in ascending order. Within each size decile, the beta of the security has been estimated from the regression of the security return on a constant and the return on the FTA Index using past data over the previous 36 to 60

months. Equally-weighted returns have been calculated each month on the 100 portfolios over the subsequent year. This procedure has been repeated for each year. The monthly risk premiums have been split into two subsamples: with positive market return and with negative market return. The test results show that there are a substantial number of down market months over the sample period, and in periods of up market months, there is a significant positive relationship between beta and return, where high beta portfolios received higher returns than low beta portfolios. In periods of down-markets, the results show that there is a significant negative relationship between beta and return where, high beta portfolios earned a lower return than low beta portfolios. In fact, it has been found that in a majority of the size deciles, the low beta portfolios had a higher mean returns than the high beta portfolios and that there is no significant positive risk premium on beta in UK stock market. The test results indicate that the relationship is stronger in down market months than up market months. As for the symmetry of the relationship between beta and returns in up market and down market months, it has been found that this relationship is unsymmetrical. It has also been found that there is little support for size effect on returns in UK stock market. In general, the evidence within the paper shows that there is a conditional relationship between beta and returns in UK stock market. Fletcher has concluded that beta may still play a useful role for portfolio managers.

The aim of the Hodoshima, Garza-Go'mez, and Kunimura's work (2000) is to examine the relationship between beta and returns in the Japanese stock market. They have supported the idea of Pettengill et al. (1995) and found that data are better explained by making a distinction between positive and negative market excess returns. In addition to the Pettengill et al. (1995) approach they have also analyzed a model including as explanatory variables, size, and book to market equity ratio, which Fama and French (1992) examined. They have compared different relationships based on summary statistics of goodness of fit and tested results obtained from the cross-sectional regression method. The test results show a positive conditional relationship between beta and returns in up market, which is offsetting a negative conditional relationship between return and beta in down market, resulting in the absence of any unconditional relation between beta and returns. Monthly stock returns listed on the first section of the Tokyo Stock Exchange (TSE) for the period of January 1956–December 1995 have been used in the study. Monthly average of the

next day call money rates with collateral, taken from the Nikkei database have been used as the risk free rate. As the Proxy of market return Hodoshima, Garza-Go'mez, and Kunimura have used two indexes: a value weighted index (VWI), provided by the Japanese Securities Research Institute (JSRI) for all the firms listed on the first section of the TSE, and an equally weighted index (EWI) of all the firms in the sample. The EWI has been used for small stocks and the VWI – for nonmanufacturing sector and financial sector. This has been explained by the fact that EWI gives a better description for small stocks and VWI is more influenced by nonmanufacturing sector and financial sector. 20 portfolios of stocks based on the ranking of the betas have been constructed. Then they have been re-estimated, using the next two years data and beta for each portfolio by the average of re-estimated betas of the stocks for that portfolio. In the next step the stocks have been assigned to the portfolios formed in the beginning and the portfolio return have been obtained by averaging returns of the stocks for each portfolio. They have used two years data to construct portfolios and two years data to estimate betas of the portfolios. To test the conditional CAPM the cross-sectional regression has been used and cross-sectional regression method of Fama and MacBeth (1973) has been used to test the unconditional CAPM. Hodoshima, Garza-Go'mez, and Kunimura have obtained the following test results: about 40% of monthly observations of the Japanese stock market excess return consist of negative returns; there exist positive and negative linear relationships between beta and returns in up market and down market; the conditional relationship is in general better fit in the down market than in the up market in terms of the goodness of fit measures given by given by the  $R^2$  and standard error; beta and book to market equity ratio is not significant, while the size is significant with a negative coefficient in the unconditional CAPM test and conversely – in conditional CAPM test. It has been concluded that it is appropriate to differentiate the returns into up and down markets for the relevance of beta, although the same does not apply to other explanatory variables as the size and book to market equity ratio, and that beta is a suitable measure of risk to explain return in the conditional relationship.

Shakrani and Ismail (2001) have tested the conditional CAPM for Islamic unit trusts in Malaysia. They have found a flat unconditional relationship between beta and returns as a result of test conducted without separating positive and negative market excess return. The

results with significant positive relationship between beta and returns have been found using the conditional CAPM test. The results of cross-sectional regression analysis show a positive relationship between beta and returns in up market and a negative – in down market. Moreover, it has been found that the relationship between beta and returns is higher in down market. The test has been conducted adopting the Pettengill et al. (1995) approach using the same cross-sectional regression method but with returns on Islamic unit trusts and taking the inter bank rate for 1 month as a risk free rate, and syariah index – as a proxy of market index. To evaluate the difference in the relationship between beta and returns the statistics as t-test, adjusted R-squared and standard error have been used. The conditional and unconditional tests have been conducted. First, the beta has been calculated for each individual Islamic unit trust by estimating the unit trusts return as a function of risk free rate and market excess return. Then, this estimated beta has been used for the whole sample (Sample A), and two sub samples, i.e. Sample B (1 May 1999 – 23 June 2000) and Sample C (24 June 2000 – 31 July 2001). Second, the estimated beta has been re-estimated for each Islamic unit trust using the average return of Islamic unit trusts. Then, this average beta and average return has been estimated using the unconditional and conditional relationship of the cross-sectional data. Summary statistics of syariah index has showed 60 positive values and 58 negative values and the market excess return with syariah index - 96 negative values and 22 positive values. The summary statistics of the time series average and standard deviation of beta and return for Islamic unit trusts shows the small positive average returns of below 1 for the market return. Standard deviations of the market return and the market excess return also become smaller after separating up markets from down markets. The cross-sectional regression (using the average return and average beta) results show the weak relationship between beta and return for unconditional test. On the contrary, the conditional test shows the positive results. So, the work of Shakrani and Ismail support the conditional CAPM and it has been concluded that the beta can be used as a measure of market risk to explain the cross-sectional differences in Islamic unit trust returns.

The paper of Brennan, Wang and Xia (2002) presents the estimation and test of a simple model of Intertemporal Capital Asset Pricing (ICAPM) and evaluation of the model ability to account for the returns on portfolios sorted according to size and book-to-market ratio,

as well as according to industry. The ICAPM suggests the pricing of risk, related to the variation in investment opportunity set, limiting the number of state variables to be considered to the two that are required to describe that set, i.e. ICAPM allows for time-variation in the real interest rate and slope of the capital market line. Authors of this paper refer to the ICAPM suggested by Merton (1973) and Fama and French (1995). The model parameters and time series of the state variables have been estimated using data on eight synthetic constant maturity zero coupon US Treasury bond yields and inflation rate calculated from the CPI for the period of 1952- 2000. For cross sectional pricing tests the returns on 25 size and book-to-market sorted value weighted portfolios, the CRSP value weighted market portfolio and the nominal short interest rate for the same period have been used. The time series of the state variables have been estimated by using the model of nominal bond yields in a Kalman filter, to extract the time series of the unobservable state variables from data on bond yields and inflation. In the simple ICAPM that has been estimated, time variation in the instantaneous investment opportunity set is described by the dynamics of the real interest rate and the maximum Sharpe ratio. It has been assumed that “these two variables follow correlated Ornstein-Uhlenbeck processes; consequently, the current values of these variables are sufficient statistics for all future investment opportunities and are the only state variables that are priced in an ICAPM setting”. Firstly, valuation model allowing for a stochastic interest rate and Sharpe ratio has been constructed. Equation of the expected return on the asset and then Sharpe ratio have been estimated. Then the stochastic process equation for real cash flow expectation has been set. In order to value nominal bonds the stochastic process for the price level has been specified. Second, by specializing the estimated pricing model (so that the innovation in the pricing kernel is an linear function of the market return and the innovations in stochastic interest rate and Sharpe ratio), a specific version of the ICAPM has been obtained. It has been indicated that the returns on Fama-French hedge portfolios are correlated with the innovations in state variables. In order to estimate the risk premia on the Fama-French hedge portfolios that are implied by simple ICAPM two separate approaches have been used: a pricing kernel based approach, which uses the pricing kernel coefficients that have been estimated from the bond yield data and a tracking portfolio approach, which uses portfolios of equities that have maximum correlation with the estimated innovations. Finally, the ability of the Fama-French hedge portfolios and the

tracking portfolios to explain the returns on 25 size and book-to-market sorted portfolios over the sample period has been compared. The tests using 30 industrial portfolios instead of the size and book-to-market sorted portfolios have also been conducted. The overall results of ICAPM test show that zero-coupon nominal bond yields are linearly related to the state variables, the real interest rate, the Sharpe ratio and expected rate of inflation. The estimated real interest rate and Sharpe ratio show strong business cycle related variation. The Sharpe ratio has found to be related to the return on market portfolio and the level of stock prices (measured by the market dividend yield) – to both the Sharpe ratio and real interest rate estimates. These findings are consistent with the model predictions. The results of formation of the size and book-to-market sorted portfolios, tracking the innovations in the state variables show that risk premia on the Fama-French hedge portfolios are explained by the ICAPM. The ICAPM is rejected only when the risk-free interest rate is allowed to differ from the Treasury Bill rate. The results of tests using 30 industrial portfolios show that the model is not rejected using these returns, although both the simple CAPM and the Fama-French model are rejected.

The studies of Elsas, El-Shaer and Theissen on beta and returns of German stock market (2003) also provide a significant relationship between beta and returns. Their observations are based on the Pettengill et al. (1995) method. In order to show how the results of Fama and McBeth test change according to the conditional nature of the relationship between beta and returns Elsas, El-Shaer and Theissen have conducted Monte Carlo simulations. In order to empirically test the relationship between beta and returns they have used data from the German stock market (1960-1995 years period). The results show the strong support of relationship between beta and returns taking into account its conditional nature. Compared to the unconditional test conducted by them, which result shows weak relation between beta and returns, the adjusted (conditional) test provides totally positive results. As in the test performed by Pettengill et al. (1995), the positive results have been also found in all subperiod and in the total sample period by Elsas, El-Shaer and Theissen when testing the German stock market. They argue that the previous tests made by other researchers on the German stock market have the result of weak relationship between beta and returns because of the periods with negative market returns. However, it has been mentioned that modifying the test for conditional nature of the relationship between beta and returns does



not resolve all problems (partially of econometric nature) with regard to the CAPM tests. It has been also stressed that the advantage of conditional test is that the hypothesis of relationship between beta and returns can be tested independently of the hypothesis of a positive market returns. So, before the empirical test Elsas, El-Shaer and Theissen have conducted Monte Carlo simulations. An artificial market, where the CAPM holds with known parameters and the market portfolio is determined and is ex-ante efficient, has been constructed. The results of two test procedure over the 1000 simulation runs indicate that the traditional two-stage approach has a weak support of the relationship between beta and returns rejecting the null hypothesis only in 177 out of 1000 cases, whereas the conditional test rejects the null hypothesis in 985 out of 1000 cases. The methodology of empirical analysis of Elsas, El-Shaer and Theissen is the same as Pettengill et al. The cross-sectional regression separately made for up market and down market and the above-mentioned equation of Pettengill et al. (1995) have been used in tests. The DAFOX index has been used as a proxy of market portfolio. The average rate for three months term deposits (Deutsche Bundesbank) has been used as a risk-free rate. Betas for each stock have been estimated by regressing the time series of stock returns against the time series of the index returns. As it is mentioned above the significant relationship between beta and returns has been found when conducting the conditional test.

Pedro B. de Ocampo (2003) have tested in his work the CAPM in the Philippine stock market using two methods: traditional approach of Fama and Macbeth (1972) and Pettengill et al. (1995) method. The strong risk return trade-off has been found under the conditional test of the CAPM. The modified three-step approach of Fama and Macbeth with cross-sectional regression has been used to conduct the unconditional CAPM test and the Pettengill et al. (1995) method has been used to conduct the conditional CAPM test. The monthly share prices for 103 stocks listed in the Philippine Stock Exchange for the period January 1992 to December 2002 have been used in the test. The 91-day Treasury-Bill rate has been used as a risk-free rate and the Phisix index – as a proxy of market index. The results of the traditional test, which show that the beta coefficient is insignificantly different from zero having the incorrect sign and the intercept is significantly different from zero, do not support the unconditional CAPM. However, as predicted by traditional CAPM, the relationship between risk and return has been found to be linear. The results

also show that factors other than beta do not affect returns. As for the conditional CAPM test results, it has been found that there is a positive relationship between risk and return where high beta stocks earn higher returns in up market and a negative relationship between risk and return where high beta stocks earn lower return in down market. The conditional CAPM test results also confirm the linear relationship between beta and returns and that factors other than beta do not affect returns. The risk-return relationship between periods of positive and negative excess market returns has been found to be symmetrical as it was found by Pettengill et al. However, it has not been found a positive reward for holding risk as predicted by Pettengill et al. method. It is explained by the fact that the stock prices for the period of January 1998 - December 2002 have been affected by the Asian economy crisis, where the 65% of risk free rates was higher than the market returns.

The recent work of Zhang and Wihlborg (2004) intended to study the conditional and unconditional CAPM in six European emerging markets making, provides a significant conditional relationship between beta and returns in domestic markets and positive results of CAPM test in two international markets. Zhang and Wihlborg have used Pettengill et al. (1995) method to test risk-return relationship and estimate cost of equity capital of firms in six emerging markets, i.e. Cyprus, Czech Republic, Greece, Hungary, Poland, Russia, and Turkey. The distinction between domestic and international CAPM has been made in order to examine whether the domestic CAPM will outperform the international one as a result of high degree of the market segmentation. Two tests of CAPM: conditional and unconditional have been conducted. Fama and McBeth (1973) two-step regression approach has been used for unconditional test and Pettengill et al. (1995) method has been used for conditional test. The monthly time series share prices of 753 firms from six emerging markets for the period of 1995 – 2002 have been used in tests. As the market proxy for each country, they have used the returns of market indexes of that country and as the risk-free rate for each country, they have used short-term Treasury-bills of that country. The Morgan Stanley world index returns have been used as the world market portfolio returns. The CPI has been used for high inflation countries to calculate the real returns. The empirical results show that the betas and returns are much higher in the above-mentioned six emerging markets than those of the world market. For four countries with the high inflation periods the real returns are much lower. The returns for six countries are not

highly correlated with each other and with the world market. It has been found that more than 50% of realized returns are negative in the most countries and the null hypothesis with mean equals zero cannot be rejected. The existence of conditional rather than unconditional relationship between risk and returns has been proved during the tests. Zhang and Wihlborg have found the positive risk-return trade-off. As for international and domestic CAPM tests, it has been found that the international CAPM performs well, especially into two countries (Czech Republic and Russia), which are more intergated in the world markets. Zhang and Wihlborg have concluded that beta is still a useful measure of risk in emerging domestic markets taking into account its conditional nature.

The aim of Sandoval and Saens study (2004) is to test the conditional and unconditional CAPM in Latin America using the data from the Argentinean, Brazilian, Chilean, and Mexican stock markets, to examine the effect on return of other risk factors as size, book-to-market ratio and momentum and to test the market integration hypothesis in the Latin American stock markets. Sandoval and Saens have noticed that since the Latin American emerging stock markets have their high industrial concentration (especially Argentinean and Mexican), the variables as size and market concentration might be an important factor in explaining the firm-specific cross-sectional return variations. The study covers the period between January 1995 and December 2002 using the weekly returns in U.S. dollars of the stocks for each country. The U.S. Treasury bill rate has been used as a risk free rate and the MSCI-LATAM stock market index has been used as proxy for the market portfolio. Fama and MacBeth (1973) approach using three steps of econometric tests has been used to analyse the unconditional CAPM. First, since Latin American's stocks are traded infrequently, Sandoval and Seans have used the aggregated coefficients method proposed by Dimson (1979) to estimate the individual betas; that is they have regressed individual security returns against five lagged, matching and five leading market return terms using both the Latin American Stock Market Index and S&P 500 Index as a proxies for the market portfolio. Then they have formed portfolios, where securities with the lowest estimated betas have been assigned to the first portfolio and securities with the highest betas - to the last portfolio. In the second stage, the portfolio betas have been estimated for each two-year period (totaling in eight portfolio beta estimation periods) and used as explanatory variables in the further next year. The last stage includes cross-

sectional regressions based on Black (1972) CAPM definition, which has been estimated by a pooled cross-sectional OLS giving estimates of the average values of weekly coefficients  $\gamma_{0t}$  and  $\gamma_{1t}$  in the testing period. Then average values of weekly coefficients have been tested to find whether they are significantly different from zero. When testing the conditional CAPM, testing periods have been divided into up and down market weeks. Tests results show the inconsistency with a positive and significant relationship between portfolio betas and returns, with the exception of the Mexican stock market. The low R-squares show that the model might be either misspecified or additional risk factors other than beta might be required to explain the relationship between risk and return. The results of conditional CAPM show that in four Latin American countries the stock markets present a significant and positive beta risk premium during up markets and a significant but negative beta risk premium during down markets, providing a strong support for a systematic but conditional relationship between portfolio betas and realized returns in each of the Latin American stock markets. It has been found that in 417 weeks over the period 1995 through 2002, the risk free rate exceeds the market return by 50%, 48%, 53% and 46% for the Argentinean, Brazilian, Chilean and Mexican stock market, respectively. Sandoval and Saens have also found that the null hypothesis of a symmetrical relationship between risk and return during periods of positive and negative market returns is rejected at the 5% level for the Latin American stock markets with exception of Mexico. It shows that Latin American stock markets react more to down than to up market. As for the test of the effect on return of other risk factors as size, book-to-market ratio and momentum, its results show that these extra risk factors is not commonly priced across the Latin American stock markets and do not contribute to explain significantly the cross sectional stock return variations in Latin American stock markets.

The paper of Tang and Shum (2004) presents the test of the risk-return relationship in the Singapore stock market for the period of 1986 – 1998 years, having the following purposes: test for a conditional relation between beta and returns; test for a positive long-run risk-return tradeoff; test whether other measures of risk in addition to beta affect asset pricing when up and down markets are split. The data collected from the Pacific-Basin Capital Markets (PACAP) Databases consist of monthly returns of 144 listed stocks and equally weighted as well as value-weighted market returns (to determine whether the

results of using these two market proxies are different). The one-month Singapore interbank offer rate (SIBOR) has been used as the risk-free rate. The 13-year period has been divided into three non-overlapping subperiods: the construction period (1986 – 1989), the estimation period (1989 – 1992), and the testing period (1992 – 1998). Betas of individual stocks have been estimated and 20 equally weighted portfolios have been formed according to the estimated betas ranking in the construction period. In the estimation period, betas and other risk measures of each portfolio formed in the construction period have been estimated. The final stage includes the regressions of portfolio returns against betas and other risk measures from the testing period. The whole process has been re-done by dropping the first month's data in the estimation period and adding the second month's data in the testing period and has been repeated up to the last month of the testing period. To compare the goodness of fit across different models of risk measures, adjusted coefficient of determination has been used. The test results of unconditional CAPM show the weak positive relation between beta and returns. The strong systematic but conditional relationship between beta and realized returns has been found as a result of conditional CAPM test using Pettengill et al. (1995) approach. It has been also found that high-beta portfolios receive a larger positive risk premium than low-beta portfolios in up market and high-beta portfolios receive higher losses than low-beta portfolios in down market. The findings reject the symmetrical risk premiums in up and down markets due to the fact that the market risk premium in up market doubles that in down market. Tang and Shum have found that unsystematic risk plays an even more significant role than beta in pricing the Singapore securities in up market and not only compensates the systematic risk but also the unsystematic risk. In fact, they have suggested for investors in the Singapore stock market do not hold diversified portfolios when the market excess returns are positive. They have also concluded that beta is still a good measure of risk and the other risk measures are also useful in explaining cross-sectional variations in stock returns.

Medvedev (2004) tested the CAPM under ambiguity. He has used Zhang (2002) model of individual choice under ambiguity in the context of Kwon (1985) market model of asset returns. The implications of the ambiguity for equilibrium asset prices have been studied,

assuming uncertain volatility setup used in the quantitative finance and proposed by Avellaneda et al (1995) under the assumption that the volatility process is not known but the volatility always lies within known bounds. The theoretical CAPM under the market model with unknown volatilities of asset returns has been derived as follows: first, the equation of individual project that generate random return has been estimated, then under some assumptions the random variable has been estimated. Then including into account multiple assets and introducing ambiguity into the market model of Kwon (1985), which allows to derive the CAPM without restrictive assumptions on the utility function or asset return distribution, the classical version of CAPM has been modified with regard of including the second factor that measures the degree of the ambiguity in returns. Some assumptions have been made to reduce the problem of choice under ambiguity to the standard problem of choice under probabilistic uncertainty. The estimated two-factor CAPM has been tested for US stocks using 48 industry portfolios constructed by Fama and French for the period of 30 years 1973 – 2003. Then the residual returns of these industry portfolios have been estimated and their volatility standard deviations have been calculated. The modified CAPM has been then tested running cross-sectional Fama-McBeth regression. The test results show that the ambiguity factor is statistically significant and the effect of market beta is insignificant. It has been concluded that these results are economically meaningful.

The studies of Lewellen and Nagel (2004) present the test whether the conditional CAPM can really explain asset-pricing anomalies. They argue that unconditional CAPM does not describe the cross section of average stock returns because the CAPM does not explain why small stocks outperform large stocks, why firms with high book-to-market ratios outperform firms with low B/M ratios. So, Lewellen and Nagel perform the test of conditional CAPM in order to learn whether it can explain these patterns. They say that if the conditional CAPM holds, only the small deviations from the unconditional CAPM should be expected to be found – much smaller than those observed empirically. For the tests the time-series CAPM regressions for stock portfolios have been used. The unconditional test has been conducted using the full time series of returns for each portfolio, restricting alpha and beta to be constant. The conditional test has been conducted using a common approach where beta is a function of observed macroeconomic variables.

In order to solve the problem of availability of the full set of state variables, the conditional alphas and betas have been estimated using short-window regressions. The CAPM regressions have been estimated separately every month, quarter, half-year and year using daily, weekly, or monthly returns. So the estimates of each quarter's conditional alpha and beta have been made without using any state variables or making any assumptions about quarter-to-quarter variation in beta, except one assumption that beta is relatively stable within the quarter and each regression can simply treat it as constant. The conditional CAPM has been tested in two ways. First, it has been directly tested whether the conditional alphas are zero. Second, it has been tested whether betas vary over time in a way that might explain stocks unconditional alphas by calculating the volatility of betas and their correlation with business conditions and the market risk premium. Since the regressions have been estimated over short intervals instead of monthly returns daily or weekly returns have been used. The estimations have been made using overlapping returns (observations overlap by four days). The tests focus on size, B/M, and momentum portfolios from 1964 to 2001 using NYSE and Amex common stock returns. The CRSP value-weighted index has been used as a market proxy and the T-bill rate has been used as a risk-free rate. The size and B/M portfolios has been constructed as those of Fama and French (1993). Twenty five size-B/M portfolios based on the intersection of five size and five B/M portfolios have been formed. So, the tests have been based on six combinations of the twenty five size-B/M portfolios. The momentum portfolios have been formed separately. The stocks have been sorted every month into deciles based on the past six-month returns. The portfolios have been hold for overlapping six-month periods. The tests use returns compounded over three horizons: daily, weekly and monthly. The test results of the conditional alphas provide strong evidence against the conditional CAPM. B/M and momentum portfolios alphas remain large, statistically significant, and close to their unconditional estimates. The test results of conditional CAPM also show that the size effect is weak, as in unconditional tests, but small stocks now show a hint of abnormal returns. It has also been found that beta vary considerably over time due to the changes in true conditional betas (not estimation error) but it is not enough to explain large unconditional pricing errors. Therefore, it has been concluded that the conditional CAPM performs about as poorly as the unconditional CAPM and does not explain the B/M and momentum effects.

The recent study of Ang and Chen (2005) is aimed to estimate and develop the conditional CAPM with time-varying betas, time-varying market risk premia, stochastic systematic volatility to test the book-to-market effect over the long run using returns of all stocks listed on the NYSE, AMEX, and NASDAQ. Ang and Chen argue that although alphas and betas are generally estimated by OLS (Ordinary Least Squares) and a large OLS of alpha is considered as an anomalous return relative to the CAPM, beta variations over time, which are correlated with time-varying market return, result in misspecified standard OLS inference, and therefore it is not possible to use it to assess the fit of the conditional CAPM. They have showed that OLS alphas and betas are biased and not a proper estimate of conditional alphas and betas. The main stress in the study has been done on the time-variation of conditional betas in estimating conditional alphas. Their estimation technique provides direct, consistent estimates of conditional alphas and of the time-series of conditional betas. The conditional CAPM has been used to portfolios sorted by book-to-market ratios, which have a spread in average returns that cannot be explained by an unconditional CAPM using an OLS regression. They argue that this can be explained by the one-factor conditional model with time-varying betas. They have also noticed the advantages of their modelling method: 1 - the time-varying betas have been considered as implicit state variables and inferred directly from stock returns; 2 - the betas have been considered as endogenous variables that vary slowly and continuously over time (rather than to assume discrete changes in betas across subsamples and constant betas within subsamples); 3 - they have included predictable time-variations in aggregate market conditions in both the conditional mean and the conditional volatility. Markov Chain Monte Carlo (MCMC) and Gibbs sampling estimation approach, which incorporates the effect of parameter uncertainty, estimates conditional alphas and betas and measures the effect of small sample bias, has been used. They have examined the sample period from July 1926 to December 2001. To test the unconditional CAPM, alphas and betas have been estimated using unconditional one-factor regression. Its result shows that the alphas from the unconditional one-factor model are insignificant for book-to-market sorted portfolios over the long run. As for the conditional one-factor model, it has been found that the model is sufficient to explain the average returns of book-to-market portfolios. Nevertheless, Ang and Chen do not posit that the conditional CAPM is the complete model for the cross-



section of stock returns and that the conditional CAPM can explain all anomalies. They argue just that the conditional one-factor model is able to explain a wider range of returns than previously thought. Their results also emphasize the importance of taking into account time-varying factor inputs before asserting the cross-sectional return pattern anomalous relative to a conditional CAPM.

## **4. TEST OF THE CAPM IN TURKISH CAPITAL MARKET**

### **4.1 Data and Methodology**

In this study two tests of the CAPM have been conducted: traditional and conditional. We use Fama and MacBeth (1973) three-step approach to conduct traditional (or unconditional) CAPM test and Pettengill et al. (1995) method for conditional test:

The sample period for this study extends from January 1995 through December 2004. Weekly returns for all securities listed on the Istanbul Stock Exchange (ISE) and ISE100 equally-weighted index, which has been used as a proxy for the market portfolio, have been obtained from the ISE database. Since there is no available data on returns adjusted for dividends and equity offerings we have used returns calculated directly from the weekly closing prices for securities. The equity offering and dividend payment days of particular securities have been dropped from the related week. In other words, security returns assumed to be zero in the equity offering and dividend payment days due to the fact that security price falls down for a big percentage depending on the equity offering or dividends payment value. As the risk-free rate we have used the three-month US Treasury bill rate adjusted for the inflation in USA and Turkey, and converted into the weekly rates (see Appendix XIII). Since Turkish Treasury bill rates were abnormally high in the period of 1990-2003, that would cause a meaningless results of the test, it has been decided to use the US Treasury bill rates, which in turn, have been adjusted for the USA inflation deducting the inflation rate from the treasury bill rate for the same period. Then using Fisher equation for nominal and real return calculation, US Treasury bill rates have been adjusted for inflation in Turkey using inflation rates for the same period as treasury bill returns. The US Treasury bill rates have been obtained from the US Federal Reserve Bulletin. The US and Turkish inflation rates have been collected from the US Bureau of Labor Statistics database and the State Institute of Statistics of Turkey respectively.

#### 4.1.1 Traditional Approach

According to the Fama and MacBeth (1973) approach the sample period should be divided into subperiods, which in turn divided into three periods: portfolio formation period, estimation period and test period. In the portfolio formation period portfolios are constructed on the basis of ranked beta values for individual securities. Fama and MacBeth argued that portfolio betas can be more precise estimates of true betas than the betas for individual securities since beta estimations cause an unavoidable “errors-in-variables” problem. Also the procedure of portfolio formation, according to the ranked beta values, reduces the loss of information in the risk-return tests caused by using portfolios rather than individual securities. However, Fama and MacBeth asserted that such a procedure could result in a serious regression phenomenon, when constructing portfolios on the basis of ranked betas causes bunching of positive and negative sampling errors between portfolios since in a cross-section of individual betas, high observed betas tend to be above the corresponding true betas and low observed individual betas tend to be below the true betas. Thus, a large portfolio betas would tend to overstate the true portfolio betas and a low portfolio betas would tend to be an underestimate. Therefore, in order to avoid a regression phenomenon, Fama and MacBeth formed portfolios from ranked individual betas calculated from the data for one time period and then used a subsequent period (estimation period) to obtain the portfolio betas. Actually in the estimation period betas of individual securities are re-estimated to compute portfolio betas.<sup>1</sup> So, it is hoped that the regression phenomenon is minimized in portfolio betas since using fresh data randomizes errors in the betas of individual securities within portfolios.<sup>2</sup> In the third time period (test period) portfolio betas, which have been estimated in the second time period (estimation period) are regressed against portfolio returns, which are calculated using the test period data.

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<sup>1</sup> Note that the number of portfolios and their contents do not change in the estimation period.

<sup>2</sup> Fama and MacBeth (1973) pointed out that: “the errors-in-variables” problem and its solving method using portfolios were first mentioned by Blume (1970); portfolio approach has also been used by Friend and Blume (1970) and Black, Jensen and Scholes (1972); the regression phenomenon was first recognized by Blume (1970) and Black, Jensen and Scholes (1972), who offered the similar problem solution as Fama and MacBeth (1973).

Our sample period of 1995-2004 has been divided into five 6-year subperiods such that the test periods overlap with the 1-year interval in order to minimize the errors in beta estimates, which arises from the beta volatility, and hence, get more precise estimates and results. Each subperiod then has been separated into three 2-year periods: portfolio formation period, beta re-estimation period and test period. Each of them consists from 104 weekly observations. In each 6-year subperiod we use securities that are available within that period and meet the data requirements. The five different portfolio formation periods, five estimation periods and testing periods with the numbers of securities available on ISE shown in Table 1.

**Table 4.1 Sample periods and number of stocks**

	Periods				
	1	2	3	4	5
Portfolio Formation Period	1995-1996	1996-1997	1997-1998	1998-1999	1999-2000
Estimation Period	1997-1998	1998-1999	1999-2000	2000-2001	2001-2002
Testing Period	1999-2000	2000-2001	2001-2002	2002-2003	2003-2004
No.of securities at the beginning of formation period	211	235	266	283	297
No.of securities meeting data requirement	167	187	197	206	225

In the portfolio formation period using 2-year period data we estimate the beta for each individual security by regressing the time series of the securities' excess returns, calculated from the natural logarithm of security prices, on the time-series of the index excess returns where excess returns have been obtained by subtracting the risk-free rate from the returns (see Appendix XII). The equation of time-series regression, also often called as the first-pass regression, is as follows:

$$R_{i,t} - R_{f,t} = \alpha_i + \beta_i (R_{m,t} - R_{f,t}) + \epsilon_{i,t} \quad (4.1)$$

where  $R_{i,t}$  is the individual security's return for the  $t$  period,  $R_{f,t}$  is the risk-free rate,  $R_{m,t}$  is the index return,  $\beta_i$  is the beta of security and  $\epsilon_{i,t}$  is the regression residual. Based on ranked beta estimates we have sorted the securities into 20 equally weighted portfolios. Portfolio 1 contains the securities with the highest betas, portfolio 20 - the securities with the lowest betas. Using the Fama and MacBeth (1973) approach the number of securities in the each portfolio have been calculated as  $\text{int}(N/20)$  for the middle 18 portfolios and  $\text{int}$

$(N/20 + \frac{1}{2} [N - 20 \text{ int}(N/20)])$  for the first and last portfolios if  $N$  is even, where  $N$  is the total number of securities to be allocated to portfolios and  $\text{int}(N/20)$  is the largest integer equal to or less than  $N/20$ . The last portfolio includes an additional security if  $N$  is odd. In the following two-year period (estimation period) the portfolio betas have been estimated by recomputing the individual security betas and calculating unweighted averages of the security betas included in each portfolio constructed in the formation period. In the test period the portfolio returns have been calculated by averaging individual security returns in a portfolio using the 2-year data of the test period. Then we run the cross-sectional regression in the following form:

$$R_{p,t} - R_{f,t} = \hat{y}_{0,t} + \hat{y}_{1,t} \beta_p + \epsilon_{i,t} \quad (4.2)$$

where  $\hat{y}_{0,t}$ ,  $\hat{y}_{1,t}$  are the regression coefficients,  $R_{p,t}$  is the portfolio return,  $\beta_p$  is the portfolio beta and  $\epsilon_{i,t}$  is the regression residual. The weekly portfolio returns have been regressed against portfolio betas, which have been computed in the estimation period.

#### 4.1.2 Conditional Approach

In the test period we have estimated the regressions both using the traditional test procedure and using the conditional approach of Pettengill et al. (1995). As Pettengill et al. (1995) argued, the test of the CAPM model needed to be modified due to the model's requirements that a part of the market return distribution is below the risk-free rate. Pettengill et al. said that since the test of relationship between beta and returns based on the assumption of a positive risk-return tradeoff, the expected market return then must be higher than the risk-free return, otherwise all investors would hold the risk-free security. Therefore, the expected return on any risky portfolio is a positive function of beta. However, Pettengill et al. argued that due the fact that CAPM test uses realized returns instead of expected returns the validity of the model would not be directly examined. According to Pettengill et al. there must be conditions under which high beta portfolios earn lower returns than low beta portfolios in order for high beta portfolios to have more risk. The CAPM equation  $(R_{i,t} - R_{f,t}) = \beta_i (R_{m,t} - R_{f,t}) + \epsilon_{i,t}$  shows an exact condition under which the realized returns for high beta portfolios are expected to be lower than the realized returns for low beta portfolios. It shows that the relationship between beta and returns for high and low beta portfolios is actually conditional on the relationship

between realized market returns and the risk-free return. If  $R_m < R_f$ , then  $\beta_p(R_{m,t} - R_{f,t})$  is  $< 0$ . In such cases, the predicted portfolio return incurs a negative risk premium that is proportionate to beta. Hence, according to Pettengill et al., if the realized market return is less than the risk-free return, an inverse relationship exists between beta and return, which means that high beta portfolios have predicted returns that are less than the predicted returns for low beta portfolios. Since the realized market returns occurs to be frequently less than risk-free rate of return, the conditional relationship will have a big effect on tests of the relationship between beta and returns. So, our first step in conducting the conditional test is to analyse the market returns over the sample periods. A week-by-week comparison of the ISE100 index and the risk-free rates over the period 1999 (first test period) through 2004 shows that the T-bill rate exceeds the market return in 155 out of total 311 observations. Therefore, the existence of a large number of negative market excess return periods suggests to modify the CAPM test to account of positive and negative market excess returns in order for test results not to be biased. The conditional test utilize the modified version of the three-step portfolio approach of Fama and MacBeth (1973). The procedure of the first two steps of portfolio formation and estimation periods are remains the same. The portfolios are constructed on the basis of ranked betas and then portfolio betas are estimated in the second period along with the portfolio returns. However, the third step (the test period), which tests the relationship between portfolio beta and returns, is modified to account for the conditional relationship between beta and *realized* returns. As it has been mentioned above, if the realized market return is above the risk-free return, portfolio betas and returns should be positively related, and if the realized market return is below the risk-free return, portfolio betas and returns should be inversely related. Consequently, to conduct the conditional CAPM test, the regression coefficients from the following equation have been examined:

$$R_{p,t} = \hat{\gamma}_{0t} + \hat{\gamma}_{1t} * \delta * \beta_{p,t} + \hat{\gamma}_{2t} * (1 - \delta) * \beta_{p,t} + \varepsilon_t \quad (4.3)$$

where  $\delta = 1$ , if  $(R_{m,t} - R_{f,t}) > 0$  (when market excess return is positive in week  $t$ ) and  $\delta = 0$ , if  $(R_{m,t} - R_{f,t}) < 0$  (when market excess return is negative in week  $t$ ). Due to the fact that  $\hat{\gamma}_1$  is estimated in periods with positive market excess returns, the expected sign of this coefficient is positive. Therefore, the following hypothesis are tested:

Ho:  $\hat{\gamma}_1 = 0$ ,

Ha:  $\hat{\gamma}_1 > 0$ .

Since  $\hat{y}_2$  is estimated in periods with negative market excess returns, the expected sign of this coefficient is negative. Hence, the following hypothesis are tested:

Ho:  $\hat{y}_2 = 0$ ,

Ha:  $\hat{y}_2 > 0$ .

The systematic conditional relationship between beta and realized returns is supported if the null hypothesis is rejected in favor of the alternate in both cases.

## 5. EMPIRICAL RESULTS

### 5.1 The results of the traditional (unconditional) test

Table 2 shows the values of the 20 portfolios betas  $\beta_{p,t}$  used in test procedure and their determination coefficients.

**Table 5.1 Portfolio betas**

Portfolio	Test Period									
	1999-2000		2000-2001		2001-2002		2002-2003		2003-2004	
	$\beta_{p,t}$	R <sup>2</sup>	$\beta_{p,t}$	R <sup>2</sup>	$\beta_{p,t}$	R <sup>2</sup>	$\beta_{p,t}$	R <sup>2</sup>	$\beta_{p,t}$	R <sup>2</sup>
1	0,866	0,381	1,037	0,565	1,040	0,237	1,012	0,953	1,175	0,669
2	0,913	0,395	0,913	0,387	0,951	0,212	1,003	0,966	1,068	0,610
3	0,767	0,421	0,820	0,461	0,846	0,175	0,998	0,936	1,070	0,607
4	0,812	0,396	0,841	0,374	0,814	0,187	0,972	0,931	1,032	0,577
5	0,810	0,445	0,893	0,452	0,747	0,159	0,999	0,946	1,003	0,544
6	0,834	0,500	0,748	0,308	0,696	0,140	0,992	0,927	0,830	0,473
7	0,795	0,361	0,703	0,342	0,740	0,143	0,984	0,936	0,837	0,545
8	0,768	0,339	0,819	0,416	0,694	0,129	0,984	0,948	0,896	0,468
9	0,950	0,506	0,737	0,369	0,762	0,156	0,996	0,924	0,907	0,424
10	0,814	0,416	0,765	0,349	0,647	0,114	0,992	0,939	0,813	0,428
11	0,801	0,386	0,731	0,424	0,679	0,127	0,988	0,904	0,849	0,445
12	0,842	0,488	0,730	0,358	0,515	0,069	0,991	0,909	0,840	0,422
13	0,814	0,408	0,736	0,317	0,640	0,106	0,992	0,918	0,902	0,361
14	0,813	0,488	0,685	0,296	0,613	0,105	0,991	0,923	0,790	0,420
15	0,792	0,350	0,729	0,343	0,551	0,082	0,972	0,917	0,843	0,380
16	0,683	0,310	0,651	0,278	0,695	0,129	1,001	0,928	0,695	0,306
17	0,759	0,378	0,699	0,265	0,562	0,078	0,987	0,924	0,722	0,290
18	0,744	0,344	0,702	0,364	0,572	0,093	0,984	0,905	0,765	0,352
19	0,805	0,447	0,495	0,205	0,362	0,048	0,981	0,924	0,771	0,255
20	0,578	0,253	0,470	0,123	0,347	0,034	0,971	0,901	0,545	0,227

The major test results of the implications of the CAPM for the total sample period are in Table 3. The results are presented for all the sample subperiods. For each period and model, the table shows:  $\hat{\gamma}_{0,t}$ ,  $\hat{\gamma}_{i,t}$  - regression coefficient estimates (regression intercept and slope);  $se(\gamma_0)$ ,  $se(\gamma_0)$  - standard errors of the regression coefficients;  $se(Rp-Rf)$  - standard



error of the portfolio premium;  $R^2$  – the determination coefficient;  $ssreg$  – the regression sum of squares;  $ssresid$  - the residual sum of squares;  $R_m - R_f$  market premium. The table also presents the standard t-statistics  $T(\gamma_0)$  and  $T(\gamma_1)$  for testing hypothesis that  $\hat{y}_{i,t} = 0$ . As predicted by the CAPM the estimated regression coefficients should be equal to:

$$\hat{y}_0 = 0$$

$$\hat{y}_1 = R_m - R_f$$

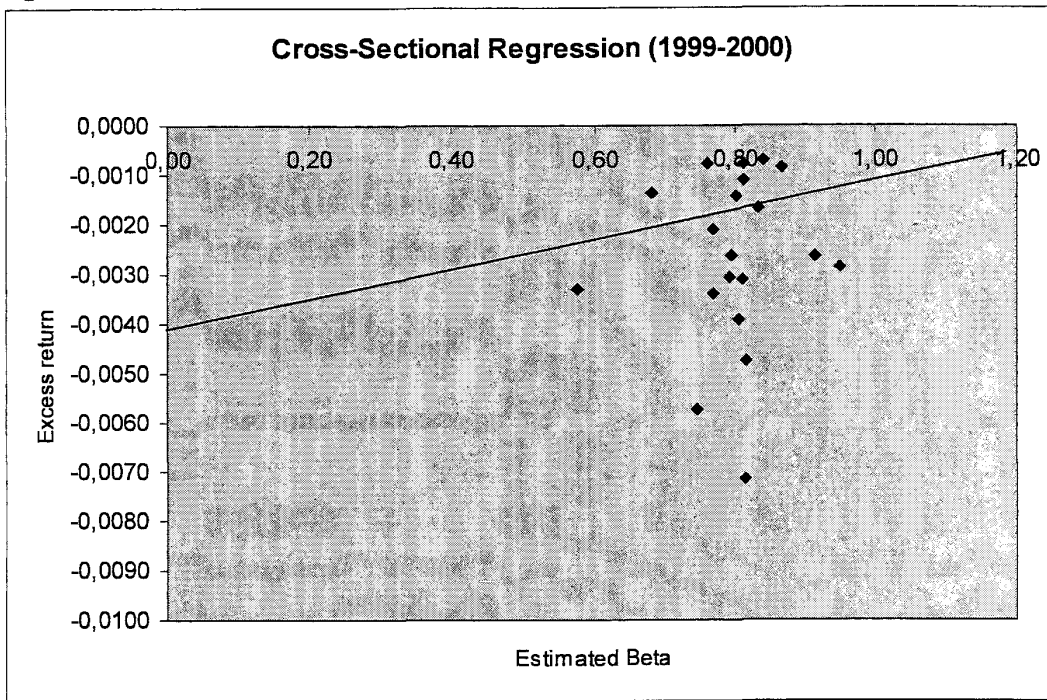
In the four of five periods  $\hat{y}_0$  is very close to zero and all of them are negative. With regard of  $\hat{y}_1$ , it has the result close to predicted by model in only the period (2001-2002). The results also show that R-squared in three periods of five are nearly zero. However, the t-statistics for the second (2000-2001) and third (2001-2002) periods are higher (-1,5690 and -2,7056) than in the other periods. For the same two periods the standard errors of coefficients are quite small and  $\hat{y}_1$  is very close to the market premium  $R_m - R_f$  as predicted by the model. Nevertheless, the  $\hat{y}_1$  in those two periods are negative, which indicates that the risk-return tradeoff not positively related as predicted by the model but inversely. The Figures 5.2, 5.3 clearly show that in the periods 2000-2001 and 2001-2002 there is strong negative inversely related relationship between beta and returns. The other periods do not show any consistent strong results in favor of the CAPM. As it is seen from the Table 3 and Figures 5.1, 5.4, 5.5 in the other three periods 1999-2000, 2002-2003, 2003-2004 the relationship between beta and return is weak although the  $\hat{y}_1$  is positive. In the first period (1999-2000) 14 portfolios, which have almost the same beta value, get the return ranging in big interval from 0 to -0,007 (see Appendix). The t-statistics shows that the value of the slope coefficient is insignificant. The t-statistics of the forth period (2002-2003) indicate the significance of the slope coefficient on the level close to zero 0,014. The last period (2003-2004) also has the small t-statistics for  $\hat{y}_1$ , which shows negligible positive relationship between beta and return and equals 0,004. In general, the overall results of the traditional CAPM test are inconsistent across subperiods and show the weak support of the model predictions. This results, however, do not mean that there is no systematic relationship between risk and return. The results, to a greater extent, might be biased due to the aggregation of positive and negative market excess return periods.

**Table 5.2 Summary results for the traditional test**

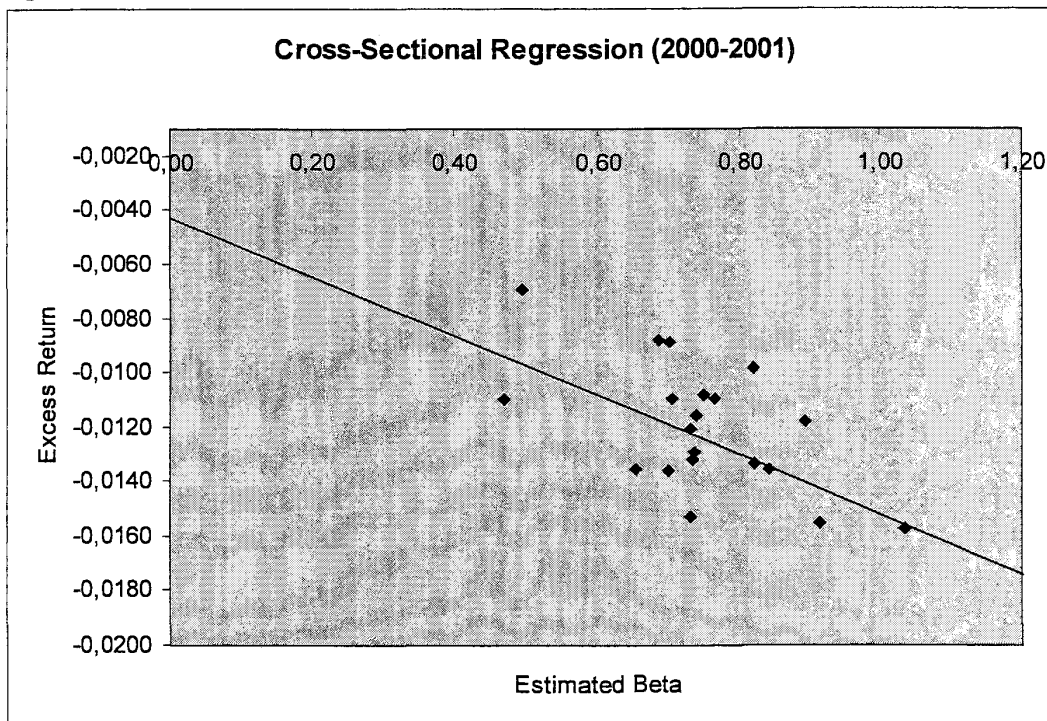
$$R_{p,t} - R_{f,t} = \hat{\gamma}_{0,t} + \hat{\gamma}_{1,t}\beta_p + \epsilon_{i,t}$$

Test Period	Statistic										
	$\gamma_0$	$\gamma_1$	se ( $\gamma_0$ )	se ( $\gamma_1$ )	se ( $R_p - R_f$ )	R <sup>2</sup>	ssreg	ssresid	T ( $\gamma_0$ )	T ( $\gamma_1$ )	Rm-Rf
1999-2000	-0,0044	0,0021	0,0042	0,0053	0,0018	0,0090	0,000001	0,000057	-1,0300	0,4046	-0,00003
2000-2001	-0,0041	-0,0106	0,0026	0,0035	0,0019	0,3425	0,000035	0,000068	-1,5690	-3,0621	-0,0126
2001-2002	-0,0040	-0,0060	0,0015	0,0021	0,0016	0,3028	0,000020	0,000046	-2,7056	-2,8185	0,0015
2002-2003	-0,0197	0,0147	0,0362	0,0365	0,0017	0,0089	0,000000	0,000051	-0,1776	0,0146	-0,3794
2003-2004	0,0023	0,0007	0,0020	0,0023	0,0015	0,0053	0,000000	0,000040	1,1069	0,3087	0,0049

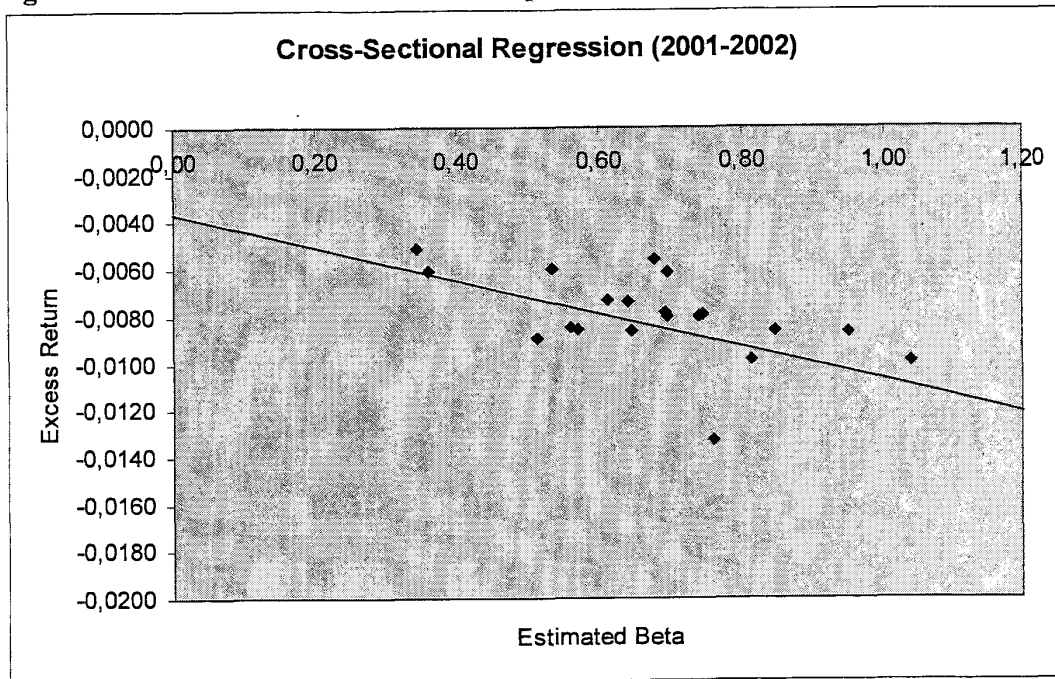
**Figure 5.1** Traditional test results for the period of 1999-2000



**Figure 5.2** Traditional test results for the period of 2000-2001



**Figure 5.3 Traditional test results for the period of 2001-2002**



**Figure 5.4 Traditional test results for the period of 2002-2003**

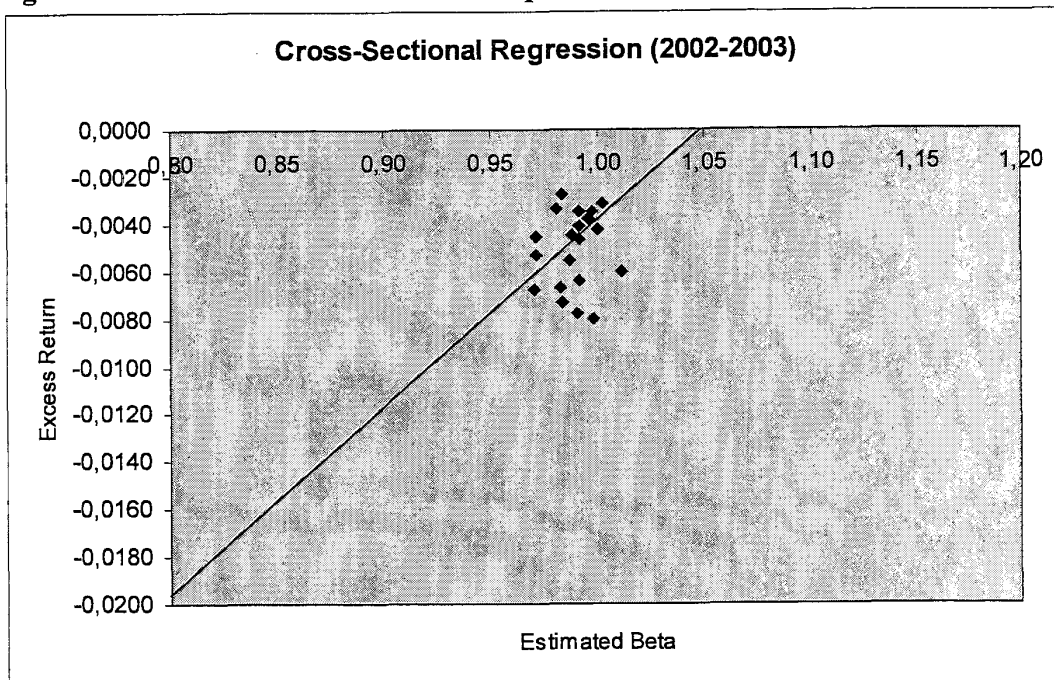
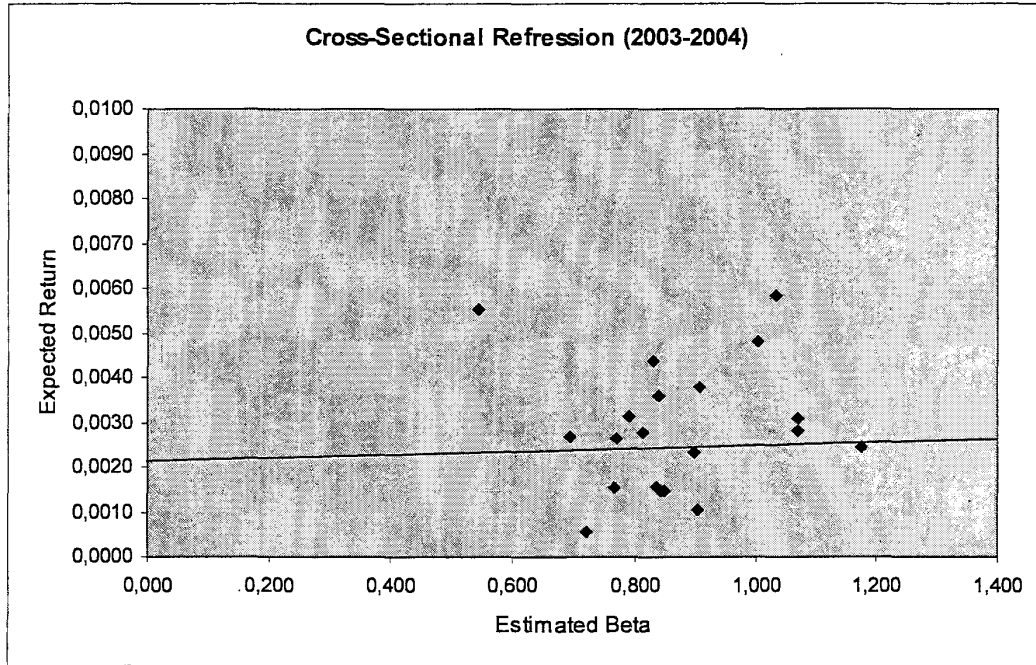


Figure 5.5 Traditional test results for the period of 2003-2004



## 5.2 The results of the conditional test

Given the conditional relation between beta and realized returns, we have tested the dual hypothesis of a positive relation between beta and returns during periods of positive market excess returns and a negative relation during the periods of negative excess returns. For the conditional CAPM test the sample period has been divided into up-market weeks and down-market weeks. The portfolio betas remain the same as reported in Table 2. The statistics of the conditional test for market with positive risk premium is presented in Tables 4. The statistics of the conditional test for market with negative risk premium is presented in Tables 5. The results of Tables 4 and 5 show that the relation between beta and return is statistically significant both in up-market and down-market, and in each subperiod. The slope coefficients  $\hat{\gamma}_1$  have the expected sign. Portfolios with higher betas have higher returns when the market risk premium is positive and lower returns when the market risk premium is negative, which are confirmed by the Figures 5.6 – 5.15. The t-statistics of  $\hat{\gamma}_0$  and  $\hat{\gamma}_1$  are significantly high in all subperiods both in up-market and down-market. Comparing this t-statistics with the t-statistics of the traditional test it can be noticed that the values of t-statistics in conditional test are considerably higher. The standard errors of the coefficients are low as in the traditional test, except the one period 2002-2003 with standard errors  $se(\hat{\gamma}_{0t}) = 0,1215$  and  $se(\hat{\gamma}_{1t}) = 0,1228$  for the up-market, and  $se(\hat{\gamma}_{0t}) = 0,1195$  and  $se(\hat{\gamma}_{1t}) = 0,1208$  for the down market. The R-squared of the regressions are high, starting from the lowest 0,4155 to the highest 0,8161. So, it has been found a highly significant relationship between beta and returns in each subperiod. For every subperiod the hypothesis of no relationship between risk and returns during periods of positive excess market returns is rejected in favor of an expected positive relationship at the 0,01 level. Likewise, for every subperiod, the hypothesis of no relationship between risk and returns during periods of negative excess market returns is rejected in favor of an expected negative relationship at the 0,01 level. Figures 5.6 – 5.15 show the realized returns for the 20 portfolios separately for months with positive and negative risk premium. Figures 5.6 – 5.15 also demonstrate the effect of neglecting the conditional nature of the relation between beta and return. Since, as mentioned earlier, the number of month with positive and negative market risk premia are almost equal in our sample, the unconditional average return of each beta portfolio is approximately equal to the unweighted average of

the two conditional return shown in the Figures 5.6 – 5.15. It is obvious that there is no relationship between the portfolio betas and these unconditional returns. So, the overall results of the conditional test support the conclusion that the betas are related to realised returns in the way predicted by the model.

**Table 5.3 Summary results for the conditional test for the up-market**

$$R_{p,t} = \hat{y}_{0t} + \hat{y}_{1t} * \delta * \beta_{p,t} + \hat{y}_{2t} * (1 - \delta) * \beta_{p,t} + \varepsilon_t$$

Test Period	Statistic										
	$\gamma_0$	$\gamma_1$	se ( $\gamma_0$ )	se ( $\gamma_1$ )	se ( $Rp-Rf$ )	R <sup>2</sup>	ssreg	ssresid	T ( $\gamma_0$ )	T ( $\gamma_1$ )	Rm-Rf
1999-2000	-0,0029	0,0541	0,0110	0,0137	0,0046	0,4644	0,0003	0,0004	-0,2600	3,9509	-0,00003
2000-2001	0,0121	0,0442	0,0053	0,0070	0,0039	0,6889	0,0006	0,0003	2,2896	6,3140	-0,0126
2001-2002	0,0133	0,0362	0,0032	0,0046	0,0034	0,7729	0,0007	0,0002	4,1791	7,7553	0,0032
2002-2003	-0,4606	0,4883	0,1215	0,1228	0,0056	0,4676	0,0005	0,0006	-3,7905	3,9761	-0,3794
2003-2004	-0,0053	0,0306	0,0042	0,0047	0,0030	0,6982	0,0004	0,0002	-1,2626	6,4536	0,0049

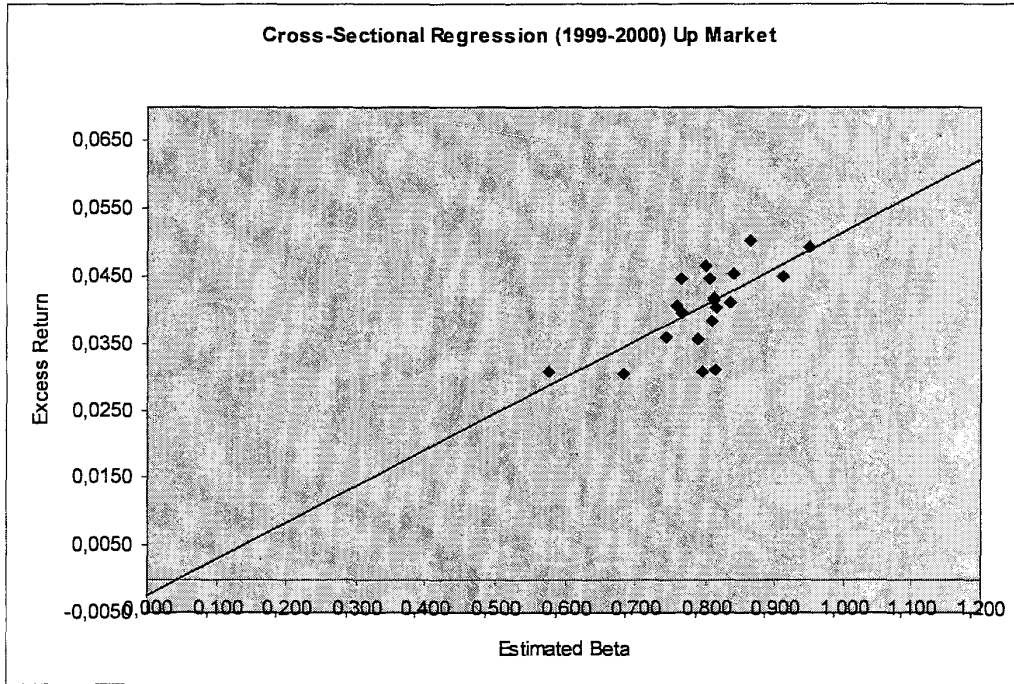
**Table 5.4 Summary results for the conditional test for the down-market**

$$R_{p,t} = \hat{y}_{0t} + \hat{y}_{1t} * \delta * \beta_{p,t} + \hat{y}_{2t} * (1 - \delta) * \beta_{p,t} + \varepsilon_t$$

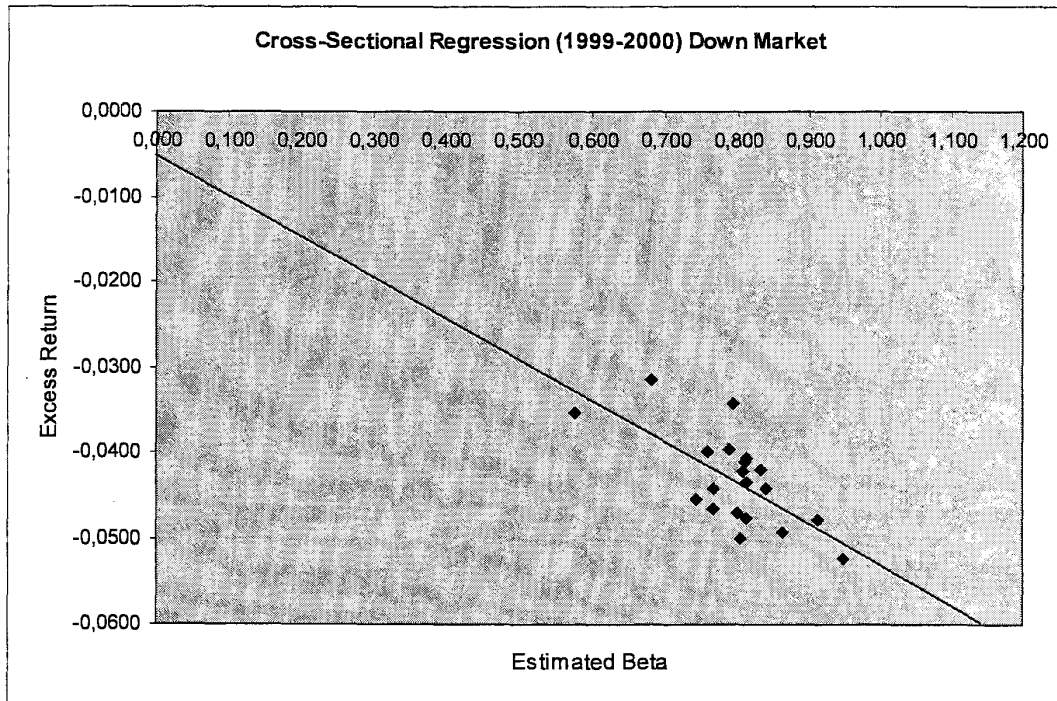
Test Period	Statistic										
	$\gamma_0$	$\gamma_1$	se ( $\gamma_0$ )	se ( $\gamma_1$ )	se ( $Rp-Rf$ )	R <sup>2</sup>	ssreg	ssresid	T ( $\gamma_0$ )	T ( $\gamma_1$ )	Rm-Rf
1999-2000	-0,0052	-0,0476	0,0097	0,0121	0,0041	0,4630	0,0003	0,0003	-0,5385	-3,9399	-0,00003
2000-2001	-0,0153	-0,0484	0,0057	0,0076	0,0042	0,6939	0,0007	0,0003	-2,6712	-6,3872	-0,0126
2001-2002	-0,0180	-0,0401	0,0033	0,0047	0,0035	0,7999	0,0009	0,0002	-5,5401	-8,4805	0,0033
2002-2003	0,3964	-0,4321	0,1195	0,1208	0,0055	0,4155	0,0004	0,0006	3,3161	-3,5771	-0,3794
2003-2004	0,0121	-0,0378	0,0037	0,0042	0,0027	0,8161	0,0006	0,0001	3,2490	-8,9388	0,0049



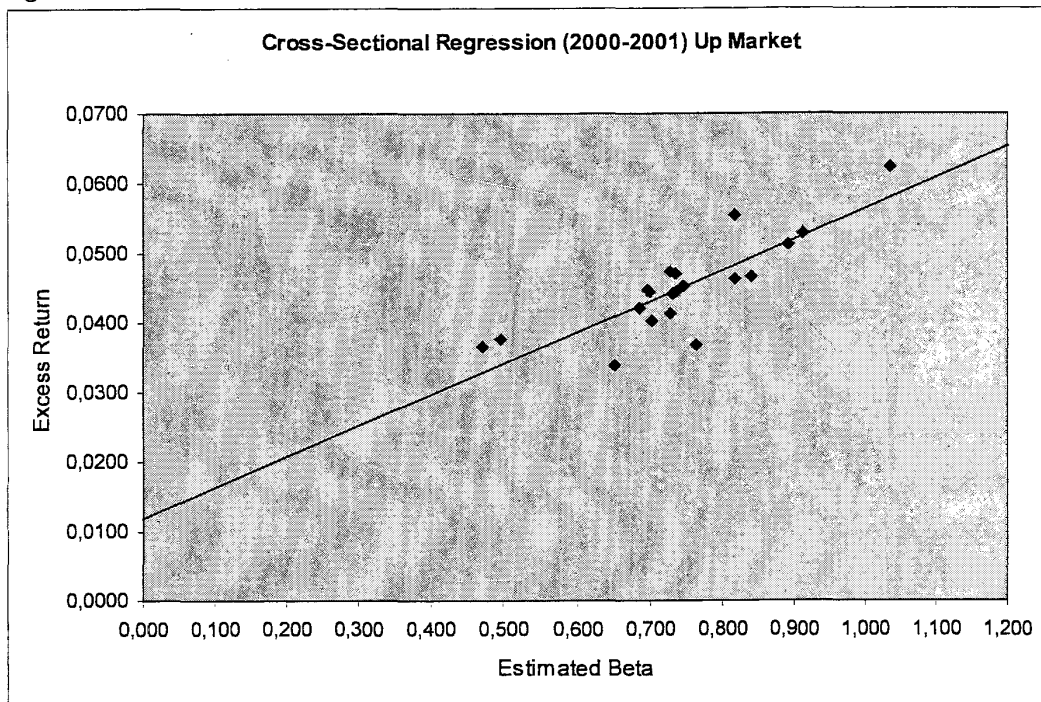
**Figure 5.6** Conditional test results for the up-market period of 1999-2000



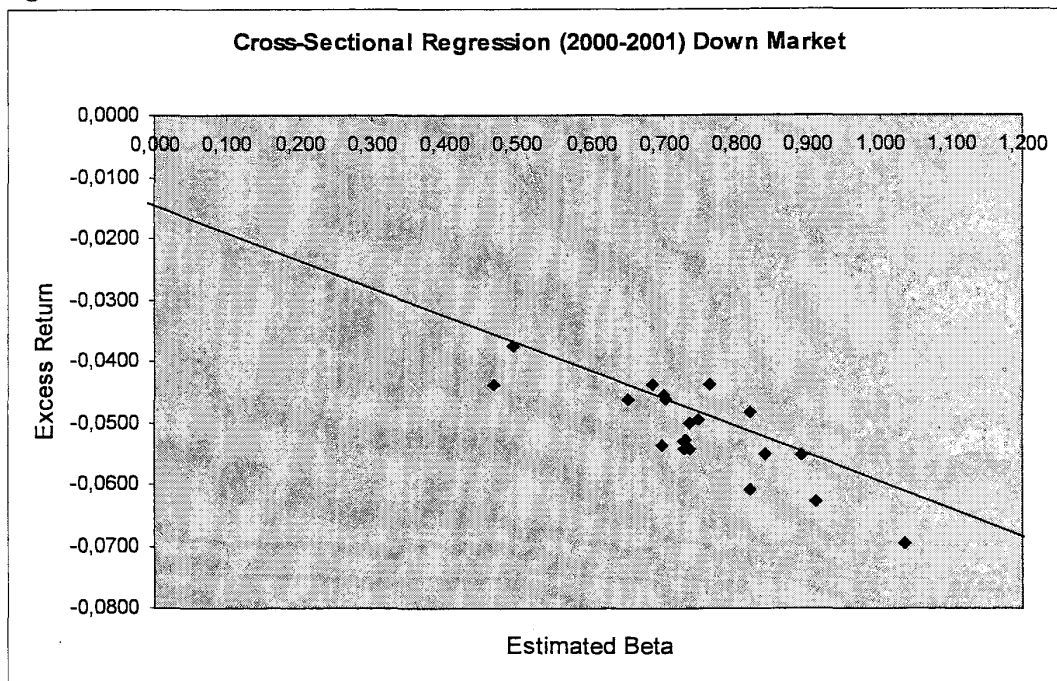
**Figure 5.7** Conditional test results for the down-market period of 1999-2000



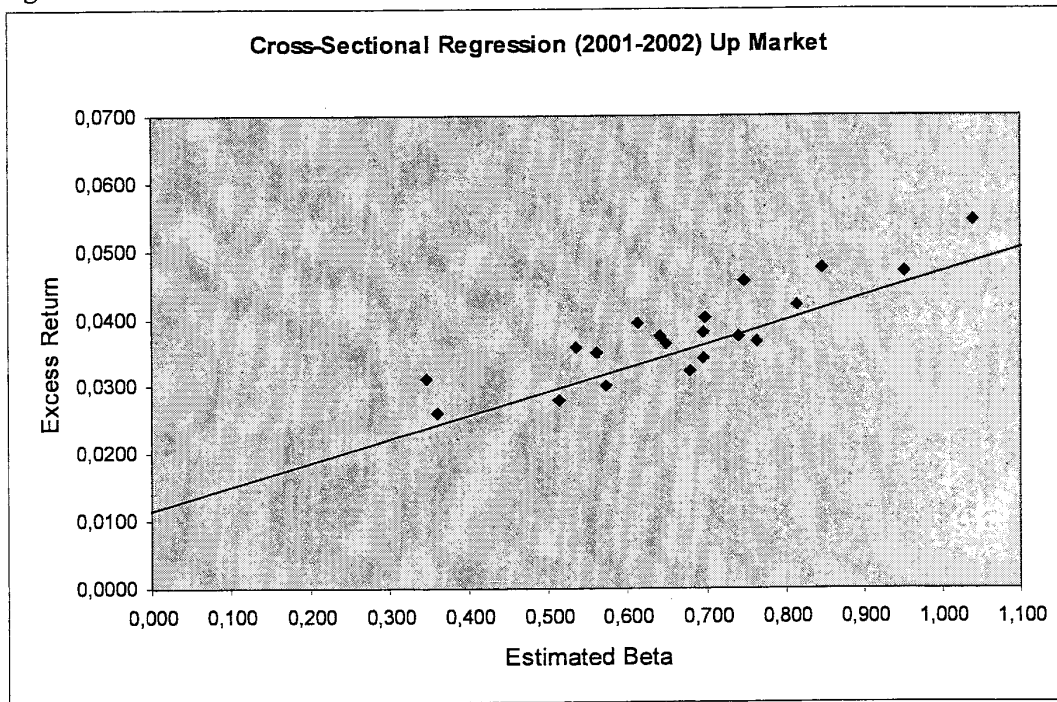
**Figure 5.8** Conditional test results for the up-market period of 2000-2001



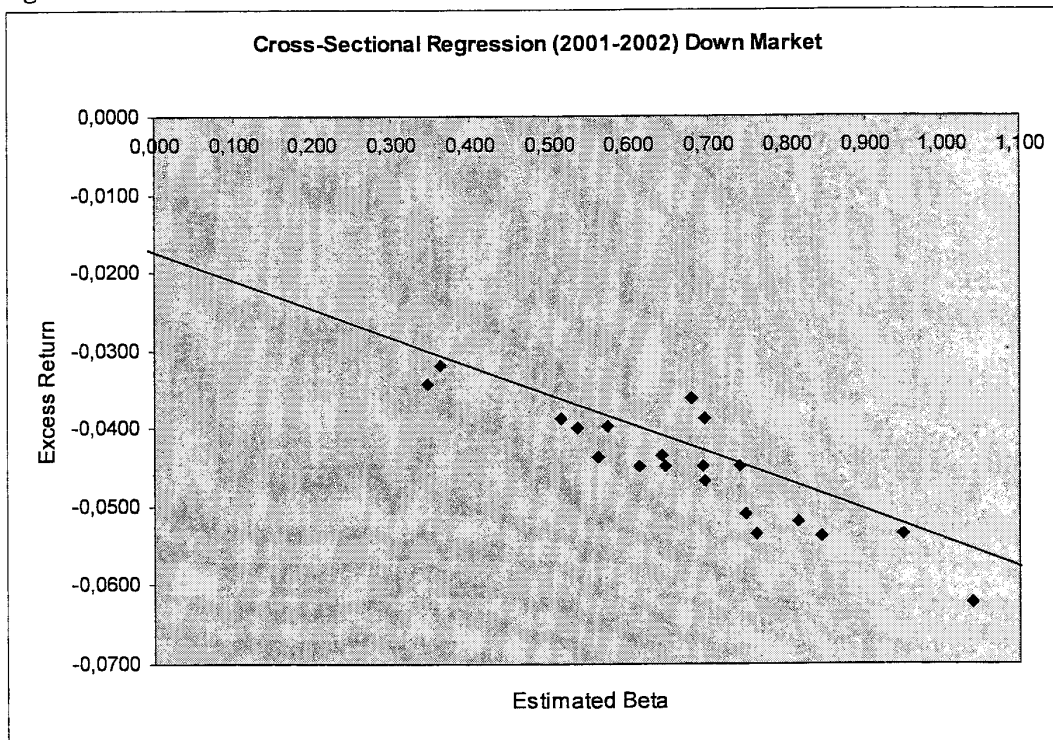
**Figure 5.9** Conditional test results for the down-market period of 2000-2001



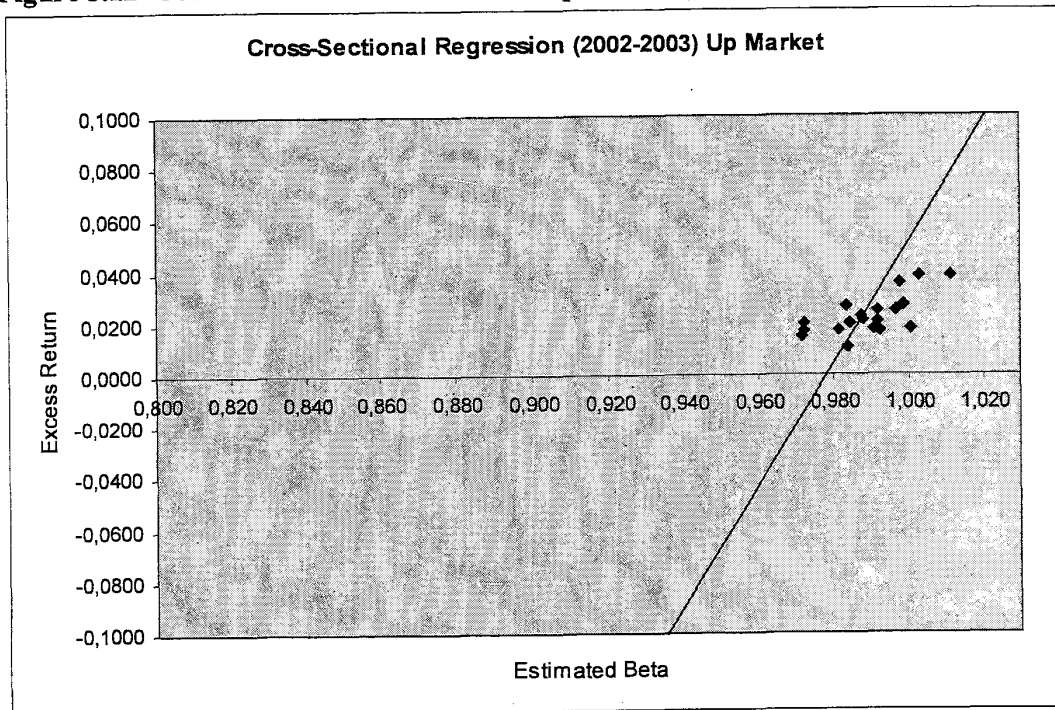
**Figure 5.10** Conditional test results for the up-market period of 2001-2002



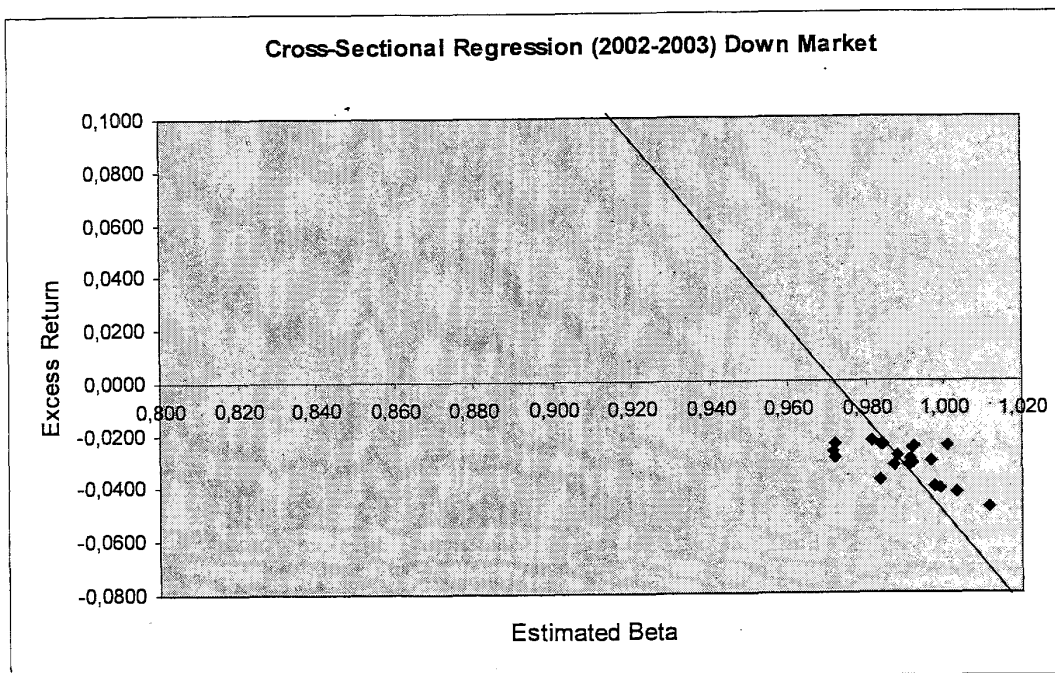
**Figure 5.11** Conditional test results for the down-market period of 2001-2002



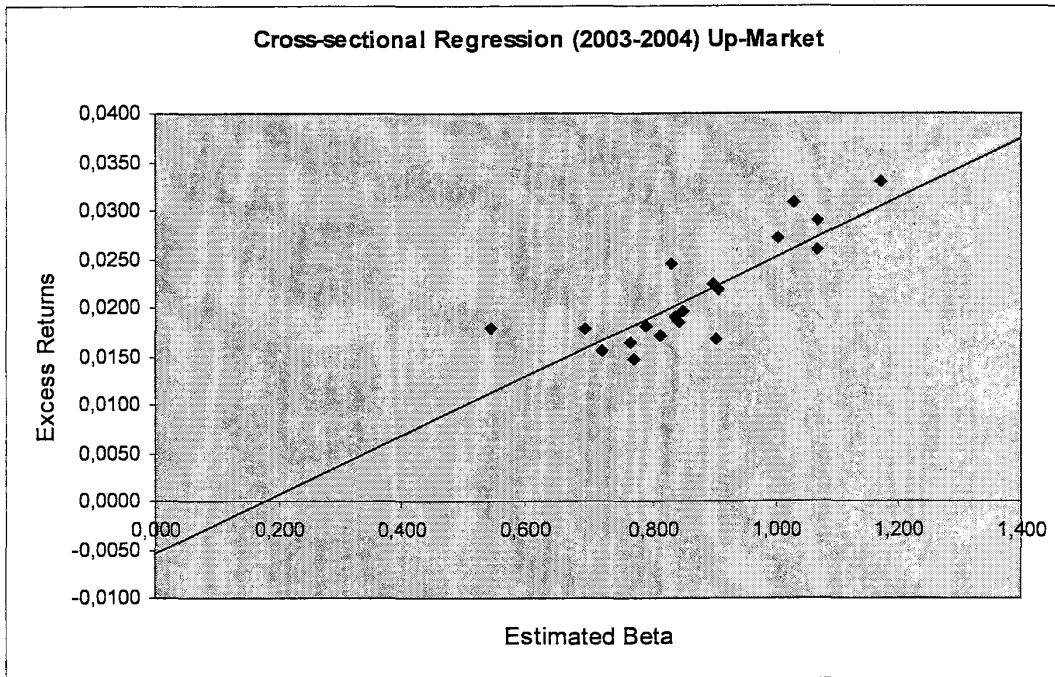
**Figure 5.12 Conditional test results for the up-market period of 2002-2003**



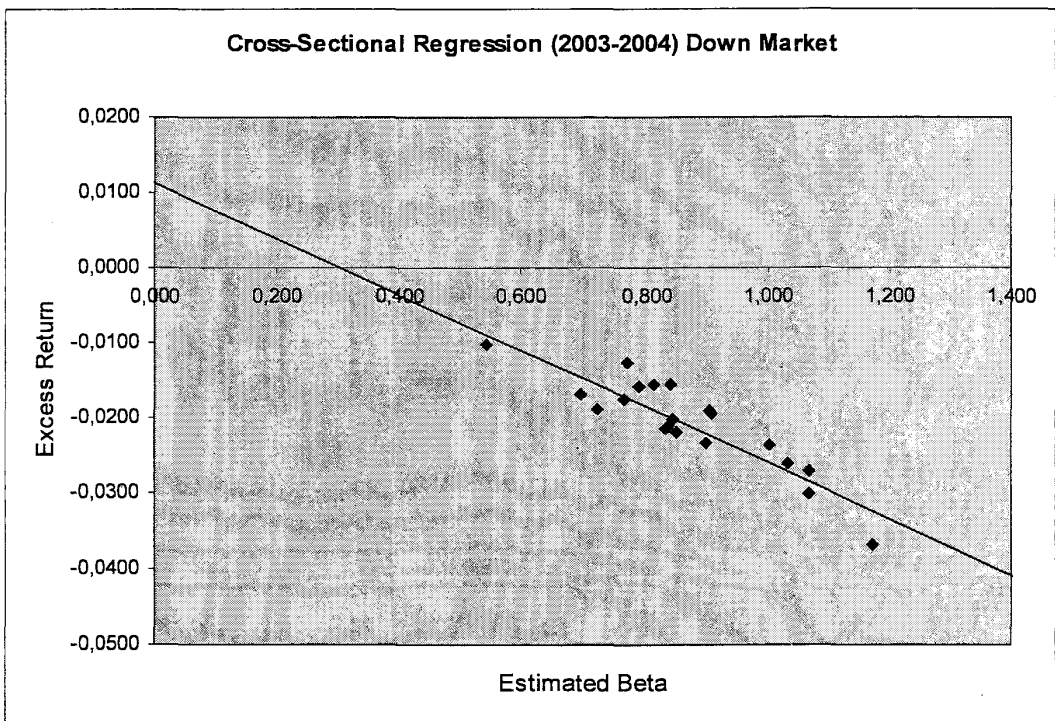
**Figure 5.13 Conditional test results for the down-market period of 2002-2003**



**Figure 5.14 Conditional test results for the up-market period of 2003-2004**



**Figure 5.15 Conditional test results for the down-market period of 2003-2004**





## 6. CONCLUSION

A number of previous studies testing systematic relationship between risk and return found weak and inconsistent results. However, the traditional test procedure based on the Fama and MacBeth (1973) approach includes a joint test of two hypotheses. The hypothesis of a relationship between beta and return and the hypothesis of a positive market risk premium are tested simultaneously. Pettigill et al. (1995) proposed a conditional test procedure, which allows to separately test the hypothesis of a relationship between beta and return. The conditional CAPM test, which is conducted using the ex-post data has the results that predict a conditional relationship between beta and return such that securities with higher beta have higher (lower) returns when the market risk premium is positive (negative).

In this study we have conducted two CAPM tests: traditional and conditional. It has been found that the CAPM can not be proven under the traditional test approach. The results show that in two of five subperiods the relationship between beta and return is inversely related, which does not support the hypothesis of positive market risk premium predicted by the CAPM. In other subperiods the results show positive but insignificant relationship between beta and return. However, when the conditional test has been applied, the statistically significant relationship between beta and return has been found in all subperiods. The market excess return analysis shows that in almost a half observations the market excess return is negative. We argue that the traditional test results have been biased due to the fact that the traditional test does not take into account the condition of positive and negative market excess returns. So, we have separated the up-market weeks and down-market week to conduct the conditional test. The results indicate that the portfolios with higher betas have higher returns when the market risk premium is positive and lower returns when the market risk premium is negative. It has been found that the relationship between beta and return is consistent across all subperiods in the sample. Therefore, we have arrived at the conclusion that the systematic conditional relationship exists between beta and returns.

Our test results support the hypothesis that the systematic risk of a security measured by its beta is indeed a relevant measure of risk and is one of the main explanatory powers on

security returns in Turkish capital market. Beta is reliably related to the security's return depending on the sign of the market risk premium. Hence the use of market beta estimated from the historical price data by portfolio managers hence seems to be justified. The contribution of this study to the investments management is the evidence provided for the role of beta in explaining returns in the Turkish capital market. Moreover, in the practical sense, the study also suggests that a portfolio of stocks may perform better when its funds are invested in high beta stocks during up market and in low beta stocks during down market.

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APPENDIX I

Portfolio Formation and Estimation for Test Period I (1995-2000)

PORTFOLIO FORMATION (1995-1996) Including 167 securities

Portfolio 1	Portfolio 2	Portfolio 3	Portfolio 4	Portfolio 5	Portfolio 6	Portfolio 7	Portfolio 8	Portfolio 9	Portfolio 10
Company Beta DOHOL 1,271 ANBRA 1,650 INTEB 1,517 MILYT 1,453 EGILC 1,425 MPAZ 1,415 PINSU 1,393 ATLAS 1,385 GLBYO 1,375	Company Beta MERKO 1,344 GUBRF 1,329 ECZYT 1,309 SNPAM 1,307 SONME 1,299 TUPRS 1,294 NTHOL 1,294 IKTFN 1,289 EVREN 1,288	Company Beta DISBA 1,287 ASELS 1,283 MMART 1,278 MAALT 1,276 DITAS 1,255 KORDS 1,254 ALTN 1,218 ARCLK 1,212 KARTN 1,198	Company Beta EGPRO 1,190 KONYA 1,185 VESTL 1,183 OKANT 1,180 TOFAS 1,177 THYAO 1,172 DEVA 1,168 TIRE 1,165 CELHA 1,162	Company Beta MAKTK 1,161 IZOCM 1,160 GENT5 1,156 AKPDP 1,153 MUTLU 1,152 DENCN 1,147 VKFTY 1,142 USAK 1,140	Company Beta KERVY 1,140 MRDIN 1,139 SAPAZ 1,136 BRISA 1,136 MYZYO 1,132 AYGAZ 1,129 BEKO 1,126 UNYEC 1,124	Company Beta PNET 1,121 EMBK 1,121 OLMKS 1,114 FACF 1,106 IZMDC 1,104 DUROF 1,104 KUTPO 1,098 FMZPZ 1,095	Company Beta EGGJIB 1,093 MEDYA 1,087 ABANA 1,086 KRTEK 1,083 FINBN 1,082 HURGZ 1,078 ADNAC 1,076 YAKFN 1,071	Company Beta GUNEY 1,071 EBEGL 1,069 TRKCM 1,064 YKBNK 1,061 PARSN 1,058 PETKM 1,056 EGEEN 1,055 EPLAS 1,035	Company Beta PROTO 1,034 PKENT 1,033 NTTUR 1,030 YASAS 1,029 RANTL 1,027 TKBNK 1,024 FKRL 1,023 PTOFS 1,019
Company Beta BSPRO 1,019 DOKTS 1,012 CBMTS 1,011 BAGFS 1,011 GOODY 1,011 DEMIR 1,009 AKALT 1,008 IHLAS 0,991	Company Beta TRNSK 0,987 SARKY 0,982 PRKAB 0,977 SABAH 0,975 BRSAN 0,974 TUDDF 0,971 ERBOS 0,969 SISE 0,953	Company Beta ANACM 0,947 RAKSE 0,943 BOLUC 0,936 DARDL 0,935 RKSEV 0,929 TUKAS 0,928 ALCAR 0,927 DERIM 0,927	Company Beta AKFN 0,926 HEKTS 0,926 KAVPA 0,920 EGSER 0,920 KLBMO 0,917 MARET 0,909 KCHOL 0,903 ADANA 0,891	Company Beta PNSUT 0,886 EMPAS 0,881 LUKSK 0,871 EDIP 0,861 BNJUN 0,861 BTCM 0,847 TEKST 0,844 GLMDE 0,843	Company Beta MIGRS 0,841 GUSGR 0,840 SKTAS 0,832 NICDE 0,824 PNUN 0,822 VKING 0,812 YUNSA 0,797 METAS 0,788	Company Beta FENIS 0,777 EMSAN 0,773 ECBRA 0,758 OZFN 0,751 ISCTR 0,750 MRSHL 0,744 NETAS 0,734	Company Beta ERCYS 0,721 TATKS 0,713 KOTKS 0,705 BURCE 0,702 TOASO 0,698 AFYON 0,697 FRIGO 0,692 CUKEL 0,660 PIMAS 0,659	Company Beta KEPEZ 0,657 CIMS A 0,647 GARAN 0,647 AKBNK 0,640 AYCES 0,639 GOLTS 0,631 ALARK 0,595 YKFN 0,554 AKSA 0,534	Company Beta TBORG 0,503 ISBTR 0,501 KENT 0,495 TSKB 0,452 ASLAN 0,437 CMNT 0,341 BU CIM 0,332 GIMMA 0,323 HLKSG 0,187

PORTFOLIO BETA RE-ESTIMATION (1997-1998) Including 167 securities

Portfolio 1	Portfolio 2	Portfolio 3	Portfolio 4	Portfolio 5	Portfolio 6	Portfolio 7	Portfolio 8	Portfolio 9	Portfolio 10
Company Beta DOHOL 1,124 ANBRA 0,726 INTEB 0,747 MILYT 0,975 EGILC 0,992 MPAZ 1,055 PINSU 0,695 ATLAS 0,836 GLBYO 0,644	Company Beta MERKO 0,477 GUBRF 0,911 ECZYT 1,093 SNPAM 0,708 SONME 0,689 TUPRS 1,288 NTHOL 1,143 IKTFN 1,052 EVREN 0,854	Company Beta DISBA 0,911 ASELS 1,015 MMART 0,685 MAALT 0,469 DITAS 0,582 KORDS 0,900 ALTN 0,643 ARCLK 1,102 KARTN 0,652	Company Beta EGPRO 0,862 KONYA 0,517 VESTL 1,078 OKANT 0,618 TOFAS 0,883 THYAO 1,007 DEVA 0,916 TIRE 0,558 CELHA 0,869	Company Beta MAKTK 0,927 IZOCM 0,715 GENTS 0,644 AKPDP 0,933 MUTLU 0,729 DENCN 0,804 VKFTY 0,877 USAK 0,848	Company Beta KERVY 0,657 MRDIN 0,851 SAPAZ 0,976 BRISA 0,867 MYZYO 0,733 AYGAZ 0,885 BEKO 0,986 UNYEC 0,719	Company Beta PNET 0,842 EMBK 0,780 OLMKS 0,756 FACF 0,666 IZMDC 1,055 DUROF 0,842 KUTPO 0,842 FMZPZ 0,724	Company Beta EGGJIB 0,671 MEDYA 0,848 ABANA 0,572 KRTEK 0,733 FINBN 0,875 HURGZ 0,906 ADNAC 0,676 YAKFN 0,859	Company Beta GUNEY 0,909 EBEGL 1,156 TRKCM 0,940 YKBNK 1,268 PARSN 0,617 PETKM 1,089 EGEEN 0,871 EPLAS 0,751	Company Beta PROTO 0,948 PKENT 0,454 NTTUR 1,002 YASAS 0,812 RANTL 0,849 TKBNK 0,854 FKRL 0,594 PTOFS 1,001
Company Beta BSPRO 0,749 DOKTS 0,898 CBMTS 0,850 BAGFS 1,031 GOODY 0,574 DEMIR 0,692 AKALT 0,714 IHLAS 0,902	Company Beta TRNSK 0,977 SARKY 0,906 PRKAB 0,723 SABAH 0,972 BRSAN 0,661 TUDDF 0,791 ERBOS 0,694 SISE 1,011	Company Beta ANACM 0,835 RAKSE 0,999 BOLUC 0,592 DARDL 0,794 RKSEV 0,963 TUKAS 0,616 ALCAR 0,889 DERIM 0,820	Company Beta AKFN 0,710 HEKTS 0,782 KAVPA 0,803 EGSER 0,785 KLBMO 0,807 MARET 0,723 KCHOL 1,098 ADANA 0,794	Company Beta PNSUT 0,962 EMPAS 0,478 LUKSK 0,679 EDIP 0,785 ENKA 1,117 BTCM 0,671 TEKST 0,753 GLMDE 1,294	Company Beta MIGRS 0,764 GUSGR 0,695 SKTAS 0,522 NICDE 0,652 PNUN 0,553 VKING 0,615 YUNSA 0,771 METAS 0,894	Company Beta FENIS 0,856 EMSAN 0,560 UCAK 0,585 ECBRA 0,672 OZFN 0,879 ISCTR 1,011 MRSHL 0,721 NETAS 0,789	Company Beta ERCYS 0,769 TATKS 0,846 KOTKS 0,470 BURCE 0,496 TOASO 0,987 AFYON 0,635 FRIGO 0,568 CUKEL 1,056 PIMAS 0,868	Company Beta KEPEZ 0,950 CIMS A 0,759 GARAN 1,217 AKBNK 0,988 AYCES 0,407 GOLTS 0,856 ALARK 0,974 YKFN 0,336 AKSA 0,760	Company Beta TBORG 0,762 ISBTR 0,552 KENT 0,672 TSKB 0,744 ASLAN 0,533 CMNT 0,404 BU CIM 0,468 GIMMA 0,903 HLKSG 0,167
Company Beta 0,801	Company Beta 0,842	Company Beta 0,814	Company Beta 0,813	Company Beta 0,792	Company Beta 0,683	Company Beta 0,759	Company Beta 0,768	Company Beta 0,950	Company Beta 0,814

**APPENDIX II**  
Portfolio Formation and Estimation for Test Period 2 (2000-2001)

PORTFOLIO FORMATION (1995-1996) including 187 securities																							
Portfolio 1		Portfolio 2		Portfolio 3		Portfolio 4		Portfolio 5		Portfolio 6		Portfolio 7		Portfolio 8		Portfolio 9		Portfolio 10					
Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta				
UACAK	1.469	ARCLK	1.149	AKBNK	1.063	KARTN	0.998	INTEM	0.941	AKIPD	0.895	UNYEC	0.835	FROTO	0.816	ESEMS	0.796	BAGFS	0.774	EMNIS	0.774		
FFKRL	1.332	TKBNK	1.139	ATSYO	1.058	TOASO	0.995	AYGAZ	0.938	DMRYO	0.895	AGIDA	0.833	ERBOS	0.811	BRASN	0.785	MERKO	0.771	BUCIM	0.419	EDIP	0.391
DOHOL	1.323	EYREN	1.130	VAKFN	1.054	PNASU	0.995	TRKCM	0.937	EPLAS	0.879	KAVPA	0.832	BRISA	0.811	YATAS	0.784	ALCAR	0.771	AVRSY	0.557	KOTKS	0.512
GLMDE	1.298	PETKM	1.129	KORDS	1.037	MLXYT	0.983	USAK	0.930	OLMKS	0.892	MRDIN	0.826	FINBN	0.783	FINBN	0.783	KBORG	0.770	ASLAN	0.505	ASLAN	0.505
THYAO	1.234	MPAZ	1.125	ASELS	1.027	MAKTK	0.981	DEVA	0.920	RAKSE	0.872	GEDIZ	0.826	ECYAP	0.780	KUPTO	0.769	CIMSA	0.801	BOSSA	0.765	YKFIN	0.292
EREGL	1.229	ISCTR	1.096	MYZYO	1.018	DIBSA	0.980	ENKA	0.920	ALNTF	0.869	DERIM	0.824	BTMCM	0.804	YKRYO	0.779	BOSSA	0.801	TIRE	0.764	YKFIN	0.292
NTHOL	1.213	PTOFS	1.077	GLBYO	1.018	ECZYT	0.970	BEKO	0.920	TOFAS	0.867	BUMYO	0.822	FMZP	0.776	BSPRO	0.776	BOSSA	0.801	TIRE	0.764	YKFIN	0.292
YKBNK	1.208	ATLAS	1.075	TRNSK	1.015	SISE	0.969	VESTL	0.919	FMZP	0.859	PRKAB	0.822	CELHA	0.798	MUTLU	0.776	BOSSA	0.801	TIRE	0.764	YKFIN	0.292
NITUR	1.180	ECILC	1.070	GARAN	1.003	SNPAM	0.960	DOKTS	0.899	HURGZ	0.843	GRNYO	0.820	NETAS	0.797	DITAS	0.776	BOSSA	0.801	TIRE	0.764	YKFIN	0.292
PORTFOLIO BETA RE-ESTIMATION (1998-1999) including 187 securities																							
Portfolio 1		Portfolio 2		Portfolio 3		Portfolio 4		Portfolio 5		Portfolio 6		Portfolio 7		Portfolio 8		Portfolio 9		Portfolio 10					
Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta				
UACAK	0.543	MMART	0.645	EGSER	0.713	KNFRT	0.697	ADNAC	0.669	BROYA	0.643	VKFTY	0.609	ARFYO	0.565	LUKSK	0.534	EMNIS	0.419	EDIP	0.391		
FFKRL	0.513	GOLTS	0.894	PINSU	0.616	MRSHL	0.602	GENSL	0.666	TUKAS	0.640	TEKST	0.608	AYCES	0.584	KAPLM	0.533	BUCIM	0.419	EDIP	0.391		
EGEBN	0.487	ALTN	0.777	VKING	0.679	YUNSA	0.688	ADSL	0.660	NGDE	0.635	CBSRO	0.603	BURCE	0.561	ISYAT	0.514	BRASN	0.488	ALCAR	0.470		
ANACM	0.488	KLBMO	0.728	OKANT	0.705	KERTV	0.688	YASAS	0.659	KRTEK	0.633	PKENT	0.602	AKSA	0.559	BHMBN	0.512	YATAS	0.486	ALCAR	0.470		
DENCM	0.749	CEMTS	0.724	ADBRG	0.705	SONME	0.686	FRGGO	0.658	DUROF	0.633	ABANA	0.588	ALARK	0.555	BISAS	0.505	FINBN	0.486	ALCAR	0.470		
ADANA	0.748	KONYA	0.721	TATKS	0.700	GOODY	0.682	DAKDL	0.655	UKIM	0.652	TNSAS	0.586	MGRS	0.555	AFYON	0.501	ASLAN	0.505	ASLAN	0.505		
FNSYO	0.740	BOLUC	0.719	PIMAS	0.699	GUSGR	0.675	HEKTS	0.655	MARET	0.628	OTKAR	0.583	AKFIN	0.551	MAALT	0.492	YKFIN	0.260	YKFIN	0.260		
GIMMA	0.739	RANIL	0.718	PAKSN	0.698	BRVAT	0.674	EGPRO	0.650	ISBTR	0.627	OTKAR	0.579	TSKB	0.548	SKTAS	0.444	PRTAS	0.251	PRTAS	0.251		
IZOCM	0.736	AKCNS	0.716	TUDDF	0.697	KENT	0.671	CARSI	0.650	OZFIN	0.609	EGGUB	0.568	HZNDR	0.544	TEZSN	0.437	BANVT	0.216	BANVT	0.216		
PORTFOLIO BETA RE-ESTIMATION (1998-1999) including 187 securities																							
Portfolio 1		Portfolio 2		Portfolio 3		Portfolio 4		Portfolio 5		Portfolio 6		Portfolio 7		Portfolio 8		Portfolio 9		Portfolio 10					
Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta				
UACAK	0.543	MMART	0.645	EGSER	0.713	KNFRT	0.697	ADNAC	0.669	BROYA	0.643	VKFTY	0.609	ARFYO	0.565	LUKSK	0.534	EMNIS	0.419	EDIP	0.391		
FFKRL	0.513	GOLTS	0.894	PINSU	0.616	MRSHL	0.602	GENSL	0.666	TUKAS	0.640	TEKST	0.608	AYCES	0.584	KAPLM	0.533	BUCIM	0.419	EDIP	0.391		
EGEBN	0.487	ALTN	0.777	VKING	0.679	YUNSA	0.688	ADSL	0.660	NGDE	0.635	CBSRO	0.603	BURCE	0.561	ISYAT	0.514	BRASN	0.488	ALCAR	0.470		
ANACM	0.488	KLBMO	0.728	OKANT	0.705	KERTV	0.688	YASAS	0.659	KRTEK	0.633	PKENT	0.602	AKSA	0.559	BHMBN	0.512	YATAS	0.486	ALCAR	0.470		
DENCM	0.749	CEMTS	0.724	ADBRG	0.705	SONME	0.686	FRGGO	0.658	DUROF	0.633	ABANA	0.588	ALARK	0.555	BISAS	0.505	FINBN	0.486	ALCAR	0.470		
ADANA	0.748	KONYA	0.721	TATKS	0.700	GOODY	0.682	DAKDL	0.655	UKIM	0.652	TNSAS	0.586	MGRS	0.555	AFYON	0.501	ASLAN	0.505	ASLAN	0.505		
FNSYO	0.740	BOLUC	0.719	PIMAS	0.699	GUSGR	0.675	HEKTS	0.655	MARET	0.628	OTKAR	0.583	AKFIN	0.551	MAALT	0.492	YKFIN	0.260	YKFIN	0.260		
GIMMA	0.739	RANIL	0.718	PAKSN	0.698	BRVAT	0.674	EGPRO	0.650	ISBTR	0.627	OTKAR	0.579	TSKB	0.548	SKTAS	0.444	PRTAS	0.251	PRTAS	0.251		
IZOCM	0.736	AKCNS	0.716	TUDDF	0.697	KENT	0.671	CARSI	0.650	OZFIN	0.609	EGGUB	0.568	HZNDR	0.544	TEZSN	0.437	BANVT	0.216	BANVT	0.216		
PORTFOLIO BETA RE-ESTIMATION (1998-1999) including 187 securities																							
Portfolio 1		Portfolio 2		Portfolio 3		Portfolio 4		Portfolio 5		Portfolio 6		Portfolio 7		Portfolio 8		Portfolio 9		Portfolio 10					
Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta				
UACAK	0.543	MMART	0.645	EGSER	0.713	KNFRT	0.697	ADNAC	0.669	BROYA	0.643	VKFTY	0.609	ARFYO	0.565	LUKSK	0.534	EMNIS	0.419	EDIP	0.391		
FFKRL	0.513	GOLTS	0.894	PINSU	0.616	MRSHL	0.602	GENSL	0.666	TUKAS	0.640	TEKST	0.608	AYCES	0.584	KAPLM	0.533	BUCIM	0.419	EDIP	0.391		
EGEBN	0.487	ALTN	0.777	VKING	0.679	YUNSA	0.688	ADSL	0.660	NGDE	0.635	CBSRO	0.603	BURCE	0.561	ISYAT	0.514	BRASN	0.488	ALCAR	0.470		
ANACM	0.488	KLBMO	0.728	OKANT	0.705	KERTV	0.688	YASAS	0.659	KRTEK	0.633	PKENT	0.602	AKSA	0.559	BHMBN	0.512	YATAS	0.486	ALCAR	0.470		
DENCM	0.749	CEMTS	0.724	ADBRG	0.705	SONME	0.686	FRGGO	0.658	DUROF	0.633	ABANA	0.588	ALARK	0.555	BISAS	0.505	FINBN	0.486	ALCAR	0.470		
ADANA	0.748	KONYA	0.721	TATKS	0.700	GOODY	0.682	DAKDL	0.655	UKIM	0.652	TNSAS	0.586	MGRS	0.555	AFYON	0.501	ASLAN	0.505	ASLAN	0.505		
FNSYO	0.740	BOLUC	0.719	PIMAS	0.699	GUSGR	0.675	HEKTS	0.655	MARET	0.628	OTKAR	0.583	AKFIN	0.551	MAALT	0.492	YKFIN	0.260	YKFIN	0.260		
GIMMA	0.739	RANIL	0.718	PAKSN	0.698	BRVAT	0.674	EGPRO	0.650	ISBTR	0.627	OTKAR	0.579	TSKB	0.548	SKTAS	0.444	PRTAS	0.251	PRTAS	0.251		
IZOCM	0.736	AKCNS	0.716	TUDDF	0.697	KENT	0.671	CARSI	0.650	OZFIN	0.609	EGGUB	0.568	HZNDR	0.544	TEZSN	0.437	BANVT	0.216	BANVT	0.216		

**APPENDIX III**  
**Portfolio Formation and Estimation for Test Period 3 (2001-2002)**

Portfolio 1		Portfolio 2		Portfolio 3		Portfolio 4		Portfolio 5		Portfolio 6		Portfolio 7		Portfolio 8		Portfolio 9		Portfolio 10			
Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta		
GLMDE	1.294	KCHOL	1.028	THYAO	1.007	MILYT	0.975	MAKTK	0.927	AYGAZ	0.885	GOLTS	0.856	TATKS	0.846	AGDA	0.804	HEKTS	0.782		
YKBNK	1.268	ECZYT	1.093	NTLUR	1.002	ALARK	0.974	DEVA	0.916	VKGYO	0.871	FENIS	0.856	KUTPO	0.842	DENCM	0.804	BROVA	0.781		
AKGRT	1.232	PEYTK	1.089	PTOFS	1.001	ALCTI	0.970	AKNS	0.881	AKNS	0.877	TKBNK	0.854	ATLAS	0.836	KAVPA	0.803	YUNSA	0.771		
ASZUJ	1.221	VESTL	1.078	RAKSE	0.999	RKSEV	0.963	DIEBA	0.911	FRBNB	0.875	EYREN	0.854	ANACM	0.835	DMSAS	0.798	MGRS	0.764		
GARAN	1.217	IZMDC	1.059	ECILC	0.992	PNSTU	0.962	GUBRF	0.911	EGBEN	0.881	YATAS	0.851	ATSDY	0.832	ADANA	0.794	TKRKG	0.762		
EREGL	1.156	AKBNK	0.988	DOKTS	0.951	SARKY	0.906	HURKZ	0.869	MRDIN	0.831	ADBL	0.825	TNSAS	0.792	AKSA	0.760	AKSA	0.760		
NTHOL	1.143	TOASO	0.987	PROTO	0.948	HURKZ	0.906	PIMAS	0.868	RAYSG	0.849	GRNYO	0.821	TUDDF	0.791	CMSA	0.759	CMSA	0.759		
DOHOL	1.124	BEKO	0.986	TRKCM	0.940	KORDS	0.900	VAKFN	0.859	RANTL	0.849	DERIM	0.820	NETAS	0.789	OLMKS	0.756	OLMKS	0.756		
ARCLK	1.102	ISCTR	0.985	SAHOL	0.940	GMA	0.893	BRISA	0.859	CEYLN	0.848	KLBMO	0.807	EGSER	0.785	EMKEL	0.753	EMKEL	0.753		
		TRNSK	0.977	AKPDP	0.913	ALCAR	0.889														

Portfolio 11		Portfolio 12		Portfolio 13		Portfolio 14		Portfolio 15		Portfolio 16		Portfolio 17		Portfolio 18		Portfolio 19		Portfolio 20			
Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta		
ECYAP	0.732	EPLAS	0.732	ALGYO	0.719	GLSGR	0.695	LKSKC	0.654	AFYON	0.654	ATEKS	0.636	ATEKS	0.636	ATEKS	0.636	ATEKS	0.636	ATEKS	0.636
BERDN	0.751	ESDMS	0.730	UNVEE	0.719	PNSTU	0.695	ADNAC	0.654	KNKRT	0.679	KNKRT	0.679	KNKRT	0.679	KNKRT	0.679	KNKRT	0.679	KNKRT	0.679
BSPRO	0.749	ANSGR	0.730	IZOCM	0.715	ERBOS	0.695	EMNS	0.654	KARTN	0.679	KARTN	0.679	KARTN	0.679	KARTN	0.679	KARTN	0.679	KARTN	0.679
INTEM	0.747	EGPRO	0.730	AKALT	0.714	DURDF	0.691	KENT	0.651	NIGDE	0.673	NIGDE	0.673	NIGDE	0.673	NIGDE	0.673	NIGDE	0.673	NIGDE	0.673
COMJN	0.745	MUTLU	0.729	AKTIN	0.711	FNSYO	0.690	BTGCM	0.651	DMRYO	0.672	DMRYO	0.672	DMRYO	0.672	DMRYO	0.672	DMRYO	0.672	DMRYO	0.672
TSKB	0.744	FMLZP	0.724	BRVAT	0.709	SONME	0.671	EGGUB	0.631	CARSI	0.671	CARSI	0.671	CARSI	0.671	CARSI	0.671	CARSI	0.671	CARSI	0.671
BPREN	0.734	MARET	0.723	KAPLM	0.708	ADBRG	0.686	CLBHI	0.644	YKRYO	0.671	YKRYO	0.671	YKRYO	0.671	YKRYO	0.671	YKRYO	0.671	YKRYO	0.671
KRTEK	0.733	PRKAB	0.723	SNPAM	0.708	GEDUZ	0.686	HZNDR	0.644	GENIS	0.668	GENIS	0.668	GENIS	0.668	GENIS	0.668	GENIS	0.668	GENIS	0.668
MYZYO	0.733	MRSHL	0.721	ARYYO	0.706	MMART	0.685	BRNSN	0.641	ALTN	0.661	ALTN	0.661	ALTN	0.661	ALTN	0.661	ALTN	0.661	ALTN	0.661
BOSSA	0.732	OTKAR	0.724	OZFIN	0.706	AKPDP	0.685	BRNSN	0.641	ALTN	0.661	ALTN	0.661	ALTN	0.661	ALTN	0.661	ALTN	0.661	ALTN	0.661

APPENDIX IV  
Portfolio Formation and Estimation for Test Period 4 (2002-2003)

PORTFOLIO FORMATION (1998-1999) including 206 securities

Portfolio 1		Portfolio 2		Portfolio 3		Portfolio 4		Portfolio 5		Portfolio 6		Portfolio 7		Portfolio 8		Portfolio 9		Portfolio 10					
Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta				
YKBNK	1.218	ALARK	1.059	THYAO	0.956	GUBRF	0.915	TNSAS	0.878	FVREN	0.857	FNIS	0.833	KLASN	0.806	ZMDC	0.791	ECYAP	0.771	UNSTAR	0.768		
GLMDE	1.197	GARAN	1.055	BERGL	0.993	AKCPD	0.903	AYGAZ	0.978	RKSEV	0.956	DOKTS	0.933	CHASA	0.903	KTEJK	0.789	MRDYN	0.768	GRNYO	0.768	MGRS	0.765
AKCRT	1.136	PETKM	1.055	PRKTE	0.992	SASA	0.900	RAYSG	0.976	KMSY	0.954	ALCAR	0.929	BRVAT	0.903	EGGUB	0.784	GRNYO	0.768	MGRS	0.765	GRNYO	0.768
DOHOL	1.133	NETAS	1.049	ISSE	0.981	UZEL	0.896	KAVPA	0.966	TRNSK	0.945	DEVA	0.927	BOSSA	0.901	KSTL	0.780	GRNYO	0.768	MGRS	0.765	GRNYO	0.768
ECZYT	1.132	BAGES	1.044	TUPRS	0.975	GOLTS	0.894	KAVPA	0.966	TRNSK	0.945	DEVA	0.927	BOSSA	0.901	KSTL	0.780	GRNYO	0.768	MGRS	0.765	GRNYO	0.768
VESTI	1.115	ISCTR	1.043	ASELS	0.967	OTIKAR	0.893	GMA	0.942	ANACM	0.948	AKESA	0.920	ALNTE	0.801	YKGYO	0.778	SARKY	0.757	KLEMO	0.754	SARKY	0.757
NTHOL	1.111	ARCLK	1.038	SAHOL	0.958	BRISA	0.890	TUDDF	0.968	AKESA	0.948	HEKTS	0.918	TRKCM	0.801	ZOCM	0.777	BRVYA	0.754	BRVYA	0.754	BRVYA	0.754
ALCTL	1.084	PROTO	1.032	FINBN	0.957	EGEEN	0.887	ATLAS	0.966	CELHA	0.943	PRMAS	0.916	ADANA	0.793	HSPRO	0.775	TSKH	0.741	ANSGR	0.740	ANSGR	0.740
KCHOL	1.062	HURCZ	1.022	MIPAZ	0.936	YATAS	0.886	NTLUR	0.966	CELHA	0.943	PRMAS	0.916	ADANA	0.793	HSPRO	0.775	TSKH	0.741	ANSGR	0.740	ANSGR	0.740
TOASO	1.063	AKBNK	1.008	MLYT	0.926	MAKTK	0.886	RAKSE	0.957	CELHA	0.943	PRMAS	0.916	ADANA	0.793	HSPRO	0.775	TSKH	0.741	ANSGR	0.740	ANSGR	0.740
ASUZU	1.072	ECULC	0.931	BEKD	0.931	MAKTK	0.886	RAKSE	0.957	CELHA	0.943	PRMAS	0.916	ADANA	0.793	HSPRO	0.775	TSKH	0.741	ANSGR	0.740	ANSGR	0.740

PORTFOLIO BETA RE-ESTIMATION (2000-2001) including 206 securities

Portfolio 11		Portfolio 12		Portfolio 13		Portfolio 14		Portfolio 15		Portfolio 16		Portfolio 17		Portfolio 18		Portfolio 19		Portfolio 20			
Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta		
YKBNK	1.027	ALARK	1.010	THYAO	1.021	GUBRF	0.963	TNSAS	1.027	FVREN	1.014	FNIS	0.978	KLASN	0.948	ZMDC	1.002	ECYAP	0.978	UNSTAR	0.963
GLMDE	1.026	GARAN	0.971	BERGL	1.006	AKCPD	0.980	AYGAZ	1.007	RKSEV	0.977	DOKTS	0.978	CHASA	0.925	MRDYN	0.989	GRNYO	0.978	MGRS	0.991
AKCRT	1.029	PETKM	0.983	PRKTE	1.006	SASA	0.982	RAYSG	0.997	CHMTS	0.977	ALCAR	0.982	BRVAT	0.924	EGGUB	0.995	GRNYO	0.978	MGRS	0.991
DOHOL	1.014	NETAS	1.042	ISSE	1.004	UZEL	0.988	KAVPA	0.990	TRNSK	0.977	DEVA	1.007	BOSSA	0.950	KSTL	0.956	GRNYO	0.978	MGRS	0.991
ECZYT	1.014	BAGES	0.989	TUPRS	0.961	GOLTS	0.955	TATES	0.972	INTEK	0.992	DEVA	1.007	BOSSA	0.950	KSTL	0.956	GRNYO	0.978	MGRS	0.991
VESTI	1.006	ISCTR	1.030	ASELS	0.985	OTIKAR	0.983	GMA	0.989	ANACM	1.009	AKESA	0.927	CARS	0.924	YKGYO	1.004	SARKY	0.980	KLEMO	0.991
NTHOL	1.024	ARCLK	0.979	SAHOL	0.985	BRISA	0.987	TUDDF	0.981	KORDS	1.002	HEKTS	1.001	TRKCM	1.001	ZOCM	0.992	BRVYA	1.029	SARKY	0.980
ALCTL	0.999	PROTO	1.024	FINBN	0.999	EGEEN	0.951	ATLAS	1.014	CULHA	0.972	PRMAS	1.014	ALGVO	0.985	ALJTN	1.008	BRVYA	1.029	SARKY	0.980
KCHOL	1.000	HURCZ	0.986	MIPAZ	1.003	YATAS	0.967	NTLUR	1.033	AKCNS	0.986	KUTA	0.944	ADANA	0.973	GUSGR	0.999	ANSGR	1.004	ANSGR	1.004
TOASO	0.990	AKBNK	0.993	MLYT	0.975	MAKTK	0.966	RAKSE	0.982	TRBNK	0.986	KUTYO	0.992	ATSYO	0.972	DISBA	0.967	TOPFN	0.994	TOPFN	0.994
ASUZU	1.009	ECULC	1.012	BEKO	0.963	MAKTK	0.966	RAKSE	0.982	TRBNK	0.986	KUTYO	0.992	ATSYO	0.972	DISBA	0.967	TOPFN	0.994	TOPFN	0.994

PORTFOLIO BETA RE-ESTIMATION (2000-2001) including 206 securities

Portfolio 11		Portfolio 12		Portfolio 13		Portfolio 14		Portfolio 15		Portfolio 16		Portfolio 17		Portfolio 18		Portfolio 19		Portfolio 20	
Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta
YKBNK	0.976	TEKST	1.009	MUTLU	1.043	MARET	1.000	ESEMS	0.972	EMNS	1.024	AGIBA	1.013	EPJAS	0.979	EPJAS	0.979	UNSTAR	0.944
GLMDE	0.972	BRSAN	1.009	BEREN	0.953	DIENCM	0.983	ERBDS	0.972	ILMKS	1.025	MRSH	1.010	SKTAS	0.986	SKTAS	0.986	AYCES	0.941
SONME	0.994	FNSYO	0.941	YUNSA	0.971	VEDZ	0.996	FMPZP	0.957	ISYAT	1.004	PTOF	0.969	KENT	0.967	KENT	0.967	AYCES	0.941
ISFR	0.994	MZHL	0.999	BANVT	1.008	VKING	1.056	KNFRT	0.945	BKCFE	0.980	GLBYO	1.011	EGPRO	0.978	EGPRO	0.978	AYCES	0.941
USAK	1.008	KERVY	0.961	ARFYO	0.987	ATERS	0.976	MMART	0.996	SNPAM	1.008	ABANA	0.964	CRSBO	0.978	CRSBO	0.978	AYCES	0.941
AKALI	0.962	EGSER	1.027	ADNAC	0.987	GENTS	0.956	MYZYO	0.964	PRSU	0.981	DIUROF	0.979	NIGDF	0.991	NIGDF	0.991	AYCES	0.941
LIKSK	0.984	PRKAB	0.971	CBMTN	1.003	BUMYO	0.956	TBRG	0.958	TIRE	1.011	DIUROF	0.979	UCAK	0.964	UCAK	0.964	AYCES	0.941
OKANT	1.024	CLEBI	0.987	PARSN	0.934	YKRYO	0.983	ADBRG	0.978	TACYO	0.954	UNYEC	0.967	KOTKS	0.910	KOTKS	0.910	AYCES	0.941
TPFAC	0.956	COMUN	0.975	OZFN	0.934	DERIM	1.027	GOODY	0.997	TACYO	0.954	UNYEC	0.967	UNYEC	0.967	UNYEC	0.967	AYCES	0.941
ARAT	0.980	ADEL	1.036	VAKEN	0.979	DERIM	1.027	GOODY	0.997	TACYO	0.954	UNYEC	0.967	UNYEC	0.967	UNYEC	0.967	AYCES	0.941

APPENDIX V

Portfolio Formation and Estimation for Test Period 5 (2003-2004)

PORTFOLIO FORMATION (1999-2000) Including 225 securities

Portfolio 1		Portfolio 2		Portfolio 3		Portfolio 4		Portfolio 5		Portfolio 6		Portfolio 7		Portfolio 8		Portfolio 9		Portfolio 10			
Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta		
TUPRS	0.747	MEMSA	0.780	BURCE	0.664	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
RAYSG	0.933	WRE	0.746	ALSD	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
GOODY	0.699	WRS	0.746	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
DITAS	0.742	AKGT	0.669	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
SASA	0.716	AKGT	0.669	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
MERKO	0.731	AKGT	0.669	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
CLEBI	0.719	AKGT	0.669	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
KRSTL	0.728	AKGT	0.669	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
ENRNS	0.728	AKGT	0.669	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
VAKFN	0.728	AKGT	0.669	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
ECZYT	0.895	SAHOL	1.076	PEIKM	0.927	ANACM	0.933	BRSAN	0.909	ECYAP	0.970	PREAS	0.835	CRSBO	0.807	YKRYO	0.778	VKFTY	0.746	YKRYO	0.778
PORTFOLIO BETA	1.175	PORTFOLIO BETA	1.058	PORTFOLIO BETA	1.070	PORTFOLIO BETA	1.032	PORTFOLIO BETA	1.003	PORTFOLIO BETA	0.830	PORTFOLIO BETA	0.837	PORTFOLIO BETA	0.896	PORTFOLIO BETA	0.907	PORTFOLIO BETA	0.971	PORTFOLIO BETA	0.941

PORTFOLIO BETA RE-ESTIMATION (2001-2003) Including 225 securities

Portfolio 1		Portfolio 2		Portfolio 3		Portfolio 4		Portfolio 5		Portfolio 6		Portfolio 7		Portfolio 8		Portfolio 9		Portfolio 10			
Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta	Company	Beta		
TUPRS	0.747	MEMSA	0.780	BURCE	0.664	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
RAYSG	0.933	WRE	0.746	ALSD	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
GOODY	0.699	WRS	0.746	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
DITAS	0.742	AKGT	0.669	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
SASA	0.716	AKGT	0.669	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
MERKO	0.731	AKGT	0.669	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
CLEBI	0.719	AKGT	0.669	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
KRSTL	0.728	AKGT	0.669	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
ENRNS	0.728	AKGT	0.669	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
VAKFN	0.728	AKGT	0.669	ENR	0.777	FNSYO	0.727	AVRSY	0.768	UKRM	0.590	KNFRY	0.406	BHMEN	0.310	FRKRL	0.731	BRMEN	0.521	FRKRL	0.731
ECZYT	0.895	SAHOL	1.076	PEIKM	0.927	ANACM	0.933	BRSAN	0.909	ECYAP	0.970	PREAS	0.835	CRSBO	0.807	YKRYO	0.778	VKFTY	0.746	YKRYO	0.778
PORTFOLIO BETA	1.175	PORTFOLIO BETA	1.058	PORTFOLIO BETA	1.070	PORTFOLIO BETA	1.032	PORTFOLIO BETA	1.003	PORTFOLIO BETA	0.830	PORTFOLIO BETA	0.837	PORTFOLIO BETA	0.896	PORTFOLIO BETA	0.907	PORTFOLIO BETA	0.971	PORTFOLIO BETA	0.941









### Appendix VII: Weekly Portfolio Returns and Market Risk Premium (Test Period 2000-2001)

Date	01.07.00	01.14.00	01.21.00	01.28.00	02.04.00	02.11.00	02.18.00	02.25.00	03.03.00	03.10.00	03.17.00	03.24.00	03.31.00	04.07.00	04.14.00	04.21.00	04.28.00	05.05.00	05.12.00
T-Bill (Inv. Adj)	0.01388226	0.01384319	0.01387911	0.01389221	0.01411788	0.01411462	0.01410156	0.01414144	0.01378089	0.01379635	0.000000	0.01382541	0.01380631	0.01390177	0.01299231	0.01298286	0.01297341	0.01297377	0.01281245
ISE100 Return	0.165215	-0.119154	0.051626	-0.074326	-0.078191	-0.015154	0.016435	0.071962	0.061115	0.000000	-0.008691	-0.104240	0.059781	0.083209	-0.078191	0.122826	-0.060063	-0.029666	
Rm - RF	0.151372	-0.115833	0.037734	-0.088444	-0.092533	-0.029236	0.002291	0.058173	0.047319	0.000000	-0.023516	-0.180406	0.046779	0.070217	-0.091173	0.109852	-0.072860	-0.042338	

Portfolio	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	
T-Bill (Inv. Adj)	0.220163	-0.124960	0.055260	-0.064708	-0.102545	-0.021709	0.009698	0.026000	0.000000	0.000000	0.000000	0.048337	-0.107618	0.063101	0.046737	-0.058362	0.086492	-0.029585	-0.017582		
ISE100 Return	0.102640	-0.124895	0.075948	-0.061946	-0.077526	0.011003	0.009654	0.126103	0.058555	0.000000	-0.027146	-0.079351	0.079590	0.049659	0.073040	-0.020708	0.165712	-0.045258	-0.016888		
Rm - RF	0.187203	-0.049172	0.089790	-0.079504	-0.133529	-0.036616	0.054855	0.020490	-0.027146	0.000000	-0.036583	-0.095719	0.049659	0.073040	0.073040	-0.020708	0.165712	-0.045258	-0.016888		
1		0.000000																			
2			0.000000																		
3				0.000000																	
4					0.000000																
5						0.000000															
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18																			0.000000		
19																				0.000000	
20																					0.000000

Date	05.19.00	05.26.00	06.02.00	06.09.00	06.16.00	06.23.00	06.30.00	07.07.00	07.14.00	07.21.00	07.28.00	08.04.00	08.11.00	08.18.00	08.25.00	09.01.00	09.08.00	09.15.00	09.22.00
T-Bill (Inv. Adj)	0.01288809	0.01286675	0.01198347	0.01198322	0.01156439	0.01156432	0.01156432	0.01156432	0.01156432	0.01156432	0.01156432	0.01156432	0.01156432	0.01156432	0.01156432	0.01156432	0.01156432	0.01156432	0.01156432
ISE100 Return	-0.065198	-0.032562	0.036856	-0.003561	-0.137166	0.039584	-0.031622	-0.008487	-0.059019	0.003873	0.038138	-0.024444	-0.038066	-0.060882	-0.011078	0.001426	-0.071156	-0.059771	-0.023453
Rm - RF	-0.078086	-0.045169	0.024873	-0.015566	-0.149165	0.027613	-0.041590	-0.020027	-0.070580	-0.005688	0.026543	-0.035460	-0.046824	-0.019098	-0.022100	-0.008774	-0.081333	-0.069913	-0.042583

Portfolio	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	
T-Bill (Inv. Adj)	-0.098657	-0.024400	0.012724	0.011589	-0.128089	0.013801	0.006491	-0.019690	-0.052558	0.017714	0.004015	-0.026416	-0.039413	0.000073	0.066623	-0.019830	-0.041054	-0.093797	-0.059938		
ISE100 Return	-0.089334	-0.024278	-0.010606	-0.002308	-0.158083	0.020016	-0.024771	-0.036886	-0.049598	0.008477	0.033806	-0.011684	-0.024272	-0.001996	0.002753	-0.002027	-0.079295	-0.074212	-0.079919		
Rm - RF	-0.094941	-0.026485	0.011150	0.048616	-0.104225	0.010130	-0.012328	-0.035180	-0.012203	-0.016017	0.026609	-0.032521	-0.032225	-0.034287	0.007114	0.007523	-0.061243	-0.074400	-0.050772		
1		0.000000																			
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### Appendix IX: Weekly Portfolio Returns and Market Risk Premium (Test Period 2002-2003)

Date	01.04.02	01.11.02	01.18.02	01.25.02	02.01.02	02.08.02	02.15.02	02.22.02	03.01.02	03.08.02	03.15.02	03.22.02	03.29.02	04.05.02	04.12.02	04.19.02	04.26.02	05.03.02	05.10.02
T-Bill (Infl Adj)	0.014105	0.014087	0.014045	0.014092	0.014092	0.014099	0.014092	0.014097	0.014097	0.014099	0.014092	0.014099	0.014099	0.014099	0.014099	0.014099	0.014099	0.014099	0.014099
ISE 100 Return	-0.047086	-0.045374	-0.045374	-0.045374	-0.045374	-0.045374	-0.045374	-0.045374	-0.045374	-0.045374	-0.045374	-0.045374	-0.045374	-0.045374	-0.045374	-0.045374	-0.045374	-0.045374	-0.045374
Rm - RF	-0.061172	-0.059949	-0.059949	-0.059949	-0.059949	-0.059949	-0.059949	-0.059949	-0.059949	-0.059949	-0.059949	-0.059949	-0.059949	-0.059949	-0.059949	-0.059949	-0.059949	-0.059949	-0.059949

Portfolio	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	
T-Bill (Infl Adj)	-0.046116	-0.011533	-0.011533	-0.011533	-0.011533	-0.011533	-0.011533	-0.011533	-0.011533	-0.011533	-0.011533	-0.011533	-0.011533	-0.011533	-0.011533	-0.011533	-0.011533	-0.011533	-0.011533	-0.011533	
ISE 100 Return	-0.044819	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548	-0.010548
Rm - RF	-0.008697	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915	-0.000915

Date	05.17.02	05.24.02	05.31.02	06.07.02	06.14.02	06.21.02	06.28.02	07.05.02	07.12.02	07.19.02	07.26.02	08.02.02	08.09.02	08.16.02	08.23.02	08.30.02	09.06.02	09.13.02	09.20.02	
T-Bill (Infl Adj)	0.008930	0.008925	0.008925	0.008925	0.008925	0.008925	0.008925	0.008925	0.008925	0.008925	0.008925	0.008925	0.008925	0.008925	0.008925	0.008925	0.008925	0.008925	0.008925	
ISE 100 Return	-0.092566	-0.026450	-0.026450	-0.026450	-0.026450	-0.026450	-0.026450	-0.026450	-0.026450	-0.026450	-0.026450	-0.026450	-0.026450	-0.026450	-0.026450	-0.026450	-0.026450	-0.026450	-0.026450	-0.026450
Rm - RF	-0.101497	-0.017526	-0.017526	-0.017526	-0.017526	-0.017526	-0.017526	-0.017526	-0.017526	-0.017526	-0.017526	-0.017526	-0.017526	-0.017526	-0.017526	-0.017526	-0.017526	-0.017526	-0.017526	-0.017526

Portfolio	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20
T-Bill (Infl Adj)	-0.100993	-0.069026	-0.047503	-0.043584	-0.046891	-0.006671	0.035568	-0.031253	-0.018232	0.086230	-0.036812	0.078648	-0.033292	-0.050350	-0.053614	-0.039971	-0.029808	-0.041070	0.014050	
ISE 100 Return	-0.108056	0.057005	-0.050014	-0.038761	-0.030751	-0.001467	0.054771	-0.016102	-0.024869	0.094820	-0.037271	0.075683	-0.037213	-0.034914	-0.058805	0.024017	0.016950	-0.058702	0.029446	
Rm - RF	-0.077063	0.067509	-0.061517	-0.042315	-0.043522	-0.003014	0.066448	-0.032305	-0.016340	0.101640	-0.044083	0.091231	-0.044187	-0.045828	-0.094619	0.039044	0.017051	-0.087474	0.033596	



Date	06.20.03	06.27.03	07.04.03	07.11.03	07.18.03	07.25.03	08.01.03	08.08.03	08.15.03	08.22.03	08.29.03	09.05.03	09.12.03	09.19.03	09.26.03	10.03.03	10.10.03	10.17.03	10.24.03	
T-Bill (Infl Adj)	0.0053998	0.0053863	0.00491264	0.0049298	0.00493102	0.00492857	0.00492535	0.00494637	0.00494637	0.00496316	0.00496437	0.00496437	0.00496437	0.00496437	0.00496437	0.00496437	0.00496437	0.00496437	0.00496437	0.00496437
ISE 100 Return	0.006518	0.018113	-0.044539	-0.004577	0.037724	-0.027597	0.003648	0.084569	0.027227	0.010353	-0.032971	0.001716	0.089056	0.071934	0.006847	0.032284	0.081743	-0.002531	-0.034096	-0.034096
Rm - RF	0.001118	0.012775	-0.049432	-0.009507	0.032793	-0.032536	0.001196	0.080105	0.022762	0.005890	-0.037436	0.084947	0.067834	0.002848	0.028605	0.078067	-0.006183	-0.037772	-0.037772	-0.037772
Portfolio	1	-0.008141	0.011282	-0.053725	-0.016632	0.031052	-0.024113	-0.013228	0.083816	0.022080	0.009530	-0.043416	-0.001054	0.100560	0.075307	0.006510	0.038424	0.096751	-0.025255	-0.026057
	2	0.007313	0.006531	-0.048713	0.000306	0.029246	-0.032615	0.011424	0.086611	0.021809	0.010695	-0.038029	-0.009192	0.095668	0.077070	0.011348	0.013612	0.066232	0.003548	-0.031403
	3	-0.010695	0.003939	-0.069666	-0.012942	0.018909	-0.019020	0.003827	0.074348	0.023399	0.018747	-0.033276	-0.002857	0.059764	0.005242	0.005342	0.020379	0.119335	-0.006595	-0.026635
	4	-0.026457	-0.009207	-0.048399	-0.021313	0.005699	-0.019977	-0.018901	0.067283	0.007159	-0.015652	0.005248	0.018474	0.027187	0.007884	0.001934	0.024523	0.001934	-0.011502	0.010633
	5	-0.021984	-0.004074	-0.056475	-0.036790	0.022638	-0.023979	-0.016052	0.050114	0.012567	0.003956	0.005361	0.006491	0.060218	0.038974	0.007887	0.022051	0.044737	-0.008024	-0.009510
	6	-0.016581	0.017068	-0.035691	-0.022678	0.001632	-0.024193	-0.007128	0.036171	0.010929	-0.004836	-0.003033	-0.016594	0.050624	0.018804	0.003160	-0.004932	0.050937	-0.008248	0.010764
	7	-0.004420	0.000854	-0.049784	-0.006988	0.034120	-0.013225	0.019453	0.041363	0.019453	-0.000087	-0.009752	0.006687	0.014233	0.046163	0.008558	0.014295	0.026536	0.004892	0.006650
	8	-0.014609	0.004680	-0.042083	-0.019220	0.010698	-0.008040	0.046176	0.027445	-0.006363	-0.028032	-0.002566	0.049255	0.053548	0.041462	-0.006170	0.006170	0.083701	-0.014851	-0.019056
	9	-0.021814	0.006032	-0.019626	-0.005250	0.012335	-0.021138	-0.000566	0.036957	-0.001261	0.003167	-0.047308	-0.001813	0.028932	0.025888	0.009379	0.021835	0.043123	-0.012493	-0.017577
	10	-0.020562	-0.013776	-0.025250	-0.023133	0.023432	-0.056258	-0.022049	-0.024562	-0.013893	-0.018942	-0.018942	0.012356	0.050127	0.035801	0.006186	-0.003463	0.051183	0.002738	-0.005752
	11	-0.010851	0.009591	-0.002634	-0.042200	0.023431	0.003344	-0.042648	0.024302	-0.010887	-0.012613	-0.019829	0.015919	0.023108	0.030998	0.003876	-0.018813	0.020628	0.014649	0.002882
	12	0.004384	-0.001869	-0.033671	-0.037697	0.011508	0.011361	-0.003111	0.045306	-0.000718	-0.013277	-0.021932	0.051342	0.019090	0.045374	0.006459	0.001593	0.033876	-0.009249	0.012159
	13	0.007399	0.011869	-0.029145	-0.014342	0.015387	-0.025363	0.014408	0.029710	0.047475	-0.036138	0.019074	0.008240	0.029187	0.000275	0.000275	0.003131	0.033756	-0.002105	0.038289
	14	0.001204	-0.001799	-0.051642	-0.017650	0.016917	-0.023245	-0.003283	0.025761	0.010255	0.002808	-0.005125	0.022088	0.015544	0.036710	0.015639	-0.011000	0.018202	0.003557	0.016121
	15	0.006190	0.009546	-0.034373	-0.009224	0.015631	-0.016422	-0.009811	0.014085	-0.007460	-0.035981	-0.021684	-0.012605	0.012139	0.029755	0.030903	-0.009062	0.037147	0.005510	0.015518
	16	0.007705	0.001007	-0.020403	-0.024126	0.016222	-0.019310	-0.012979	0.042106	-0.017317	0.000614	0.014474	0.002369	0.039145	-0.003974	-0.014616	0.043063	-0.017970	0.003243	0.012159
	17	-0.009598	0.026756	-0.061686	0.004695	0.027339	-0.026532	-0.030307	0.034226	0.049938	0.012656	0.010482	-0.009910	0.041638	0.058375	-0.012600	0.002820	0.064377	-0.011849	0.015991
	18	-0.014472	0.014231	-0.048420	-0.027583	0.044247	0.002395	-0.003612	0.013825	0.000427	-0.008268	0.019557	0.020250	0.010591	0.000762	-0.000607	-0.005487	0.032261	-0.003286	0.004492
	19	-0.011520	-0.005698	-0.050875	-0.031125	0.033808	-0.012095	0.024855	0.009229	0.001245	-0.009848	0.011366	-0.004533	0.014348	0.022500	0.017097	0.006318	0.072559	0.011419	-0.015568
	20	0.049893	-0.008280	-0.030039	-0.007163	0.002274	-0.026099	-0.012734	0.024696	0.022939	0.026241	-0.059749	-0.014811	0.051766	0.060423	-0.015401	0.013721	0.034032	-0.012867	0.013576

Date	10.31.03	11.07.03	11.14.03	11.21.03	11.28.03	12.05.03	12.12.03	12.19.03	12.26.03
T-Bill (Infl Adj)	0.00368034	0.00338979	0.00338979	0.00338979	0.000000	0.00321695	0.00321126	0.00320784	0.00320443
ISE 100 Return	0.058018	-0.002941	0.023443	-0.018862	0.000000	0.044895	0.026923	0.071182	-0.011563
Rm - RF	0.054338	-0.006332	0.020054	-0.022252	0.000000	0.041678	0.023712	0.067974	-0.014767
Portfolio	1	0.045474	-0.003566	0.041887	-0.017385	0.000000	0.043039	0.056249	-0.013866
	2	0.070723	0.010172	0.019788	-0.015859	0.000000	0.073388	0.028571	0.054224
	3	0.021128	-0.024177	0.058897	-0.021803	0.000000	0.059419	0.025627	0.060108
	4	0.005769	0.010632	0.023948	-0.003976	0.000000	0.059500	0.011488	0.048843
	5	0.020298	-0.012139	0.053470	0.014570	0.000000	0.031433	0.011687	0.049741
	6	0.017244	0.010550	0.009519	0.003764	0.000000	0.066614	0.041986	0.056259
	7	0.003306	0.021728	0.017234	-0.006203	0.000000	0.039318	0.014537	-0.001565
	8	0.023888	-0.015030	0.036440	-0.014570	0.000000	0.055353	0.038264	0.061089
	9	0.039436	-0.019560	0.011042	-0.011448	0.000000	0.056706	0.020224	0.049429
	10	0.007702	-0.028852	0.018514	-0.002143	0.000000	0.032274	0.001960	0.056780
	11	-0.004994	-0.027618	0.002773	-0.010539	0.000000	0.038812	0.055911	0.046071
	12	0.049706	-0.001149	0.038508	-0.002976	0.000000	0.024107	0.006589	0.044855
	13	0.015712	-0.012006	0.004437	-0.011299	0.000000	0.060205	-0.004042	0.029267
	14	0.024705	-0.030370	0.011169	-0.011445	0.000000	0.054127	-0.004257	0.048216
	15	0.006090	-0.007860	0.036376	-0.035274	0.000000	0.038121	0.024890	0.034461
	16	0.028483	-0.020939	0.019436	0.008186	0.000000	0.071546	0.024190	0.044070
	17	0.026582	-0.030678	0.003598	-0.013508	0.000000	0.032171	0.035056	0.036250
	18	0.007571	0.011074	0.023591	-0.006725	0.000000	0.021043	0.014676	0.044001
	19	-0.005313	0.014122	0.026982	-0.014784	0.000000	0.038744	0.018719	0.030930
	20	-0.003329	-0.025361	0.026532	-0.000417	0.000000	0.017564	0.011075	0.009302

**Appendix X: Weekly Portfolio Returns and Market Risk Premium (Test Period 2003-2004)**

Date	T-bill (Infl Adj)	ISE 100 Return	Rm - RF	05.16.03	05.23.03	05.30.03	06.06.03	06.13.03	06.20.03	07.04.03	07.11.03	07.18.03	07.25.03	08.01.03	08.08.03	08.15.03	08.22.03	08.29.03	09.05.03	09.12.03	09.19.03
1	0.00480284	0.00480162	0.00013116	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162
2	0.0060230	0.0060230	0.0012072	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162
3	0.0065303	0.0065303	0.0017145	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162	0.00480162
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Date	T-bill (Infl Adj)	ISE 100 Return	Rm - RF	02.14.03	02.21.03	02.28.03	03.07.03	03.14.03	03.21.03	03.28.03	04.04.03	04.11.03	04.18.03	04.25.03	05.02.03	05.09.03
1	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
2	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
3	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
4	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
5	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
6	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
7	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
8	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
9	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
10	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
11	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
12	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
13	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
14	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
15	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
16	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
17	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
18	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
19	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000
20	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000	0.000000

Date	09.26.03	10.03.03	10.10.03	10.17.03	10.24.03	10.31.03	11.07.03	11.14.03	11.21.03	11.28.03	12.05.03	12.12.03	12.19.03	12.26.03	01.02.04	01.09.04	01.16.04	01.23.04	01.30.04
T-Bill (1m/1Ard)	0.0049946	0.00367918	0.00367221	0.00367221	0.00367221	0.00367221	0.00367221	0.00367221	0.00367221	0.00367221	0.00367221	0.00367221	0.00367221	0.00367221	0.00367221	0.00367221	0.00367221	0.00367221	0.00367221
ISE 100 Return	0.006947	0.032284	0.081743	-0.025111	-0.034096	0.008018	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979
Rm - Rf	0.002848	0.028605	0.078667	-0.06183	-0.037772	0.054338	0.006312	0.020054	-0.022352	0.006000	0.041678	0.023712	0.067974	-0.014767	0.059274	0.037146	-0.087795	0.009071	-0.073114

Portfolio	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20
T-Bill (1m/1Ard)	0.0049946	0.029832	0.120010	-0.026359	-0.044434	0.049183	-0.009207	0.030557	-0.022481	0.000000	0.044657	0.051073	0.090307	0.002006	0.063596	0.058242	-0.079746	-0.005683	-0.089082	-0.075985
ISE 100 Return	0.006947	0.032284	0.081743	-0.025111	-0.034096	0.008018	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979
Rm - Rf	0.002848	0.028605	0.078667	-0.06183	-0.037772	0.054338	0.006312	0.020054	-0.022352	0.006000	0.041678	0.023712	0.067974	-0.014767	0.059274	0.037146	-0.087795	0.009071	-0.073114	-0.073114

Date	02.06.04	02.13.04	02.20.04	02.27.04	03.05.04	03.12.04	03.19.04	03.26.04	04.02.04	04.09.04	04.16.04	04.23.04	04.30.04	05.07.04	05.14.04	05.21.04	05.28.04	06.04.04	06.11.04
T-Bill (1m/1Ard)	0.00235539	0.00235539	0.00235539	0.00235539	0.00235539	0.00235539	0.00235539	0.00235539	0.00235539	0.00235539	0.00235539	0.00235539	0.00235539	0.00235539	0.00235539	0.00235539	0.00235539	0.00235539	0.00235539
ISE 100 Return	-0.017147	0.11262	-0.021096	0.015223	0.014532	0.010315	0.031483	0.039768	-0.016994	-0.049015	-0.012158	0.008374	-0.075964	-0.058302	-0.028076	0.005881	0.041182	0.021709	-0.002470
Rm - Rf	-0.019502	0.110906	-0.023430	0.012865	0.012656	0.008421	0.031386	0.037874	-0.018580	-0.050598	-0.013738	0.006790	-0.076856	-0.059651	-0.029440	0.004521	0.039820	0.020320	-0.003880

Portfolio	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20
T-Bill (1m/1Ard)	0.0049946	0.029832	0.120010	-0.026359	-0.044434	0.049183	-0.009207	0.030557	-0.022481	0.000000	0.044657	0.051073	0.090307	0.002006	0.063596	0.058242	-0.079746	-0.005683	-0.089082	-0.075985
ISE 100 Return	0.006947	0.032284	0.081743	-0.025111	-0.034096	0.008018	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979	0.0038979
Rm - Rf	0.002848	0.028605	0.078667	-0.06183	-0.037772	0.054338	0.006312	0.020054	-0.022352	0.006000	0.041678	0.023712	0.067974	-0.014767	0.059274	0.037146	-0.087795	0.009071	-0.073114	-0.073114

Date	06.18.04	06.25.04	07.02.04	07.09.04	07.16.04	07.23.04	07.30.04	08.06.04	08.13.04	08.20.04	08.27.04	09.03.04	09.10.04	09.17.04	09.24.04	10.01.04	10.08.04	10.15.04	10.22.04
T-Bill (1m1.A4d)	0.00144293	0.00142722	0.00155763	0.00155023	0.00155236	0.00155238	0.00155238	0.00166881	0.00166987	0.00166987	0.0016794	0.0016794	0.0016794	0.00150835	0.00150835	0.00168241	0.00168241	0.00168135	0.00170037
ISE 100 Return	-0.040457	0.022792	-0.004862	0.044304	-0.013295	0.024898	-0.010024	-0.019024	0.024159	0.029071	0.045282	0.045282	0.045282	-0.009174	0.068374	-0.026572	0.055029	-0.020881	-0.010426
Rm - Rf	-0.041859	0.021364	-0.006412	0.057794	-0.014847	0.023325	-0.011693	-0.020694	0.022489	0.027392	0.041785	0.041785	0.041785	-0.009683	0.066855	-0.028259	0.053346	-0.022562	-0.012126

Portfolio	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20
T-Bill (1m1.A4d)	-0.052495	0.006210	0.059123	-0.001389	0.046378	-0.016378	0.016429	0.003339	-0.009079	0.004046	0.039946	0.016047	0.021388	-0.017771	0.062069	-0.033227	0.033227	-0.033227	-0.033227	-0.033227
ISE 100 Return	-0.044068	0.013355	0.047309	0.002042	0.031507	-0.011632	0.026612	0.000024	-0.029928	0.036965	0.023113	0.034099	0.020842	-0.015254	0.061088	-0.012169	0.033227	-0.033227	-0.033227	-0.033227
Rm - Rf	-0.044068	0.013355	0.047309	0.002042	0.031507	-0.011632	0.026612	0.000024	-0.029928	0.036965	0.023113	0.034099	0.020842	-0.015254	0.061088	-0.012169	0.033227	-0.033227	-0.033227	-0.033227

## APPENDIX XI

**Portfolio Returns and Betas (Test Period 1999-2000)**

		RF= 0,01243		
Portfolio	Rp	Rp-Rf	Beta	
1	0,011581	-0,000854	0,865977	
2	0,009792	-0,002643	0,912686	
3	0,010313	-0,002122	0,766603	
4	0,011307	-0,001128	0,811927	
5	0,009325	-0,003110	0,809544	
6	0,010768	-0,001667	0,834214	
7	0,009820	-0,002615	0,794672	
8	0,009041	-0,003394	0,767550	
9	0,009581	-0,002854	0,950337	
10	0,007713	-0,004722	0,814291	
11	0,011006	-0,001429	0,801159	
12	0,011749	-0,000686	0,841804	
13	0,005284	-0,007151	0,813572	
14	0,011670	-0,000765	0,812659	
15	0,009367	-0,003068	0,791574	
16	0,011069	-0,001366	0,683157	
17	0,011643	-0,000792	0,759195	
18	0,006707	-0,005728	0,743984	
19	0,008511	-0,003924	0,803318	
20	0,009148	-0,003287	0,578120	

**Portfolio Returns and Betas (Test Period 2000-2001)**

		RF= 0,01100		
Portfolio	Rp	Rp-Rf	Beta	
1	-0,004722	-0,015725	1,037127	
2	-0,004544	-0,015548	0,913098	
3	-0,002356	-0,013359	0,819785	
4	-0,002600	-0,013603	0,841048	
5	-0,000822	-0,011825	0,892936	
6	0,000124	-0,010880	0,748306	
7	0,000007	-0,010997	0,703160	
8	0,001163	-0,009841	0,819488	
9	-0,000634	-0,011637	0,736878	
10	0,000003	-0,011001	0,764706	
11	-0,002247	-0,013251	0,731445	
12	-0,004316	-0,015320	0,729554	
13	-0,001982	-0,012986	0,735761	
14	0,002156	-0,008848	0,685386	
15	-0,001106	-0,012110	0,729028	
16	-0,002542	-0,013546	0,651461	
17	-0,002643	-0,013647	0,699272	
18	0,002115	-0,008889	0,701576	
19	0,004033	-0,006971	0,495358	
20	0,000014	-0,010990	0,469694	

**Portfolio Returns and Betas (Test Period 2001-2002)**

		RF= 0,00976		
Portfolio	Rp	Rp-Rf	Beta	
1	-0,000233	-0,009998	1,039979	
2	0,001025	-0,008739	0,951202	
3	0,001129	-0,008635	0,845569	
4	-0,000156	-0,009920	0,814211	
5	0,001827	-0,007937	0,746618	
6	0,001725	-0,008040	0,695702	
7	0,001745	-0,008020	0,739757	
8	0,001884	-0,007881	0,693654	
9	-0,003581	-0,013346	0,761567	
10	0,001138	-0,008626	0,646790	
11	0,004188	-0,005577	0,678704	
12	0,000797	-0,008967	0,514682	
13	0,002379	-0,007386	0,640276	
14	0,002452	-0,007313	0,612963	
15	0,003738	-0,006026	0,550951	
16	0,003653	-0,006112	0,695316	
17	0,001273	-0,008491	0,562035	
18	0,001215	-0,008550	0,572442	
19	0,003731	-0,006034	0,362297	
20	0,004648	-0,005117	0,346676	

Portfolio Returns and Betas (Test Period 2002-2003)

Rf = 0,00683

Portfolio	Rp	Rp-Rf	Beta
1	0,000821	-0,006014	1,011574
2	0,003678	-0,003157	1,002865
3	0,003392	-0,003443	0,997578
4	0,001544	-0,003291	0,972237
5	-0,001210	-0,008045	0,998831
6	0,002716	-0,004119	0,991606
7	0,004069	-0,002765	0,984201
8	0,000119	-0,006715	0,983529
9	0,003062	-0,003773	0,996438
10	0,002196	-0,004639	0,992138
11	0,002373	-0,004462	0,988024
12	0,000402	-0,006433	0,991471
13	0,003336	-0,003499	0,991694
14	-0,000986	-0,007821	0,990875
15	0,002315	-0,004520	0,971976
16	0,002587	-0,004248	1,000512
17	0,001309	-0,005526	0,987091
18	-0,000483	-0,007318	0,983990
19	0,003406	-0,003429	0,981329
20	0,000004	-0,006831	0,971452

Portfolio Returns and Betas (Test Period 2003-2004)

Rf = 0,003175

Portfolio	Rp	Rp-Rf	Beta
1	0,005635	0,002459	1,174786
2	0,006286	0,003110	1,068352
3	0,005994	0,002818	1,069546
4	0,009029	0,005853	1,031967
5	0,007998	0,004823	1,002752
6	0,007531	0,004355	0,830292
7	0,004714	0,001538	0,837292
8	0,005504	0,002328	0,896021
9	0,006981	0,003805	0,906960
10	0,005938	0,002762	0,813022
11	0,004642	0,001466	0,849007
12	0,006761	0,003585	0,840340
13	0,004222	0,001046	0,902145
14	0,006333	0,003158	0,789791
15	0,004639	0,001463	0,842745
16	0,005866	0,002691	0,695004
17	0,003757	0,000581	0,721553
18	0,004737	0,001562	0,765460
19	0,005820	0,002644	0,770852
20	0,008714	0,005539	0,545317



## APPENDIX XII: Example of Beta Estimation

0,740 Beta of ABANA

0,770 R-squared of the regression

Date	ISE100	ABANA	T-bill	R (ISE 100)	R (ABANA)	ER (ISE 100)	ER (ABANA)
01.06.95	26949,82	2025	0,025504				
01.13.95	26548,19	1925	0,025544	-0,015015	-0,050644	-0,040559	-0,076187
01.20.95	25519,86	1400	0,025500	-0,039505	-0,318454	-0,065005	-0,343954
01.27.95	25246,08	1400	0,025513	-0,010786	0,000000	-0,036299	-0,025513
02.03.95	25463,39	1350	0,024820	0,008571	-0,036368	-0,016250	-0,061188
02.10.95	26918,00	1375	0,024838	0,055553	0,018349	0,030716	-0,006488
02.17.95	28214,19	1300	0,024833	0,047030	-0,056089	0,022197	-0,080923
02.24.95	29720,10	1450	0,024799	0,051999	0,109199	0,027199	0,084400
03.03.95	29781,63	1350	0,024251	0,002068	-0,071459	-0,022183	-0,095710
03.10.95	33232,88	1275	0,024268	0,109648	-0,057158	0,085380	-0,081427
03.17.95	33802,38	1275	0,024264	0,016991	0,000000	-0,007272	-0,024264
03.24.95	36693,19	1375	0,024264	0,082060	0,075508	0,057796	0,051244
03.31.95	39837,33	1350	0,024213	0,082213	-0,018349	0,058000	-0,042562
04.07.95	44622,88	1950	0,018074	0,113442	0,367725	0,095369	0,349651
04.14.95	48203,25	2275	0,018052	0,077180	0,154151	0,059128	0,136099
04.21.95	54653,93	2425	0,018001	0,125595	0,063851	0,107594	0,045850
04.28.95	46615,20	2250	0,018037	-0,159094	-0,074901	-0,177132	-0,092939
05.05.95	51227,44	2450	0,016366	0,094349	0,085158	0,077983	0,068792
05.12.95	51227,44	2450	0,016328	0,000000	0,000000	-0,016328	-0,016328
05.19.95	47181,80	2200	0,016355	-0,082267	-0,107631	-0,098622	-0,123986
05.26.95	46500,33	1975	0,016359	-0,014549	-0,107889	-0,030908	-0,124248
06.02.95	48653,10	2250	0,016511	0,045256	0,130362	0,028745	0,113850
06.09.95	53781,29	1300	0,016456	0,100210	-0,548566	0,083754	-0,565022
06.16.95	52489,34	1050	0,016487	-0,024316	-0,213574	-0,040803	-0,230061
06.23.95	53042,90	1175	0,016449	0,010491	0,112478	-0,005958	0,096029
06.30.95	48233,01	1050	0,016411	-0,095057	-0,112478	-0,111468	-0,128889
07.07.95	46357,35	1175	0,016453	-0,039664	0,112478	-0,056117	0,096025
07.14.95	47884,48	1175	0,016408	0,032412	0,000000	0,016003	-0,016408
07.21.95	47486,61	1450	0,016429	-0,008344	0,210295	-0,024773	0,193866
07.28.95	52777,52	1600	0,016432	0,105638	0,098440	0,089205	0,082008
08.04.95	49103,94	1775	0,016941	-0,072146	0,103797	-0,089087	0,086856
08.11.95	45888,04	1900	0,016930	-0,067735	0,068053	-0,084665	0,051123
08.18.95	45801,88	2025	0,016934	-0,001879	0,063716	-0,018813	0,046782
08.25.95	45542,98	1775	0,016937	-0,005669	-0,131769	-0,022606	-0,148707
09.01.95	45690,62	1700	0,017416	0,003237	-0,043172	-0,014179	-0,060588
09.08.95	46194,89	1850	0,017402	0,010976	0,084557	-0,006425	0,067156
09.15.95	44782,31	1875	0,017416	-0,031056	0,013423	-0,048472	-0,003993
09.22.95	41608,43	1775	0,017384	-0,073510	-0,054808	-0,090894	-0,072192
09.29.95	41707,62	1725	0,017344	0,002381	-0,028573	-0,014963	-0,045918
10.06.95	45767,72	2375	0,017094	0,092895	0,319770	0,075801	0,302676
10.13.95	43577,71	2575	0,017084	-0,049033	0,080852	-0,066117	0,063769
10.20.95	49029,62	3650	0,017087	0,117879	0,348878	0,100792	0,331791
10.27.95	47145,87	3450	0,017052	-0,039178	-0,056353	-0,056230	-0,073405
11.03.95	44928,30	2950	0,016547	-0,048179	-0,156569	-0,064726	-0,173116
11.10.95	42370,63	2450	0,016572	-0,058612	-0,185717	-0,075184	-0,202289
11.17.95	42833,18	2550	0,016596	0,010858	0,040005	-0,005738	0,023409
11.24.95	42159,46	3150	0,016565	-0,015854	0,211309	-0,032419	0,194744
12.01.95	39337,14	2450	0,015478	-0,069290	-0,251314	-0,084768	-0,266792
12.08.95	41846,42	2450	0,015467	0,061837	0,000000	0,046370	-0,015467
12.15.95	39669,01	2350	0,015471	-0,053436	-0,041673	-0,068907	-0,057144
12.22.95	42236,25	2225	0,015420	0,062709	-0,054658	0,047289	-0,070078
12.29.95	40024,58	2100	0,015339	-0,053785	-0,057820	-0,069124	-0,073158

**APPENDIX XIII: Example of T-bill Rate Adjustment**

Date	T-bill (USA)	USA Inflation		Inflation (Turkey)	Turkish Inflation Added T-bill Rate	Weekly Adjusted T-bill Rate
		Inflation Rate (USA)	Deducted T-bill Rate			
01.06.95	0,0578	0,0280	0,0298	1,2589	1,3262	0,025504
01.13.95	0,0587	0,0280	0,0307	1,2589	1,3283	0,025544
01.20.95	0,0577	0,0280	0,0297	1,2589	1,3260	0,025500
01.27.95	0,0580	0,0280	0,0300	1,2589	1,3267	0,025513
02.03.95	0,0579	0,0280	0,0299	1,2242	1,2907	0,024820
02.10.95	0,0583	0,0280	0,0303	1,2242	1,2916	0,024838
02.17.95	0,0582	0,0280	0,0302	1,2242	1,2913	0,024833
02.24.95	0,0574	0,0280	0,0294	1,2242	1,2896	0,024799
03.03.95	0,0573	0,0280	0,0293	1,1967	1,2611	0,024251
03.10.95	0,0577	0,0280	0,0297	1,1967	1,2619	0,024268
03.17.95	0,0576	0,0280	0,0296	1,1967	1,2617	0,024264
03.24.95	0,0576	0,0280	0,0296	1,1967	1,2617	0,024264
03.31.95	0,0564	0,0280	0,0284	1,1967	1,2591	0,024213
04.07.95	0,0576	0,0280	0,0296	0,8841	0,9398	0,018074
04.14.95	0,0570	0,0280	0,0290	0,8841	0,9387	0,018052
04.21.95	0,0556	0,0280	0,0276	0,8841	0,9361	0,018001
04.28.95	0,0566	0,0280	0,0286	0,8841	0,9379	0,018037
05.05.95	0,0574	0,0280	0,0294	0,7981	0,8510	0,016366
05.12.95	0,0563	0,0280	0,0283	0,7981	0,8490	0,016328
05.19.95	0,0571	0,0280	0,0291	0,7981	0,8505	0,016355
05.26.95	0,0572	0,0280	0,0292	0,7981	0,8507	0,016359
06.02.95	0,0564	0,0280	0,0284	0,8073	0,8586	0,016511
06.09.95	0,0548	0,0280	0,0268	0,8073	0,8557	0,016456
06.16.95	0,0557	0,0280	0,0277	0,8073	0,8573	0,016487
06.23.95	0,0546	0,0280	0,0266	0,8073	0,8553	0,016449
06.30.95	0,0535	0,0280	0,0255	0,8073	0,8534	0,016411
07.07.95	0,0553	0,0280	0,0273	0,8063	0,8556	0,016453
07.14.95	0,0540	0,0280	0,0260	0,8063	0,8532	0,016408
07.21.95	0,0546	0,0280	0,0266	0,8063	0,8543	0,016429
07.28.95	0,0547	0,0280	0,0267	0,8063	0,8545	0,016432
08.04.95	0,0544	0,0280	0,0264	0,8325	0,8809	0,016941
08.11.95	0,0541	0,0280	0,0261	0,8325	0,8804	0,016930
08.18.95	0,0542	0,0280	0,0262	0,8325	0,8806	0,016934
08.25.95	0,0543	0,0280	0,0263	0,8325	0,8807	0,016937
09.01.95	0,0534	0,0280	0,0254	0,8584	0,9056	0,017416
09.08.95	0,0530	0,0280	0,0250	0,8584	0,9049	0,017402
09.15.95	0,0534	0,0280	0,0254	0,8584	0,9056	0,017416
09.22.95	0,0525	0,0280	0,0245	0,8584	0,9040	0,017384
09.29.95	0,0514	0,0280	0,0234	0,8584	0,9019	0,017344
10.06.95	0,0534	0,0280	0,0254	0,8421	0,8889	0,017094
10.13.95	0,0531	0,0280	0,0251	0,8421	0,8883	0,017084
10.20.95	0,0532	0,0280	0,0252	0,8421	0,8885	0,017087
10.27.95	0,0522	0,0280	0,0242	0,8421	0,8867	0,017052
11.03.95	0,0529	0,0280	0,0249	0,8153	0,8605	0,016547
11.10.95	0,0536	0,0280	0,0256	0,8153	0,8617	0,016572
11.17.95	0,0543	0,0280	0,0263	0,8153	0,8630	0,016596
11.24.95	0,0534	0,0280	0,0254	0,8153	0,8614	0,016565
12.01.95	0,0532	0,0280	0,0252	0,7605	0,8048	0,015478
12.08.95	0,0529	0,0280	0,0249	0,7605	0,8043	0,015467
12.15.95	0,0530	0,0280	0,0250	0,7605	0,8045	0,015471
12.22.95	0,0515	0,0280	0,0235	0,7605	0,8018	0,015420
12.29.95	0,0491	0,0280	0,0211	0,7605	0,7976	0,015339
01.05.96	0,0504	0,0300	0,0204	0,7811	0,8174	0,015719

## CURRICULUM VITAE

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